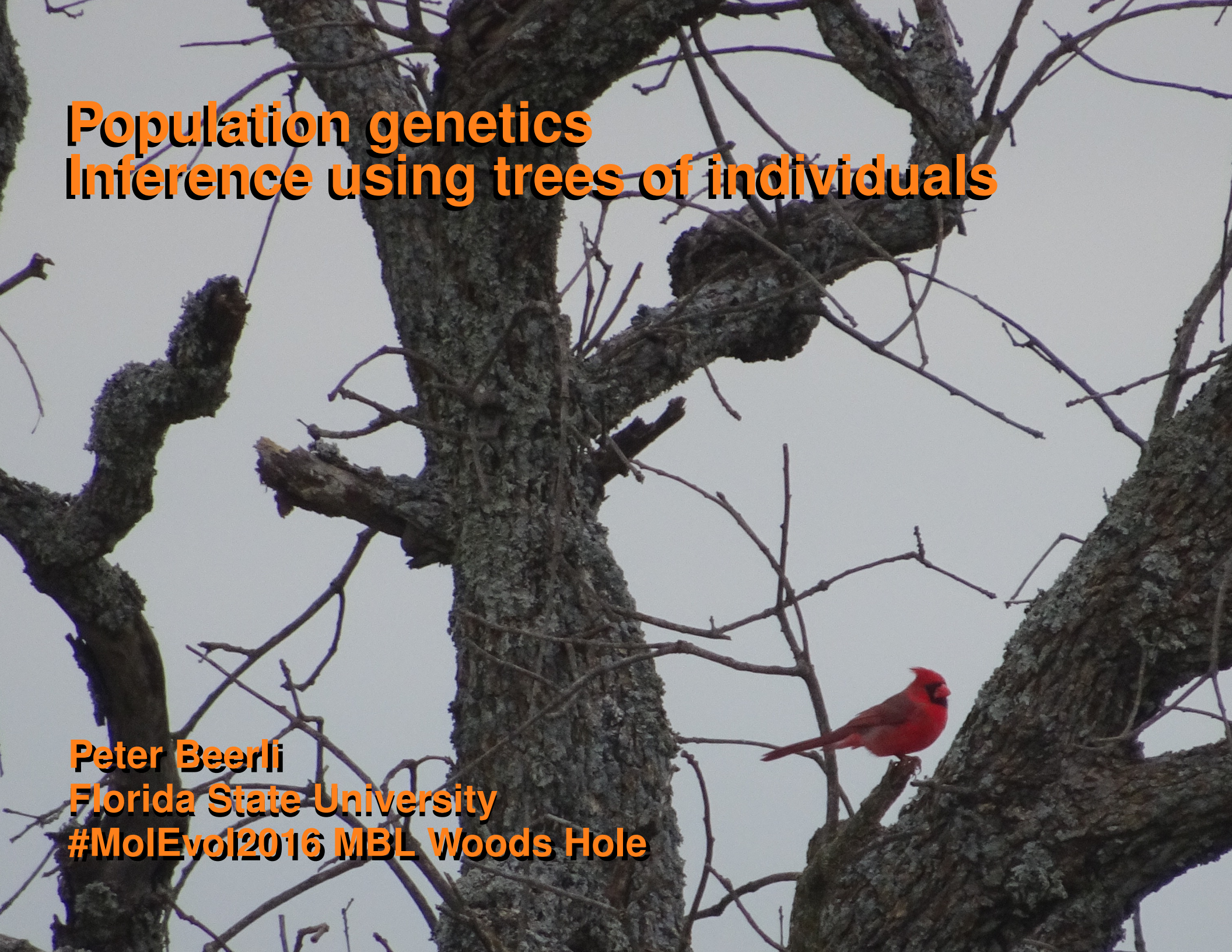
Population genetics
Inference using trees of individuals



Peter Beerli
Florida State University
#MolEvol2016 MBL Woods Hole

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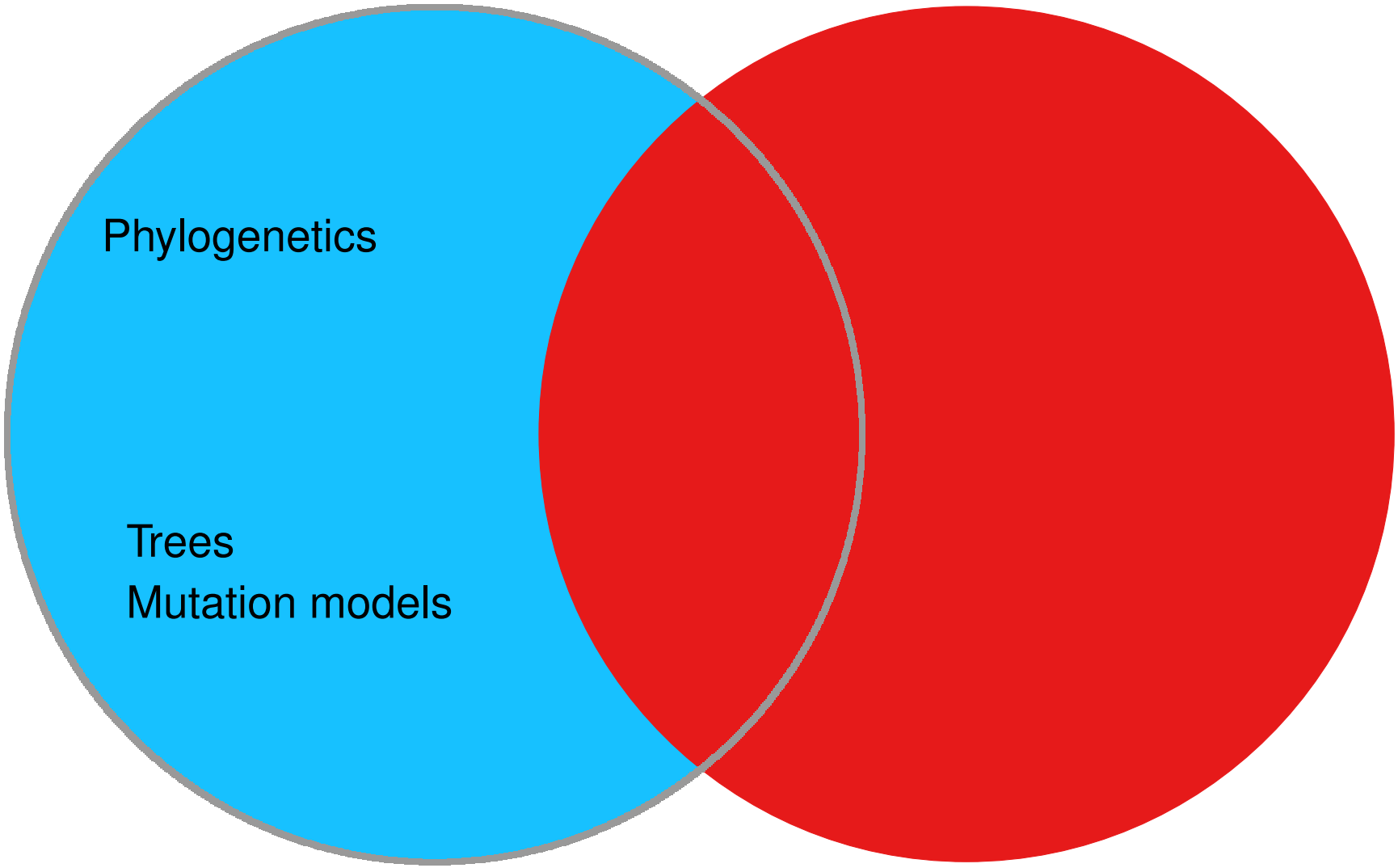
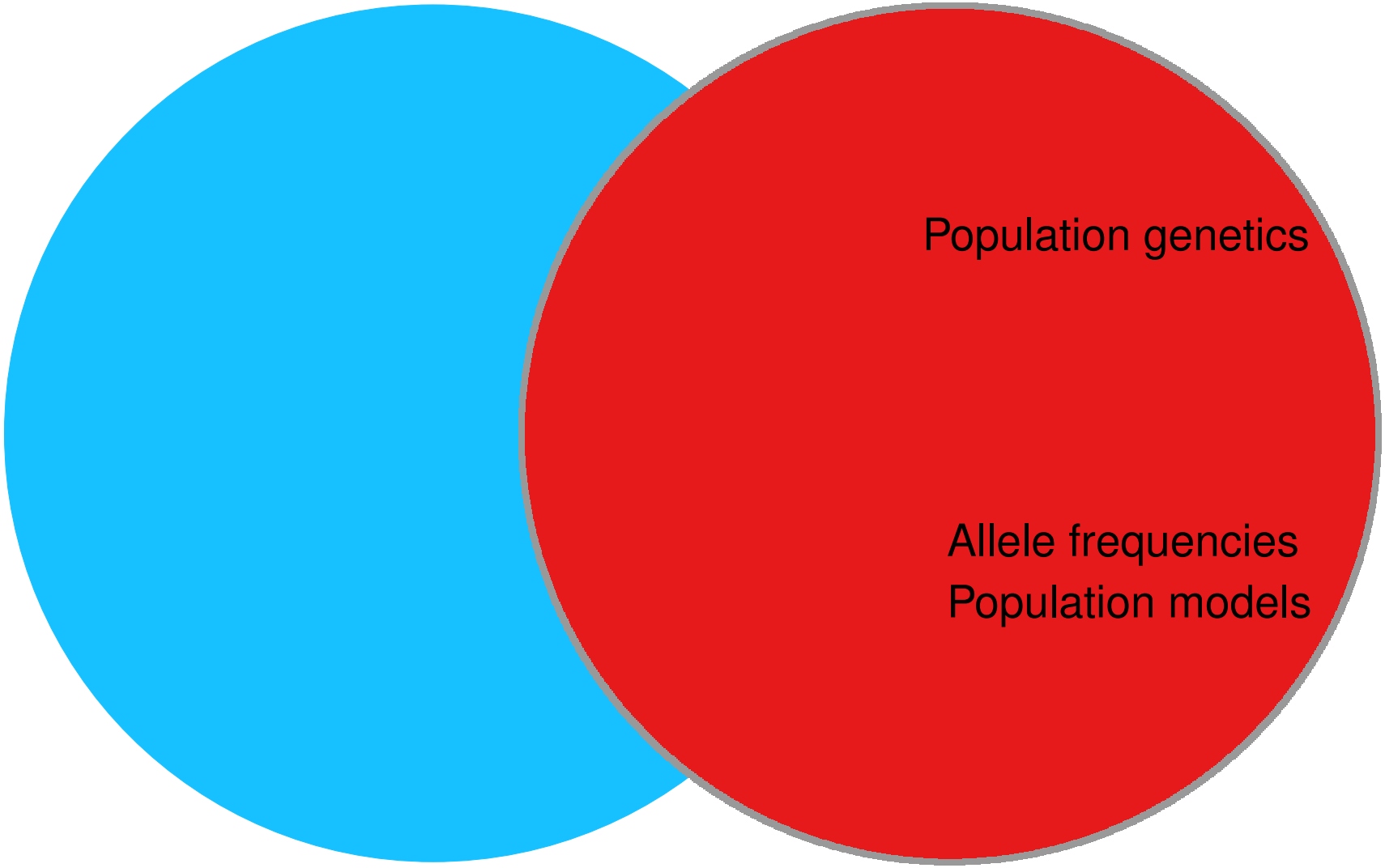
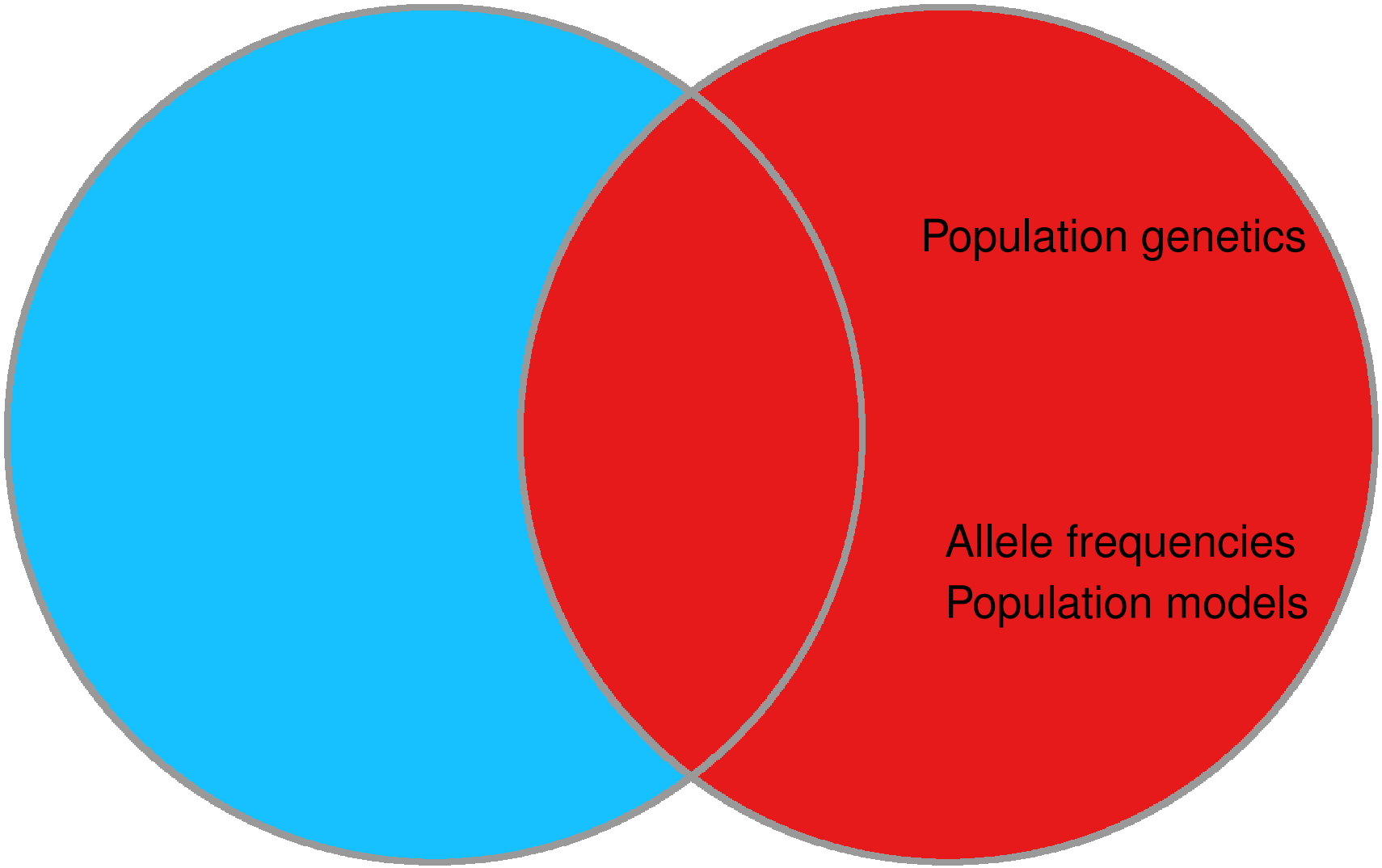
The big overview



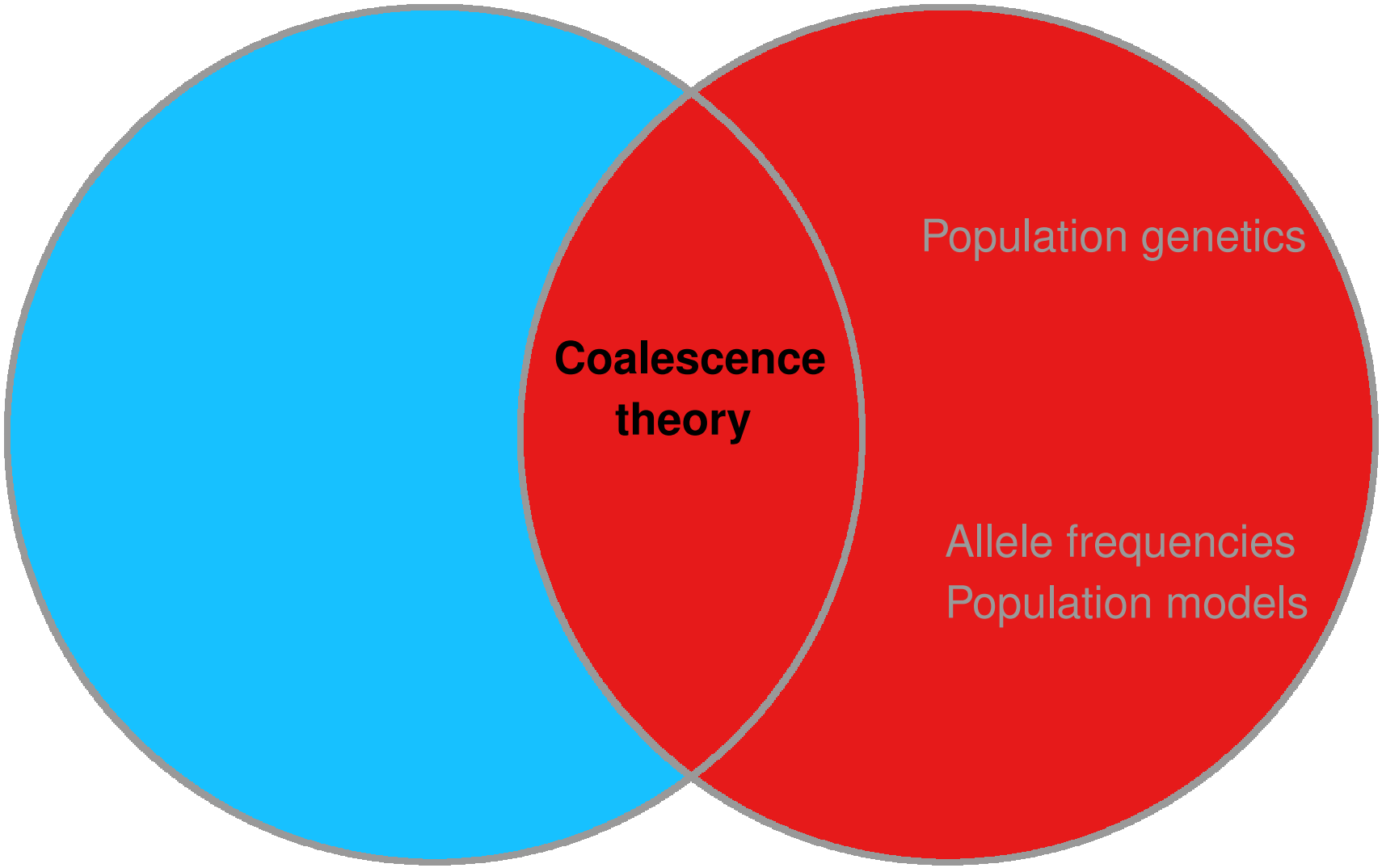
The big overview



The big overview



The big overview



Coalescence theory as a tool for population genetics

co•a•lesce |ˌkōəˈles|

verb [intrans.]

come together and form one mass or whole *: the puddles had* ***coalesced into*** *shallow streams | the separate details coalesce to form a single body of scientific thought.*

• [trans.] combine (elements) in a mass or whole *: to help coalesce the community, they established an office.*

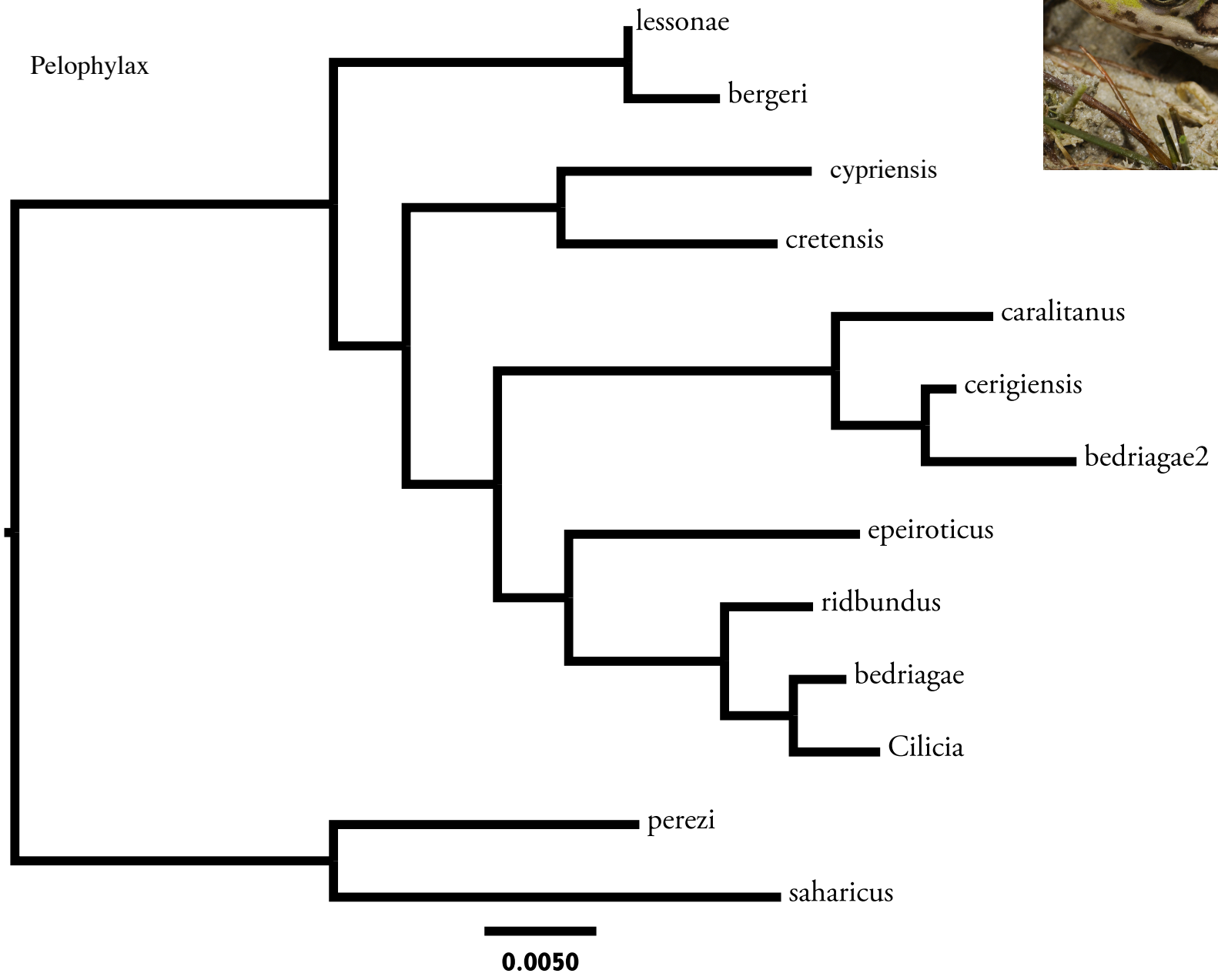
DERIVATIVES

co•a•les•cence |-ˈlesəns| noun

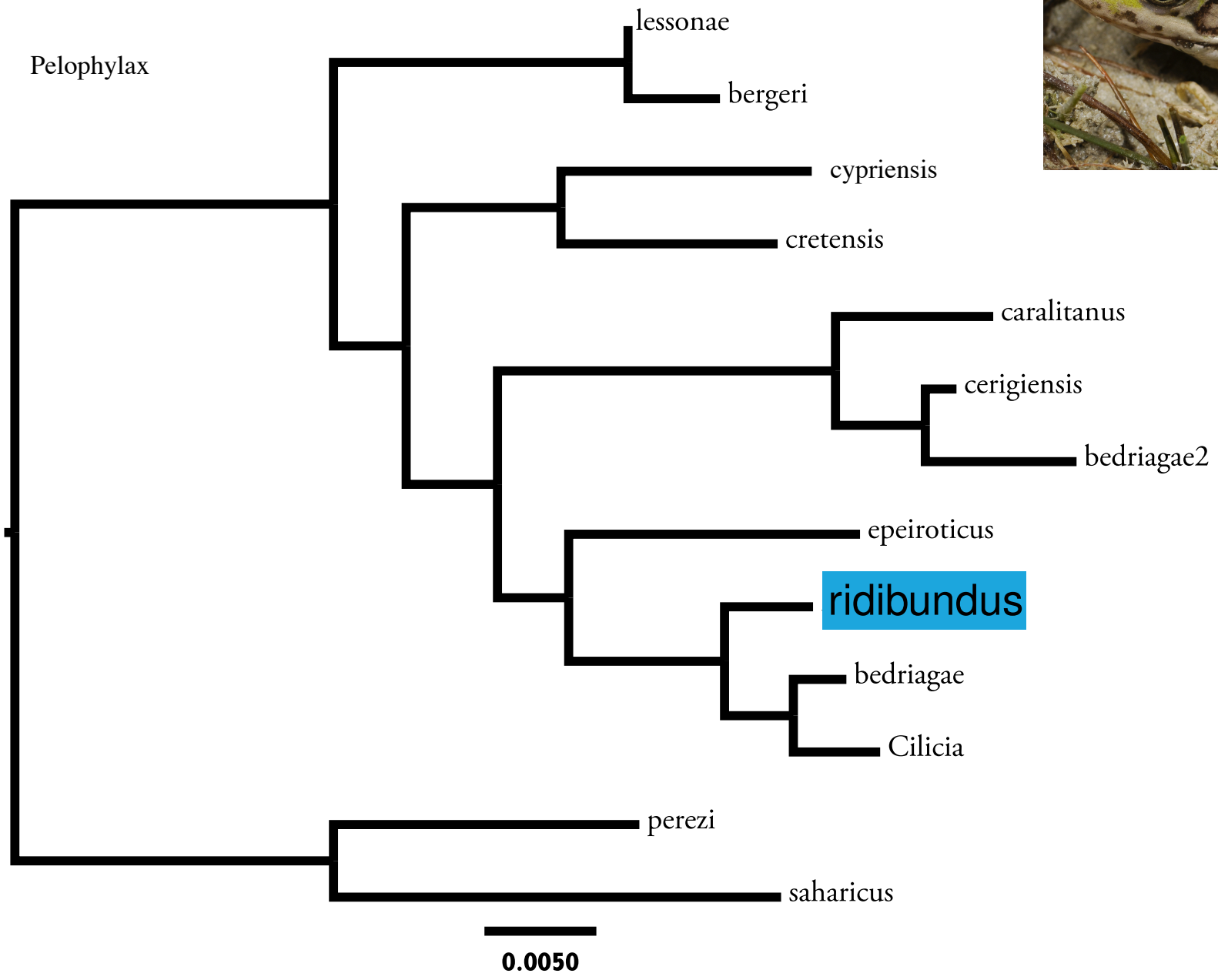
co•a•les•cent |-ˈlesənt| adjective

ORIGIN mid 16th cent. (in the sense [bring together, unite]): from Latin ***coalescere***, from ***co-*** (from ***cum 'with'***) + ***alescere 'grow up'*** (from ***alere 'nourish'***).

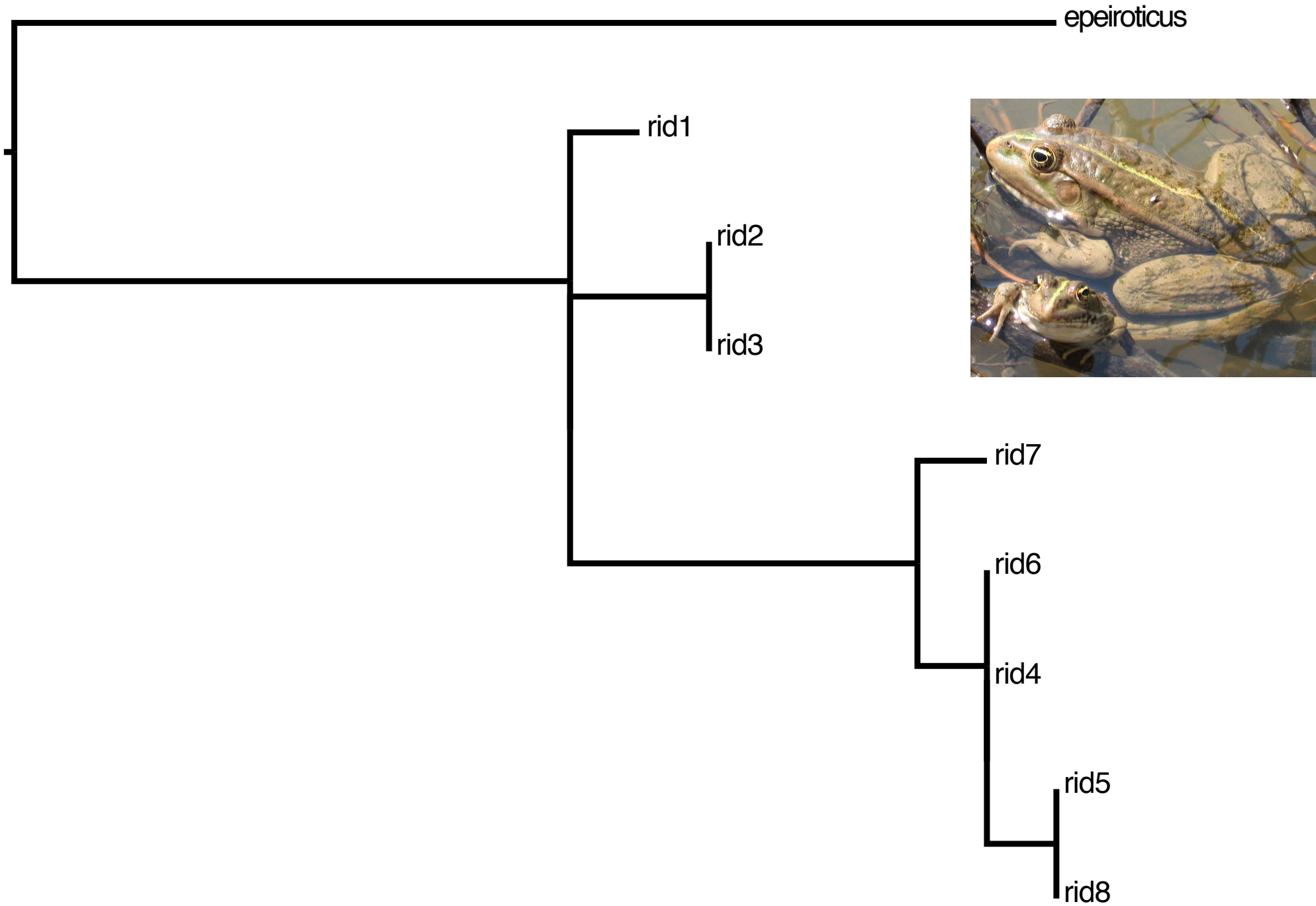
Species trees



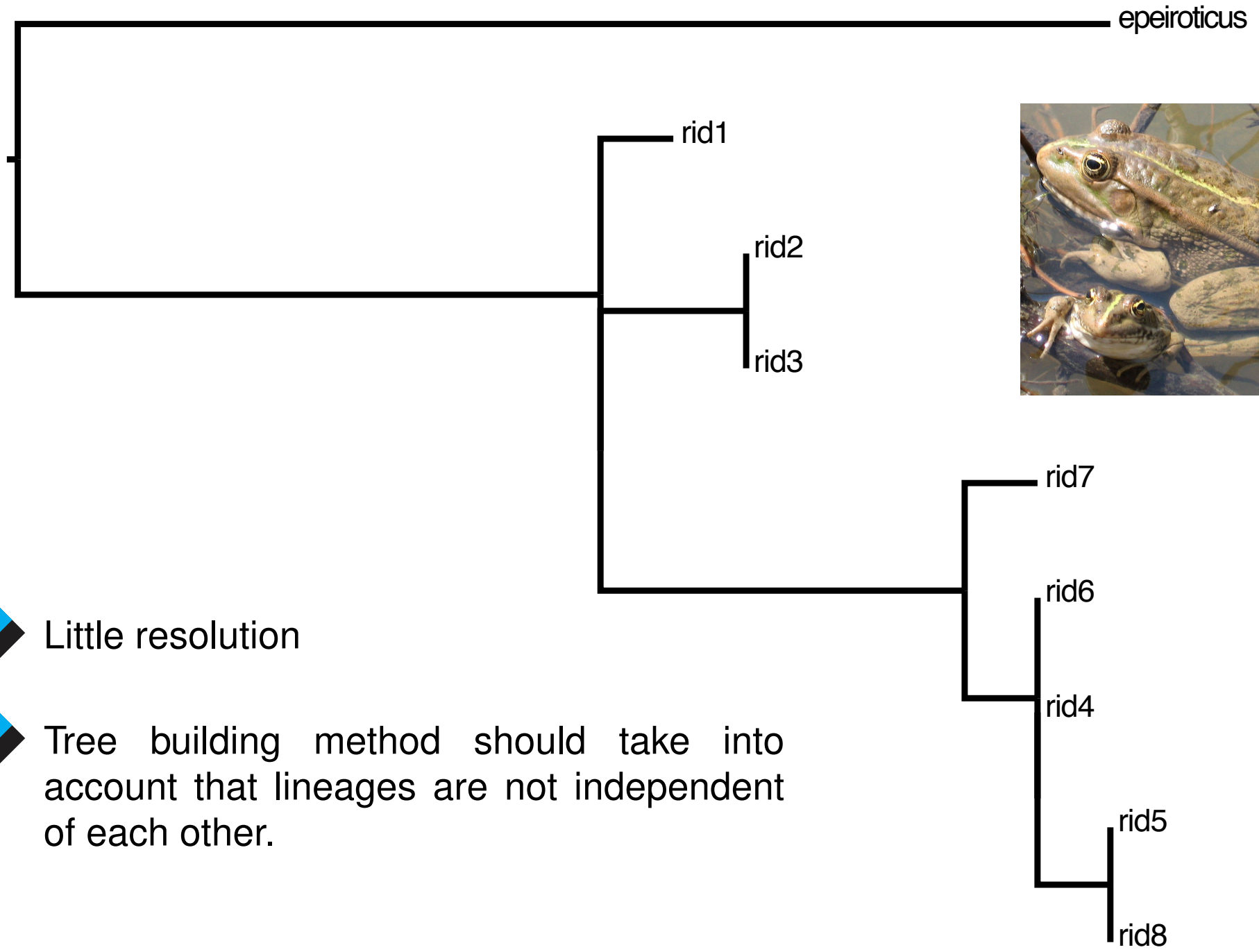
Species trees



Tree of individuals of same species

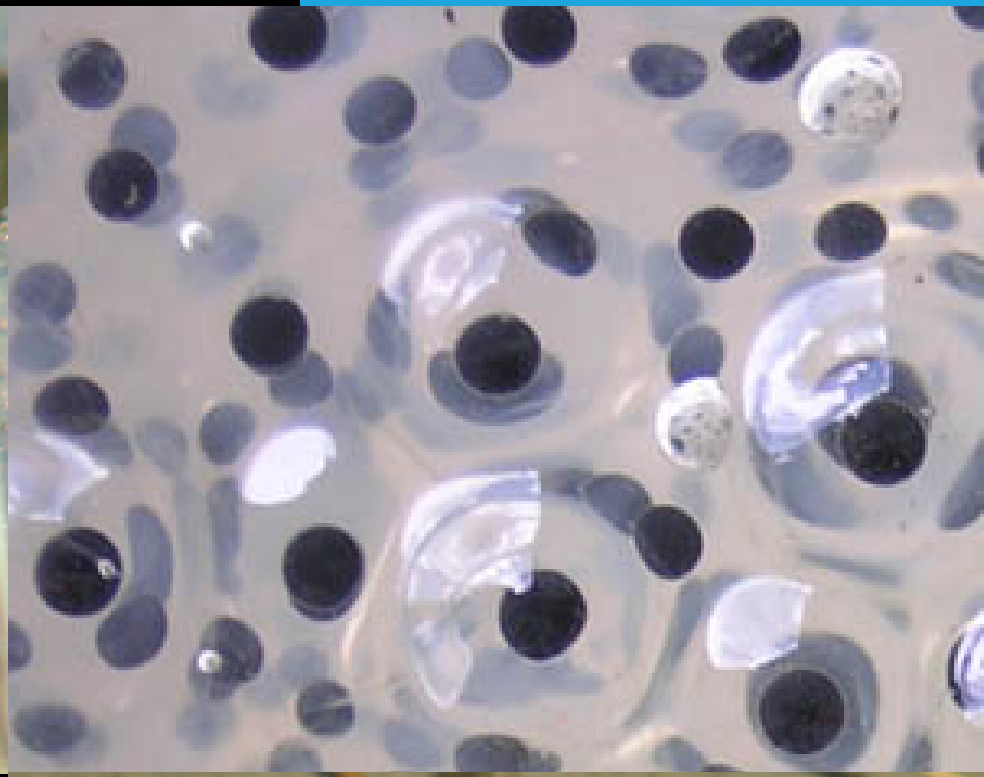


Tree of individuals of same species

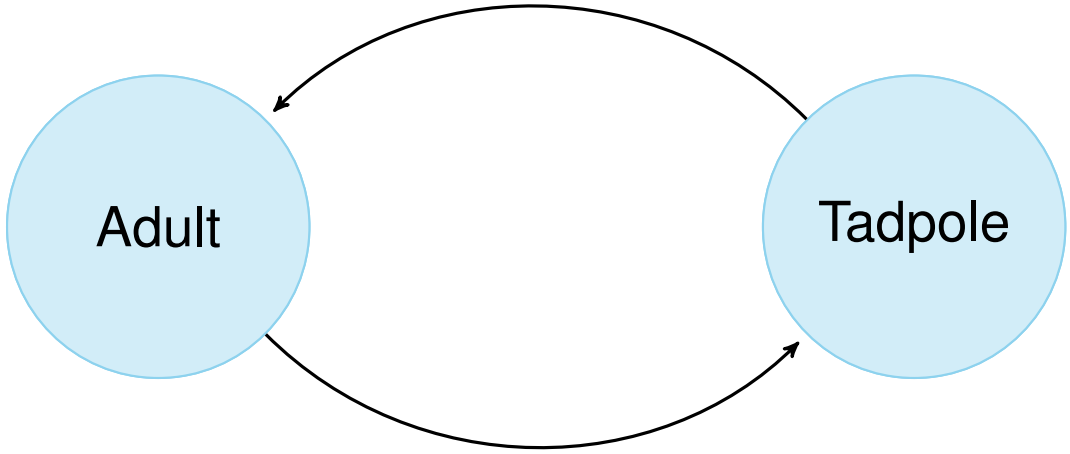


- Little resolution
- Tree building method should take into account that lineages are not independent of each other.

Interaction among individuals



Interaction among individuals



Wright-Fisher population model

- All individuals live one generation and get replaced by their offspring
- All have same chance to reproduce, all are equally fit
- The number of individuals in the population is constant

Population model

Past

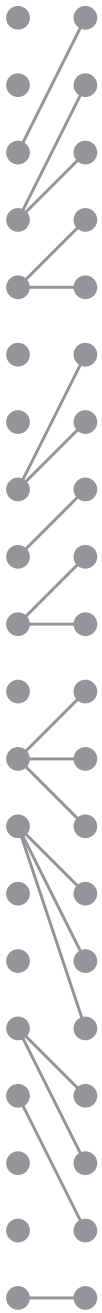
Present

Population model

Past

Present

Population model

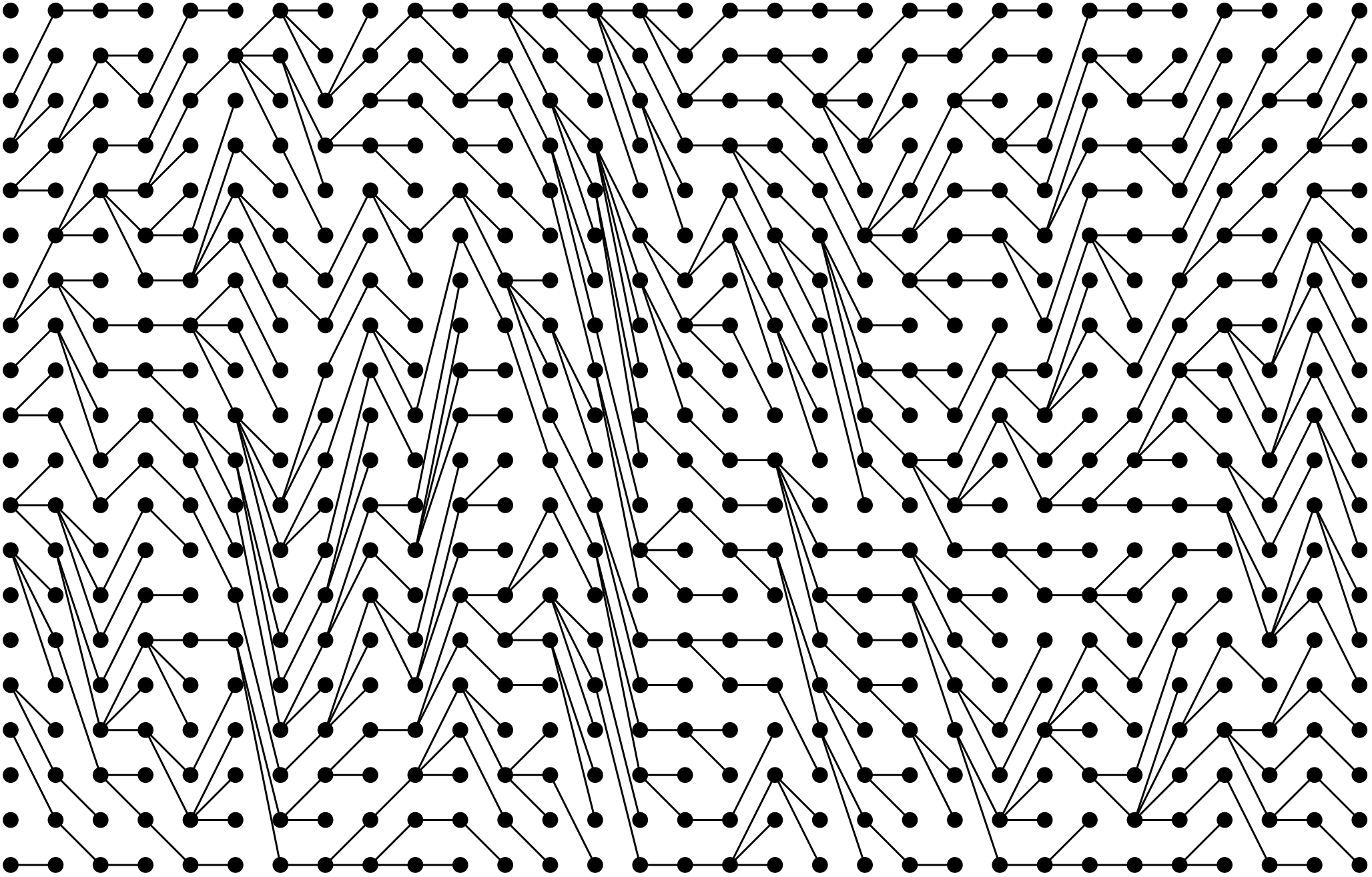


Past

Present

Population model

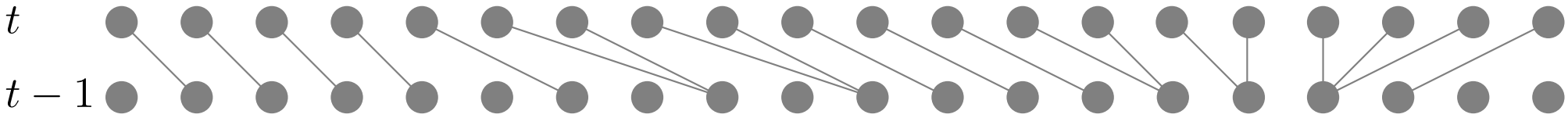
Wright-Fisher



Population model

Wright

Sewall Wright evaluated the probability that two randomly chosen individuals in generation t have a common ancestor in generation $t-1$. If we assume that there are $2N$ chromosomes then the probability of sharing a common ancestor in the last generation is



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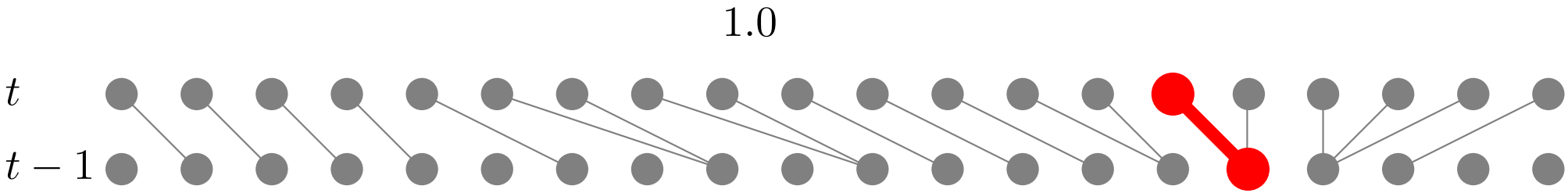
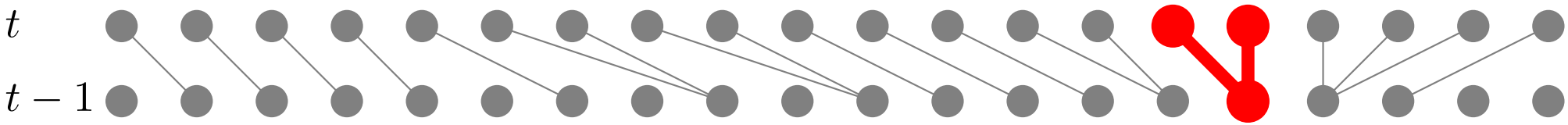


Population model

Wright

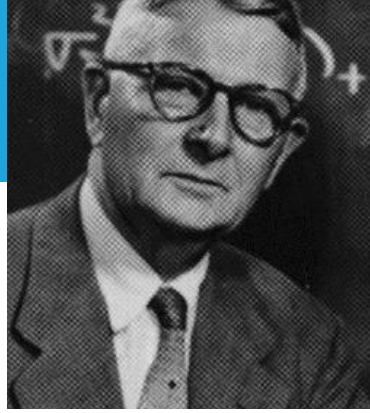
Sewall Wright evaluated the probability that two randomly chosen individuals in generation t have a common ancestor in generation $t-1$. If we assume that there are $2N$ chromosomes then the probability of sharing a common ancestor in the last generation is

$$1.0 \times \frac{1}{2N}$$



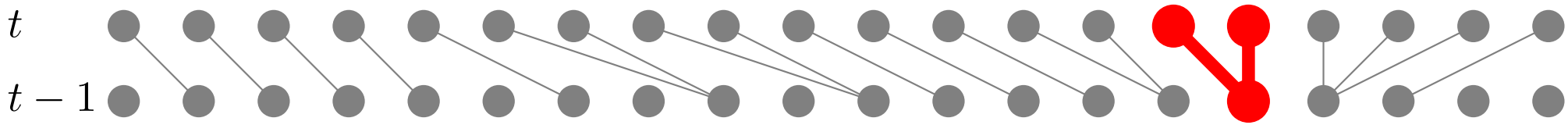
Population model

Wright



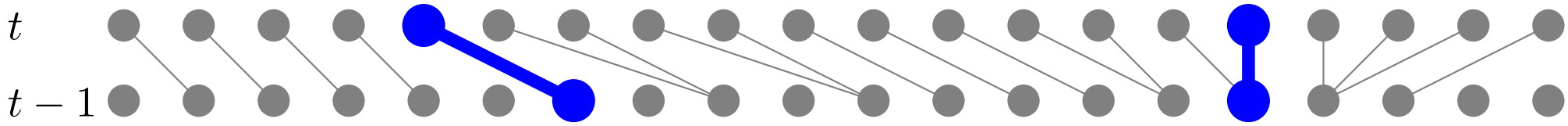
Sewall Wright evaluated the probability that two randomly chosen individuals in generation t have a common ancestor in generation $t-1$. If we assume that there are $2N$ chromosomes then the probability of sharing a common ancestor in last generation is

$$\frac{1}{2N}$$



The probability that two randomly picked chromosome do not have a common ancestor is

$$1 - \frac{1}{2N}$$



Population model

Wright

If we know the genealogy of the two individuals then we can calculate the probability as

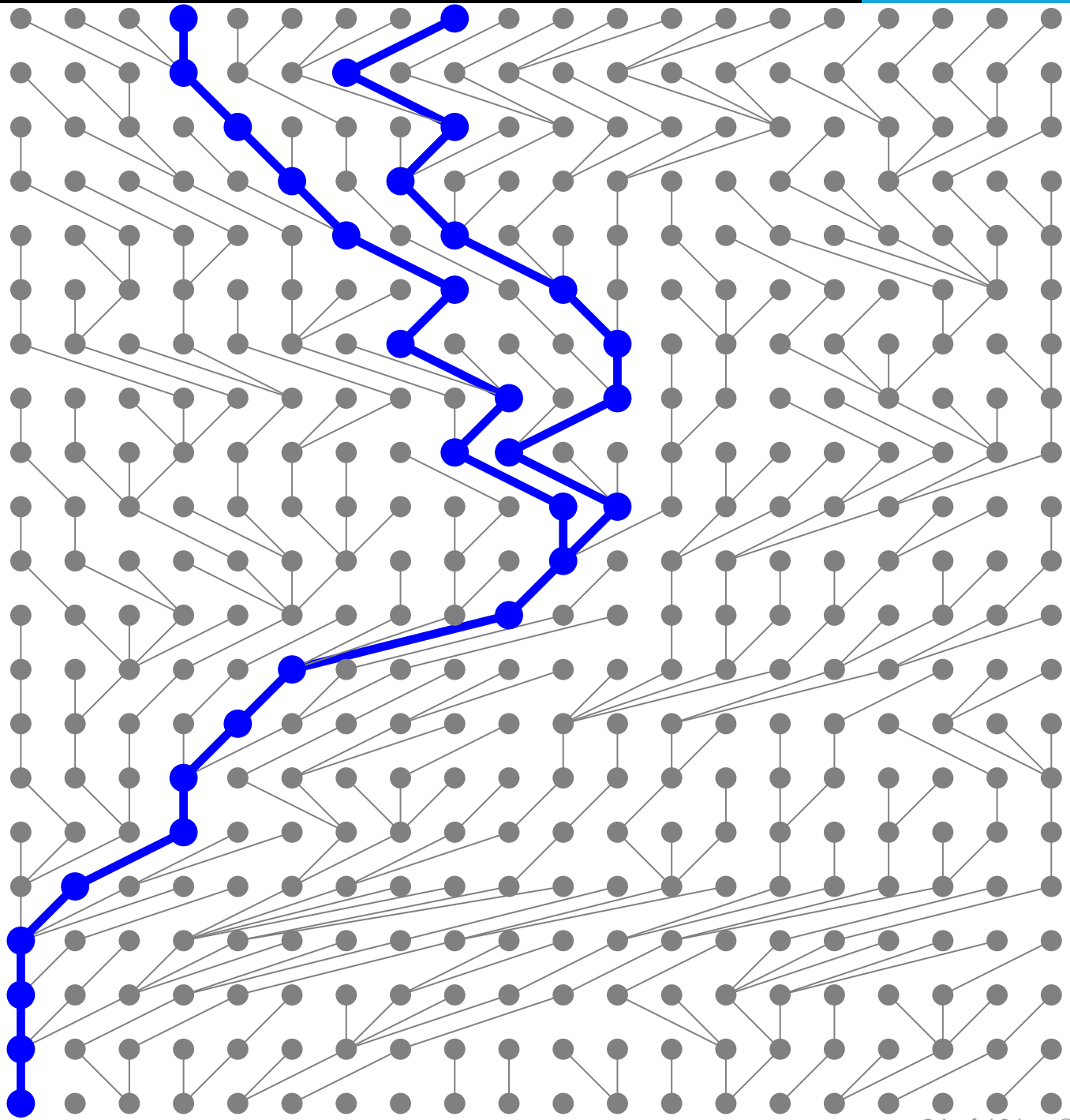
$$\mathrm{P}(\tau|N) = \left(1 - \frac{1}{2N}\right)^{\tau} \left(\frac{1}{2N}\right)$$

where τ is the number of generations with no coalescence. This formula is the Geometric Distribution and we can calculate the expectation of the waiting time until two random individuals coalesce:

$$\mathbb{E}(\tau) = 2N$$

Population model

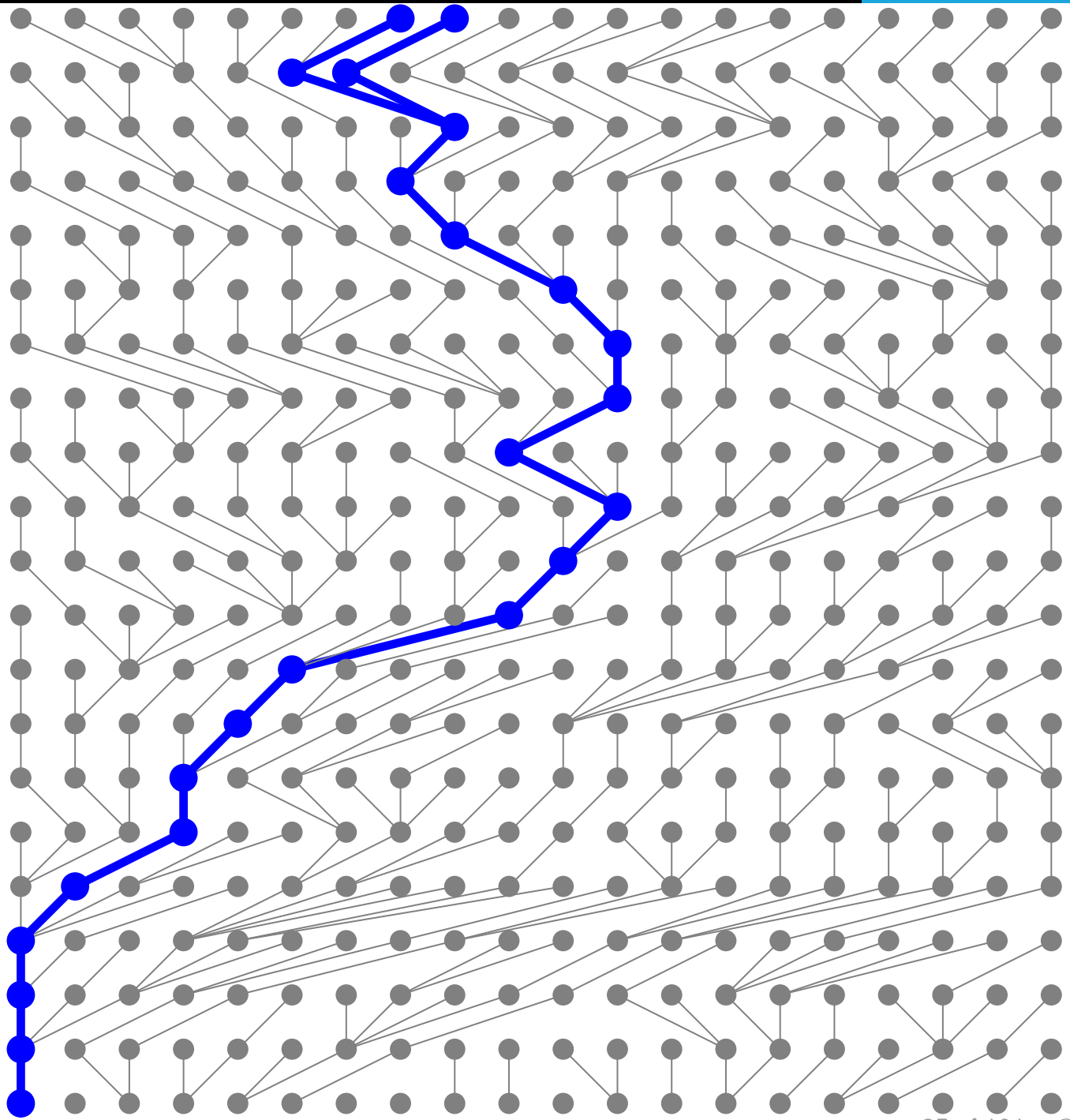
Present



Past

Population model

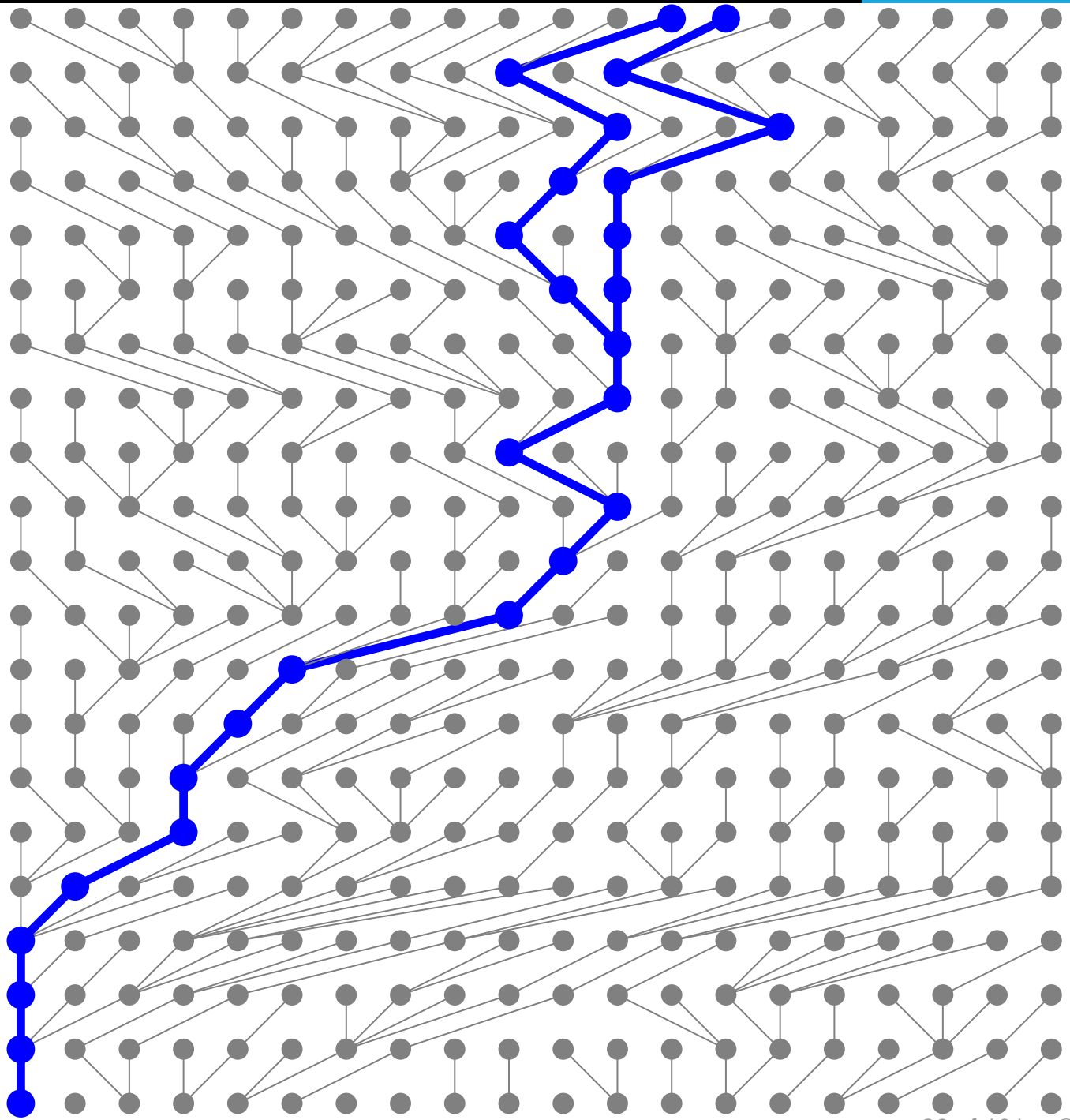
Present



Past

Population model

Present

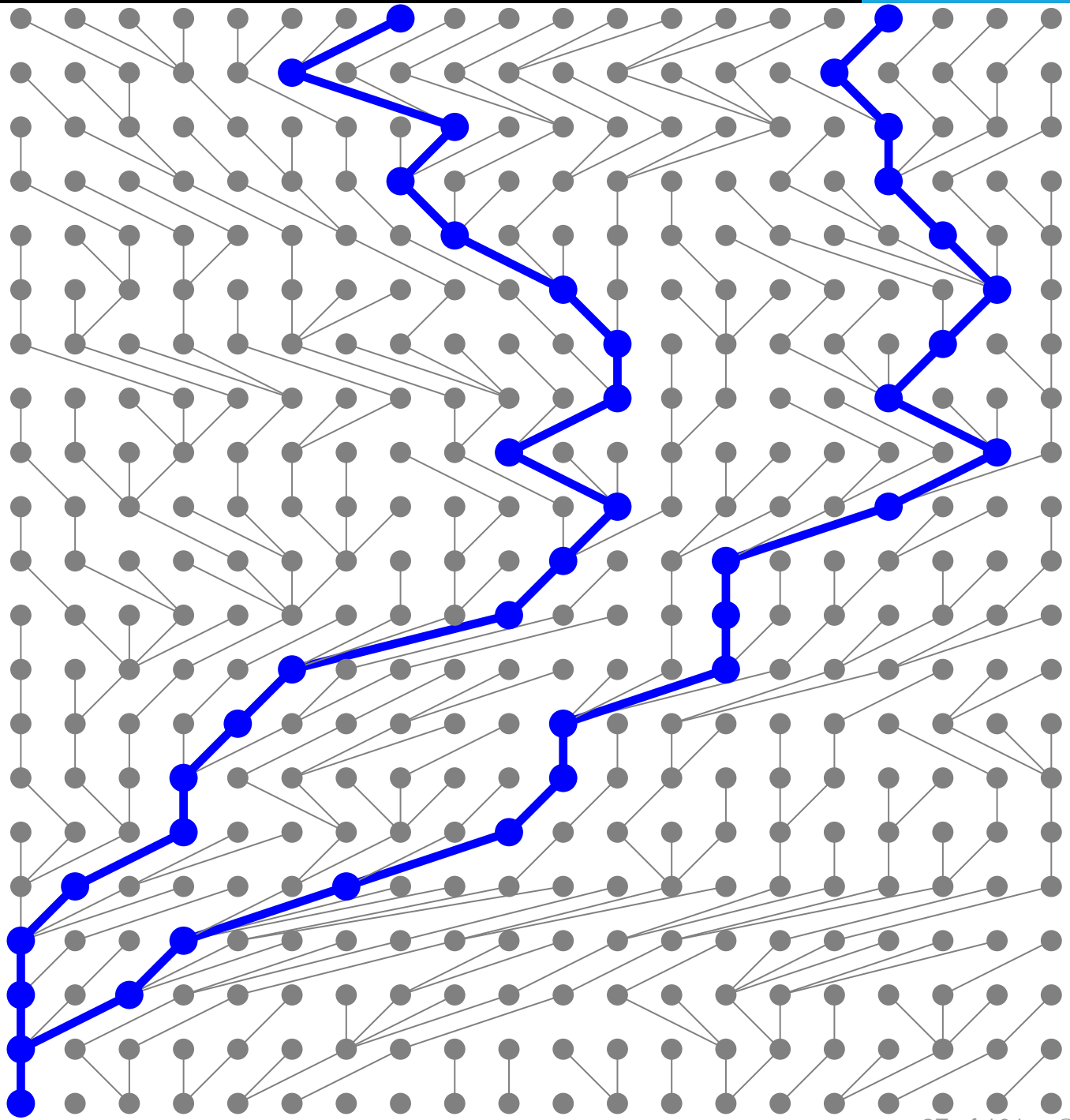


Past

Population model

Wright-Fisher

Present

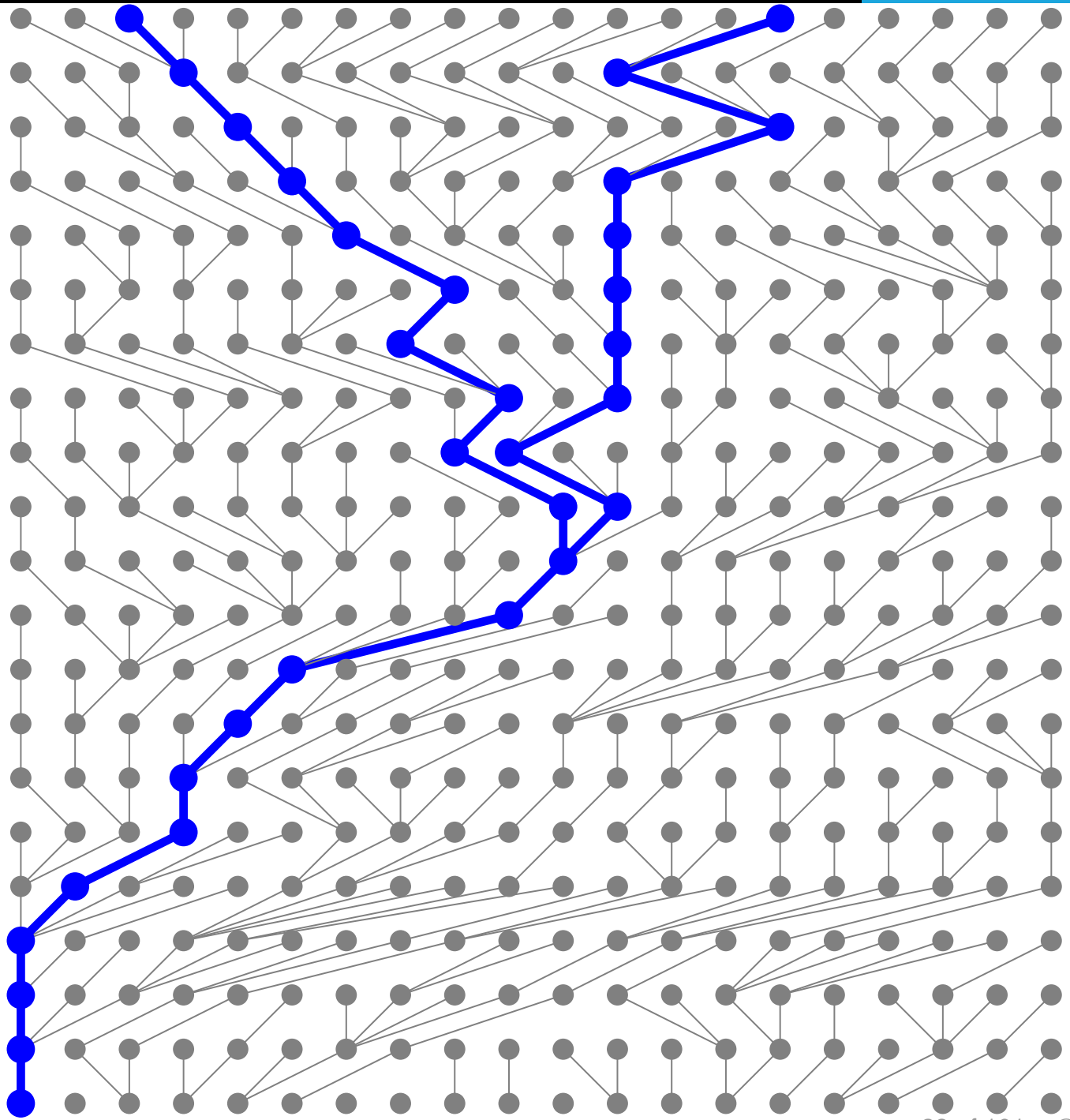


Past

Population model

Wright-Fisher

Present

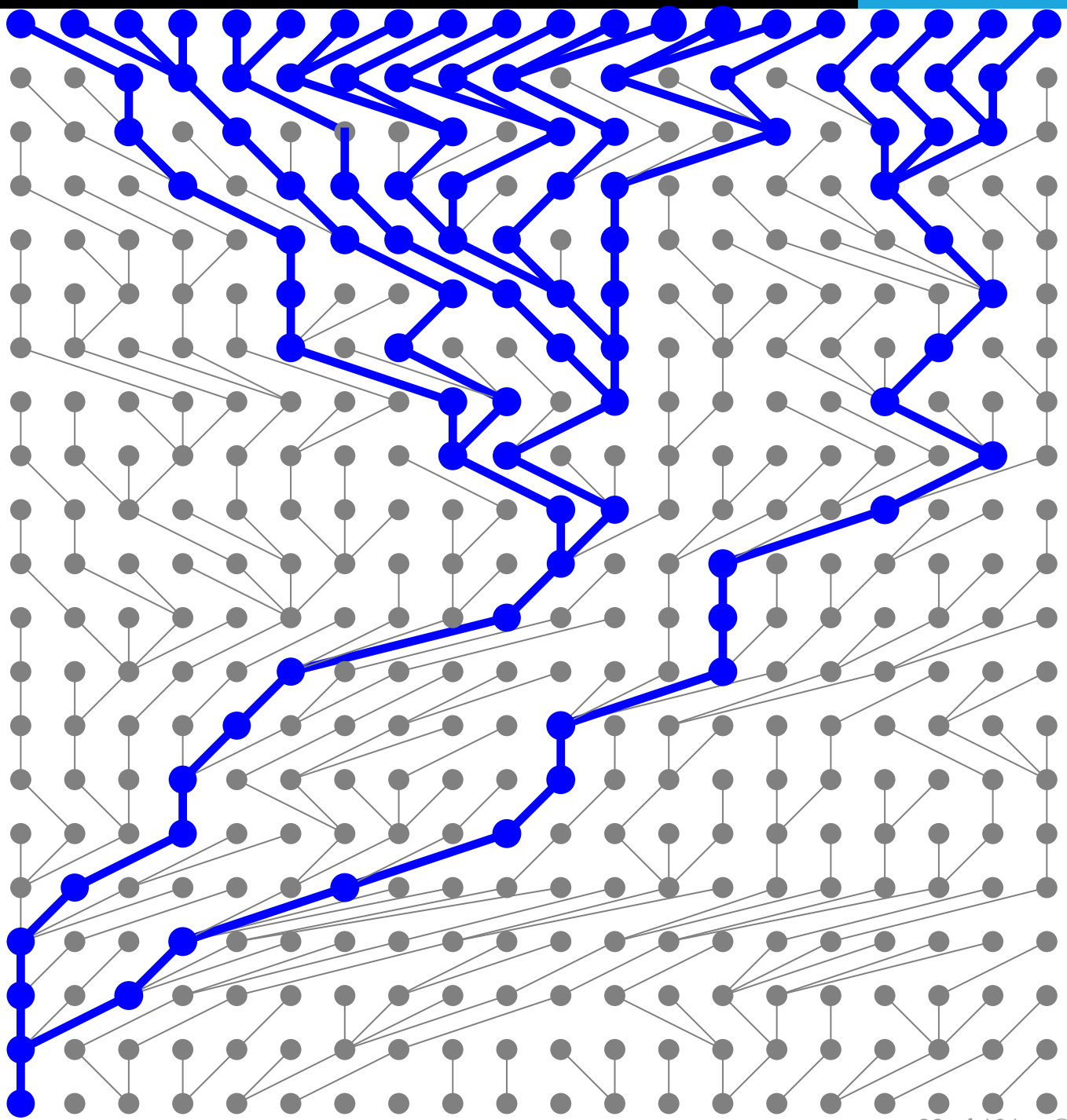


Past

Population model

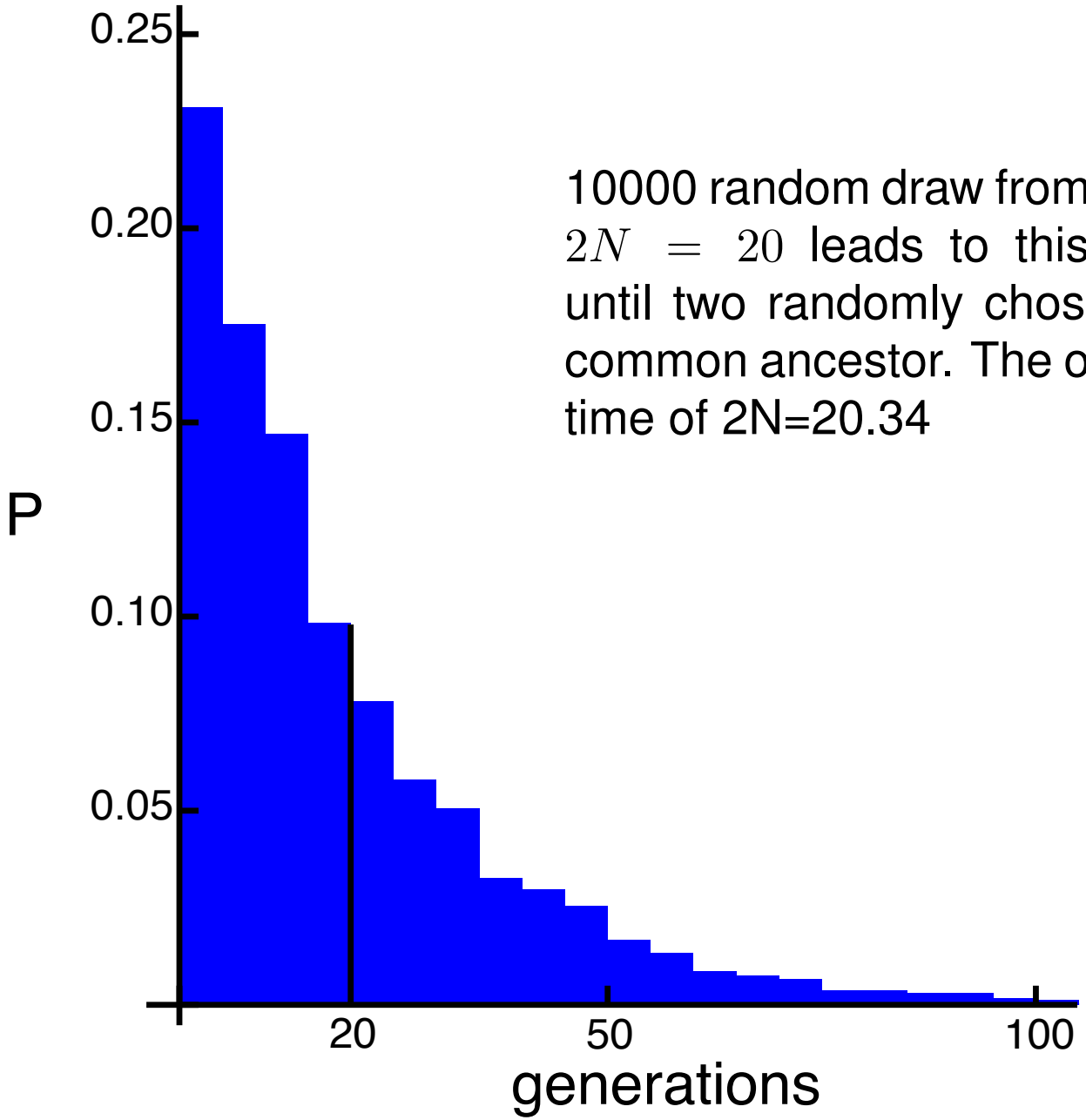
Wright-Fisher

Present



Past

Probability Distribution



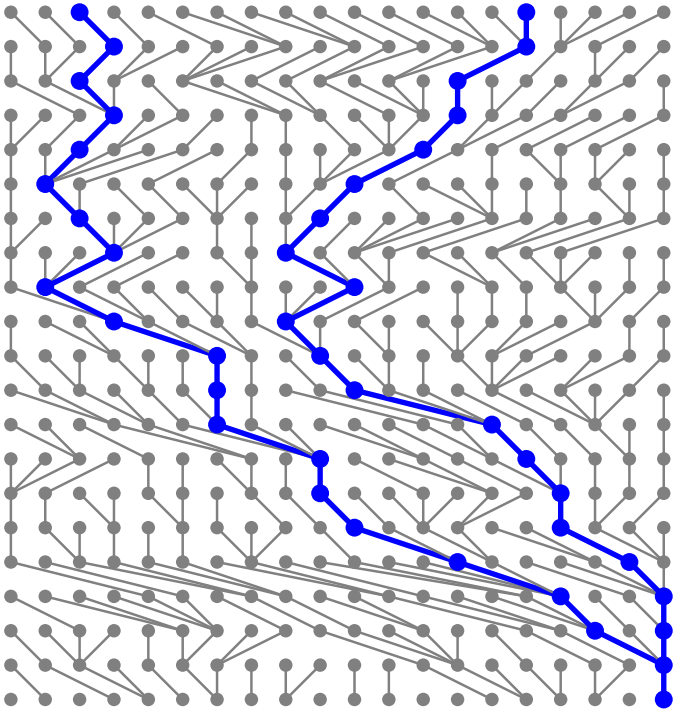
10000 random draw from a population with size $2N = 20$ leads to this distribution of times until two randomly chosen individuals have a common ancestor. The observed mean waiting time of 2N=20.34

Observations

- For the time of coalescence in a sample of TWO , we will wait on average $2N$ generations assuming it is a Wright-Fisher population
- The model assumes that the generations are discrete and non-overlapping
- Real populations do not necessarily behave like a Wright-Fisher (the *'ideal' population)*
- *We assume that calculation using Wright-Fisher populations can be extrapolated to real populations.*

Other population models

Wright-Fisher

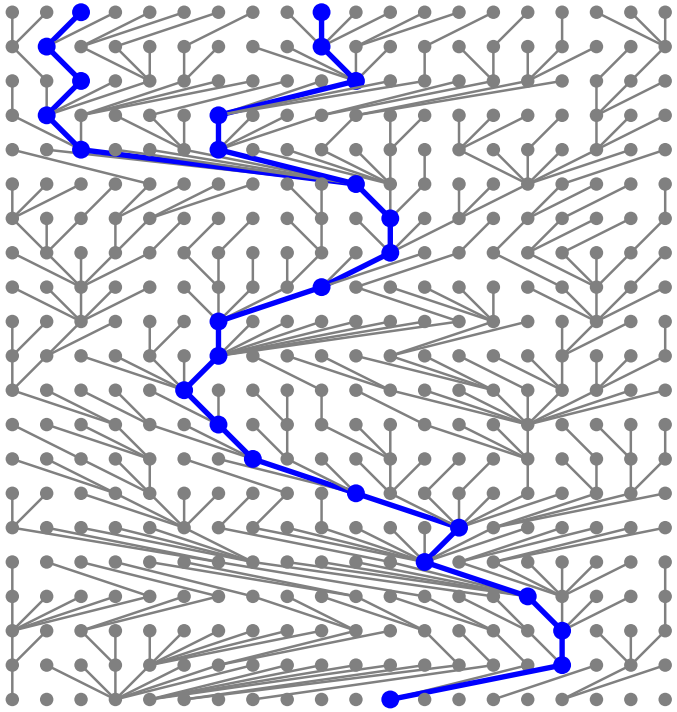


$\sigma^2_{\text{offspring}} \simeq 1$

$\mathbb{E}(\tau) = 2N$

generation time $g = 1$

Canning

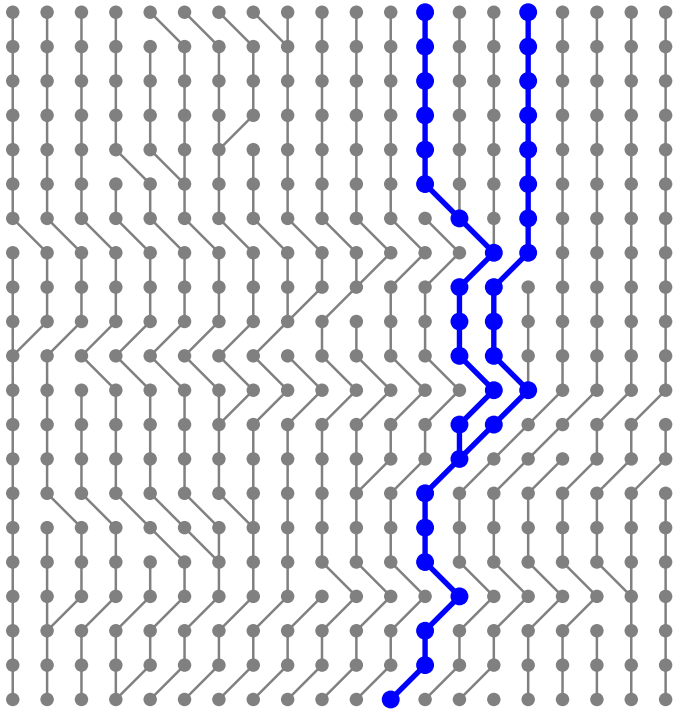


$\sigma^2_{\text{offspring}} = x$

$\mathbb{E}(\tau) = 2N/x$

$g = 1$

Moran



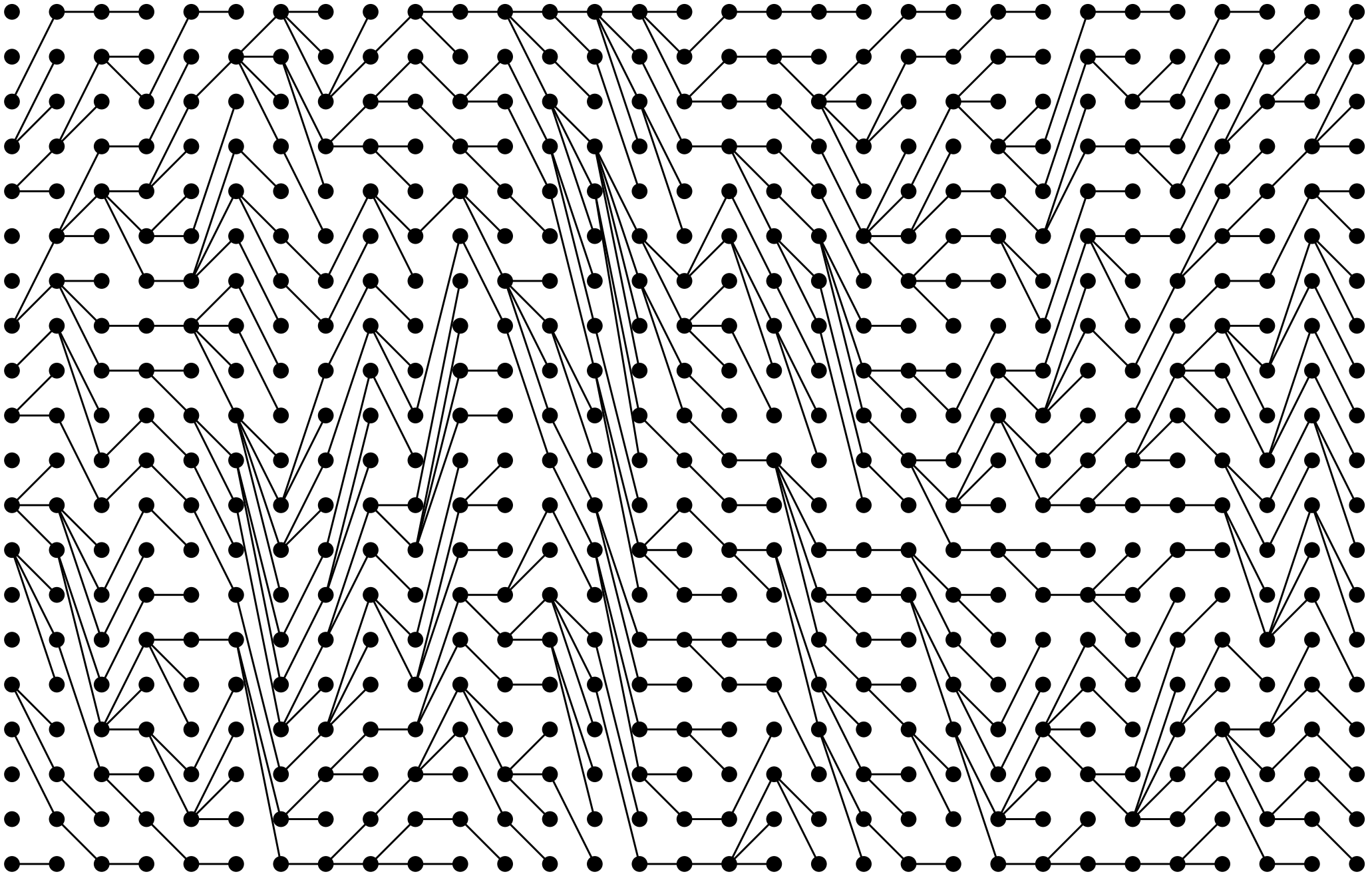
$\sigma^2_{\text{offspring}} = \frac{2}{2N}$

$\mathbb{E}(\tau) = \frac{1}{2}(2N)^2$

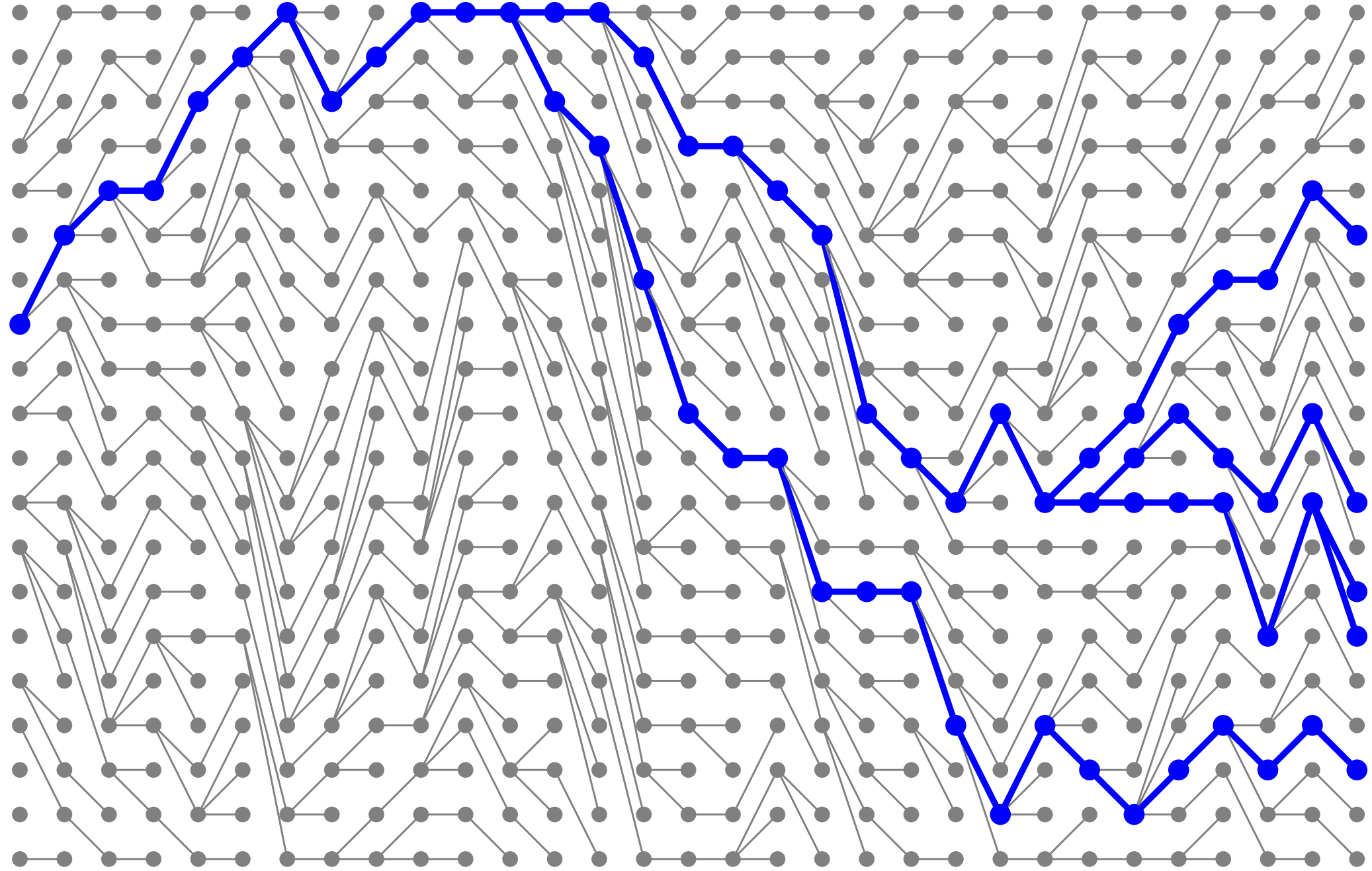
$g = 2N$

You can generate graphs like this using the python program *popsim* (check out my faculty page for the link)

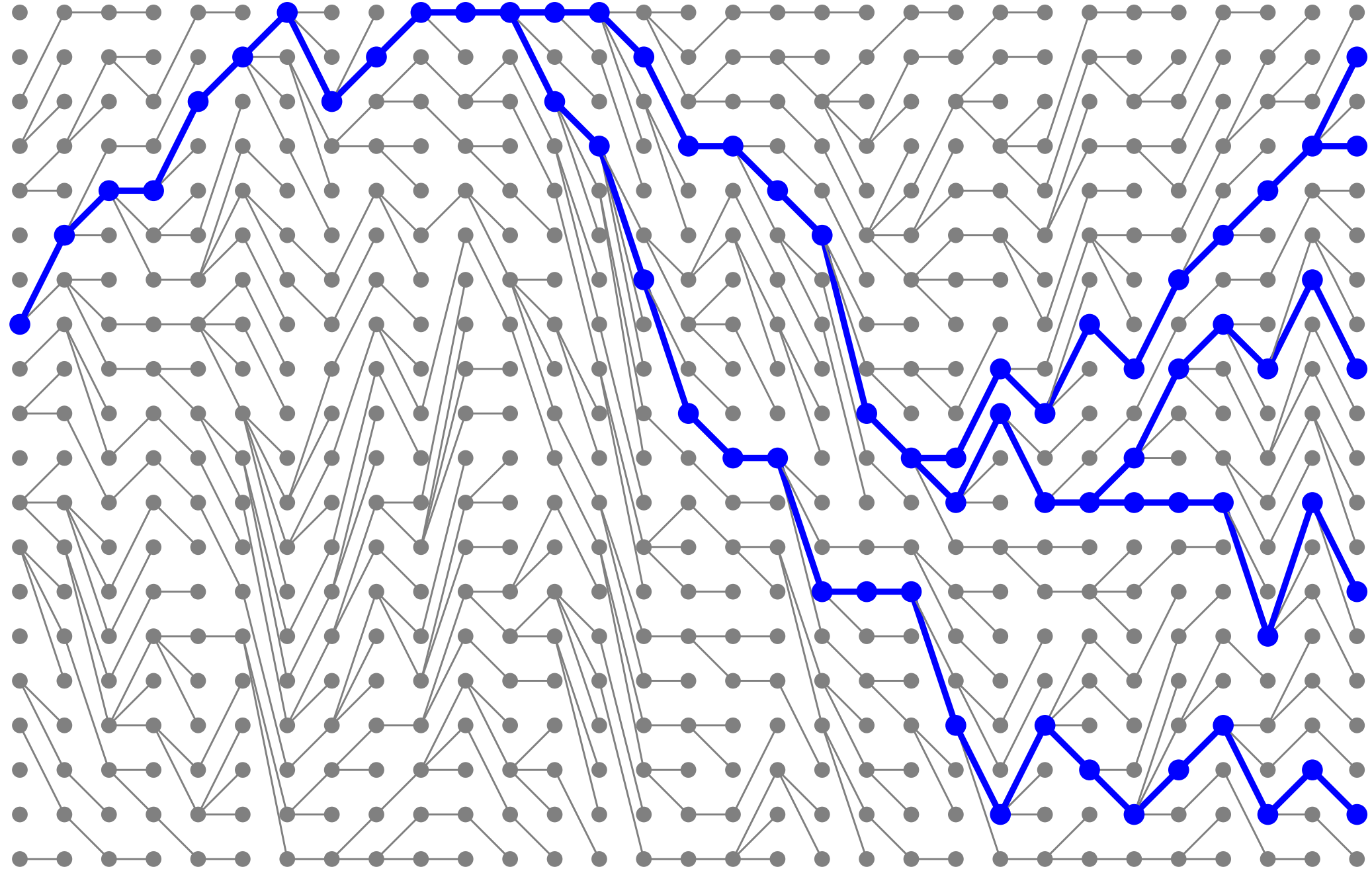
Sample larger than TWO



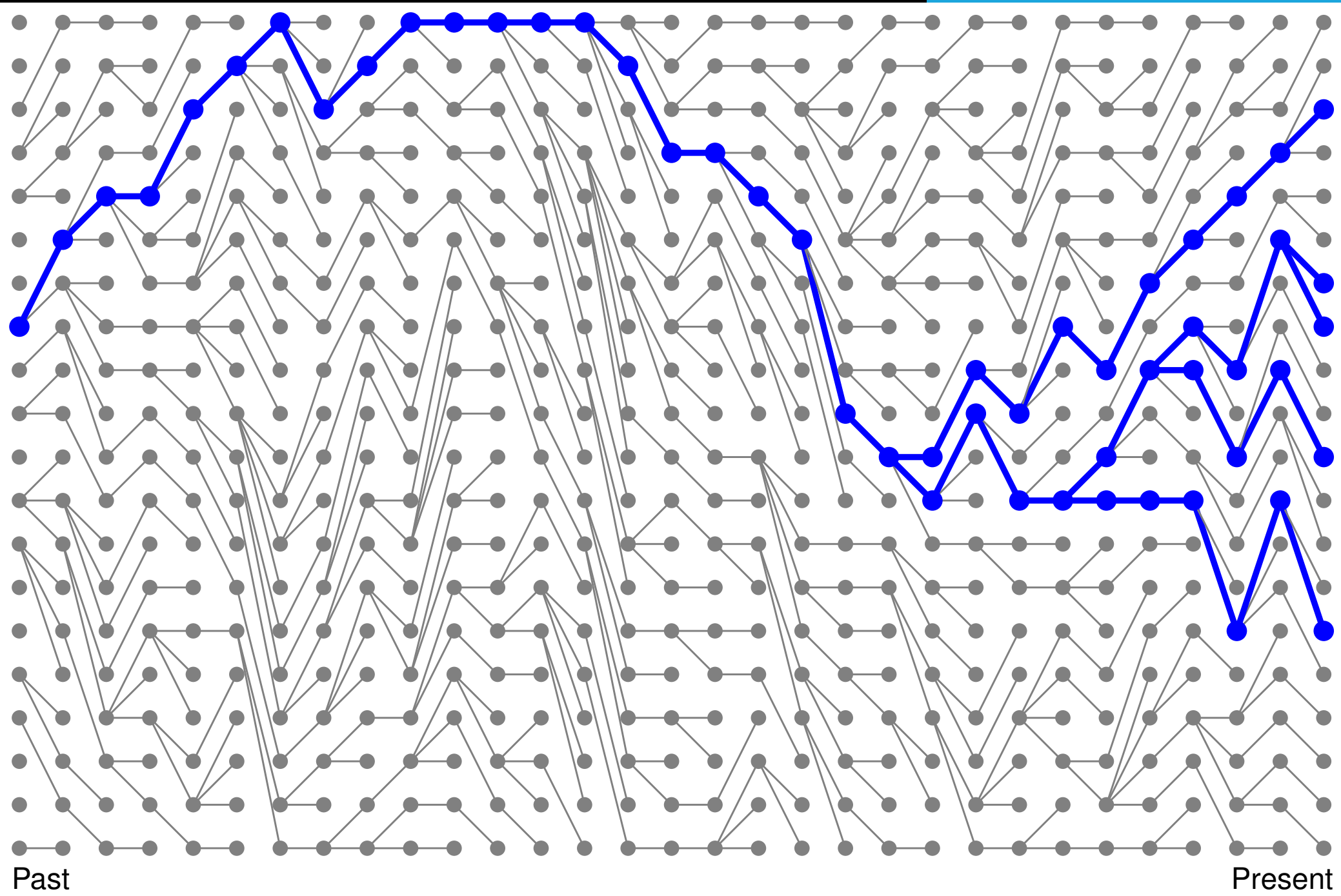
Sample larger than TWO



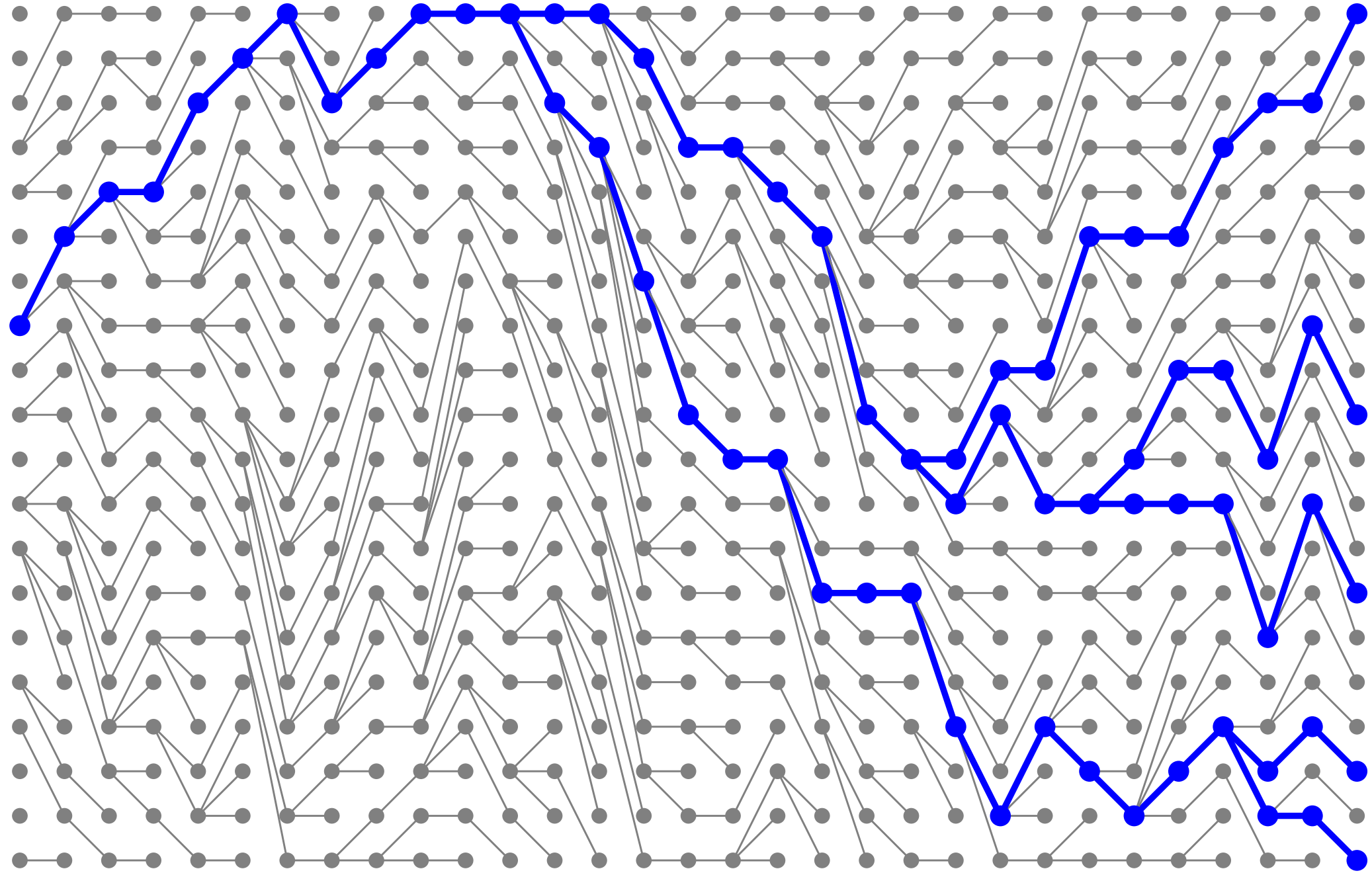
Sample larger than TWO



Sample larger than TWO



Sample larger than TWO



Samples larger than two

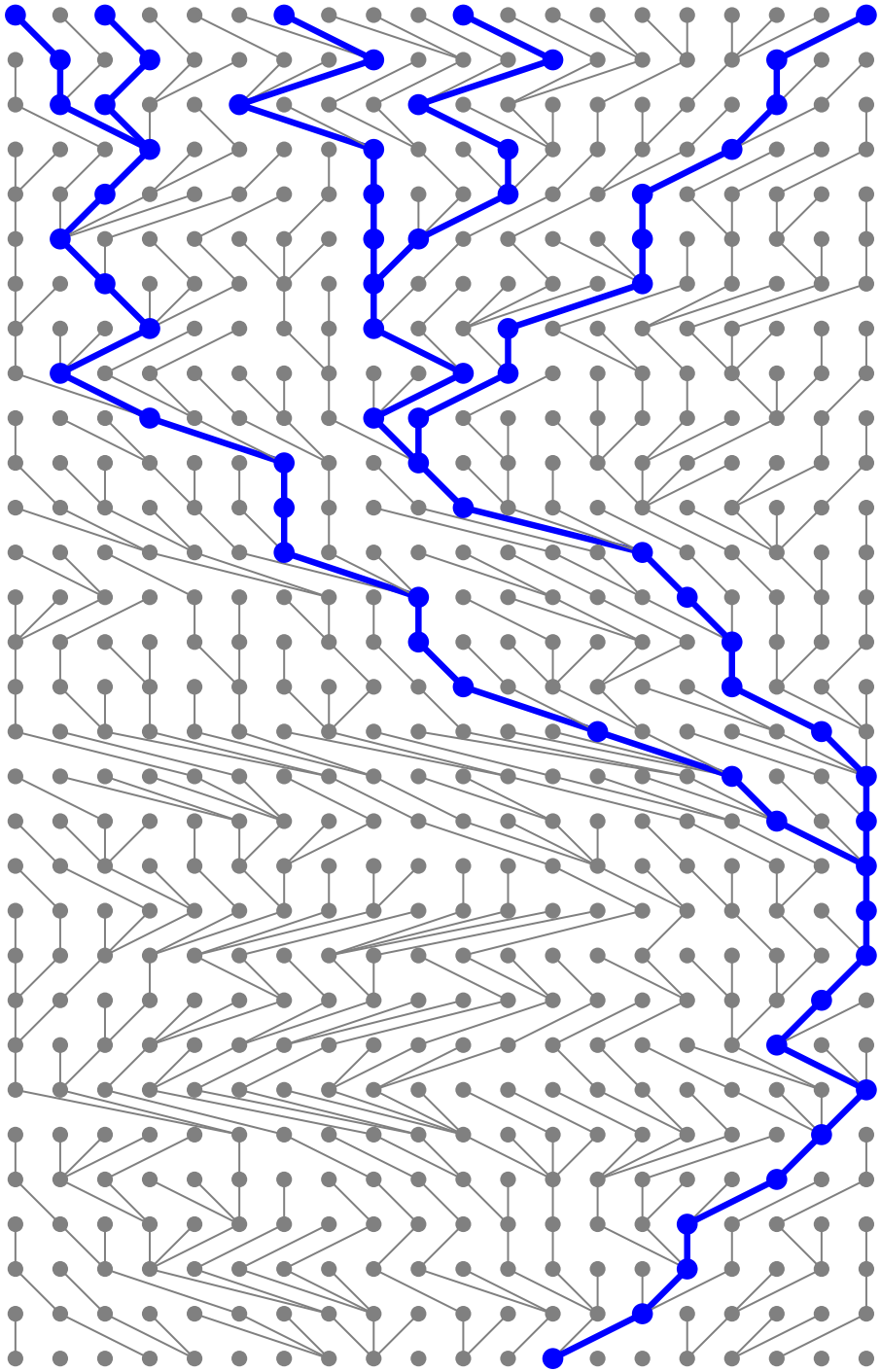
Sir J. F. C. Kingman described in 1982 the n-coalecent. He showed the behavior of a sample of size n, and its probability structure looking backwards in time.

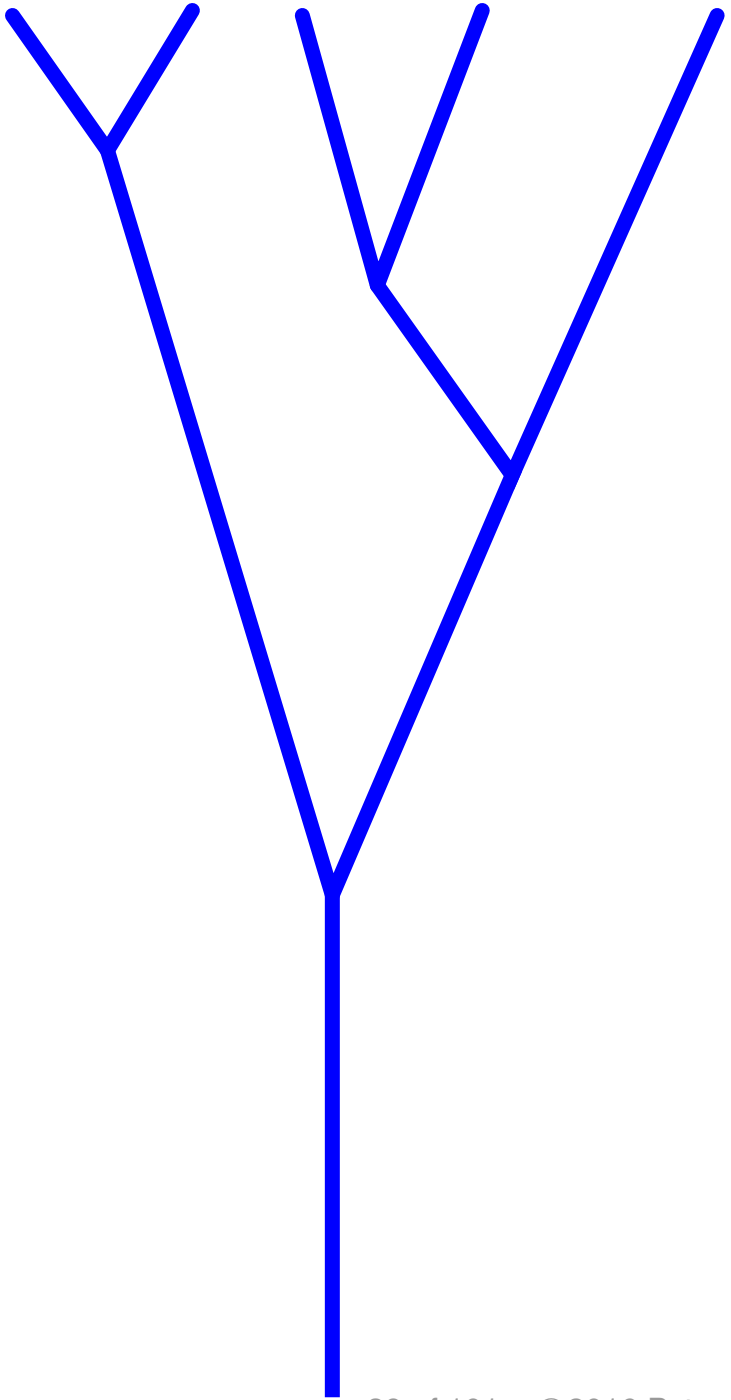
General findings:

$$\text{coalescence rate} = \binom{n}{2} = \frac{n(n-1)}{2}$$

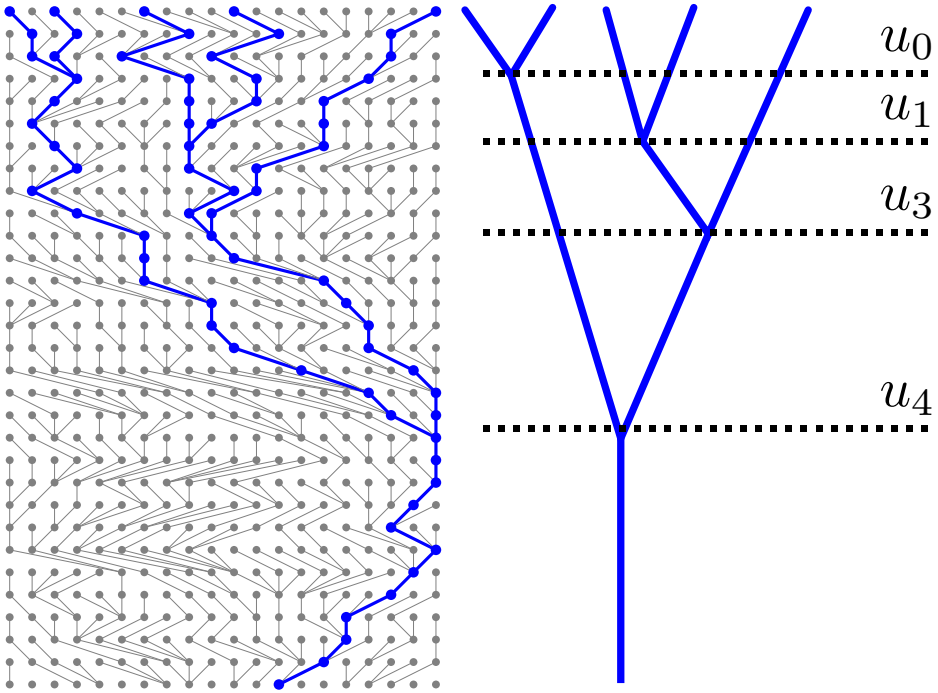
Once a coalescence happened n is reduced to $n-1$ because two lineages merged into one. He then imposed a continuous approximation of the Canning's exchangeable model to get results.

Samples larger than two



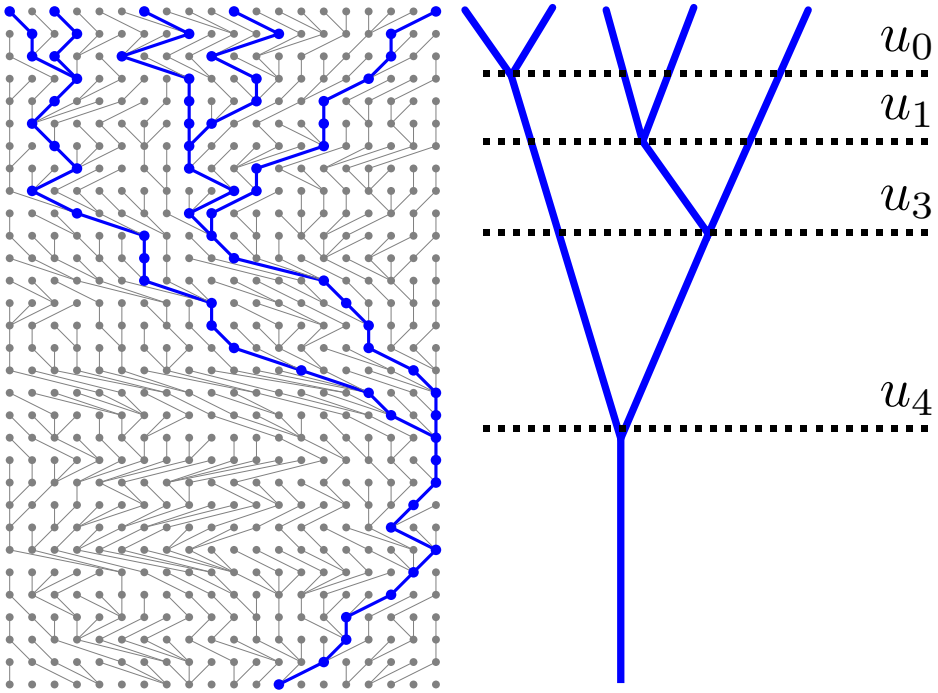


Samples larger than two



Looking backward in time, the first coalescence between two random individuals is the result of a waiting process that depends on the sample n and the total population size N.

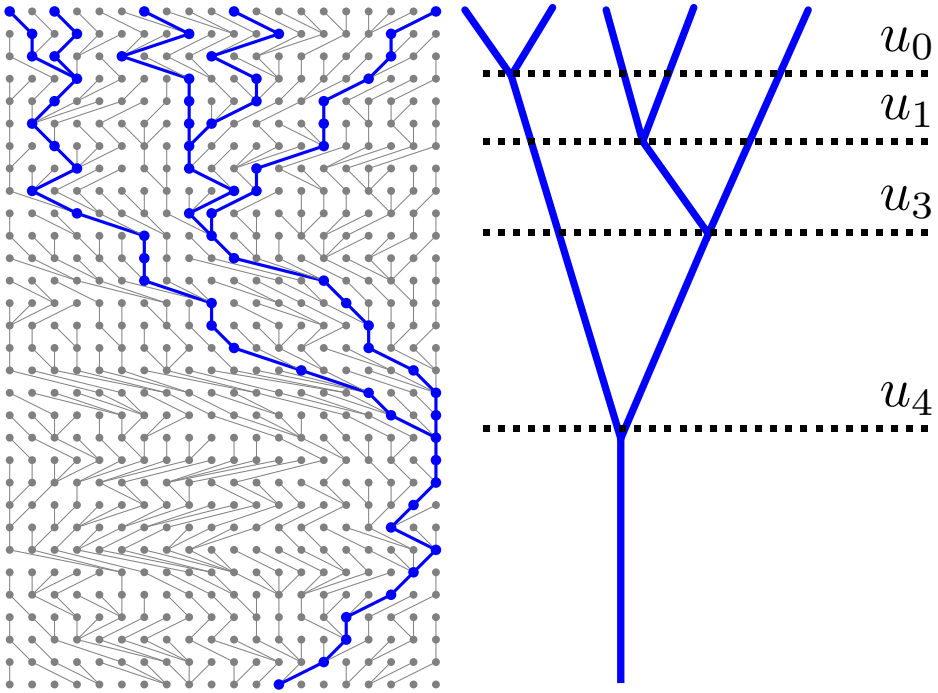
Samples larger than two



Looking backward in time, the first coalescence between two random individuals is the result of a waiting process that depends on the sample n and the total population size N.

Using Kingman's coalescence rate and imposing a time scale we can approximate the process with a exponential distribution:

Samples larger than two



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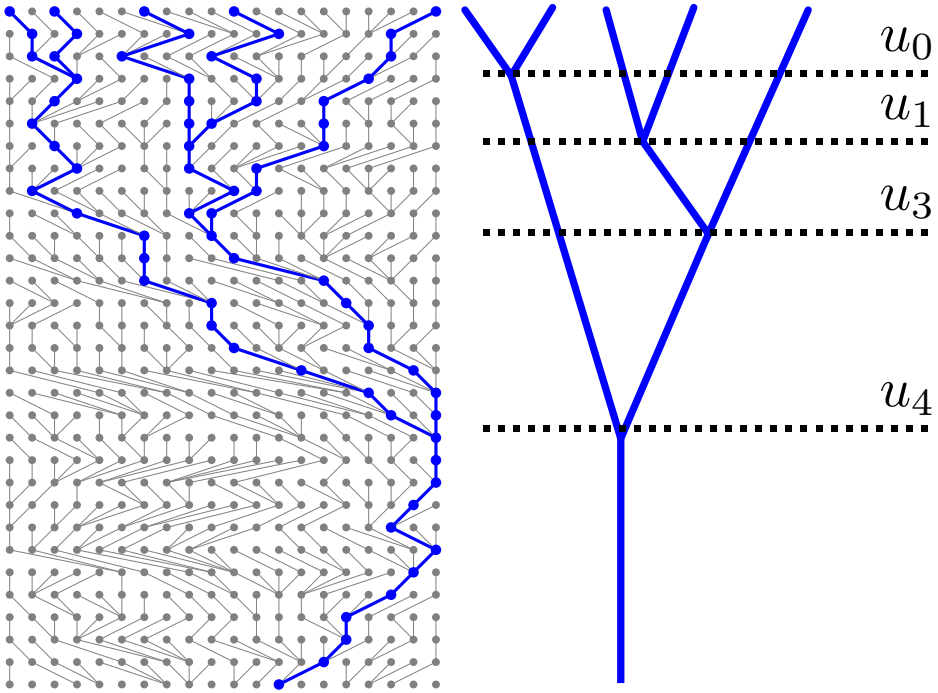
Using Kingman's coalescence rate and imposing a time scale we can approximate the process with an exponential distribution:

$$\mathrm{P}(u_j|N) = e^{-u_j\lambda}\lambda$$

with the scaled coalescence rate

$$\lambda = \binom{k}{2}\frac{1}{2N} \times \mathrm{Prob}(\text{others do not coalesce})$$

Samples larger than two



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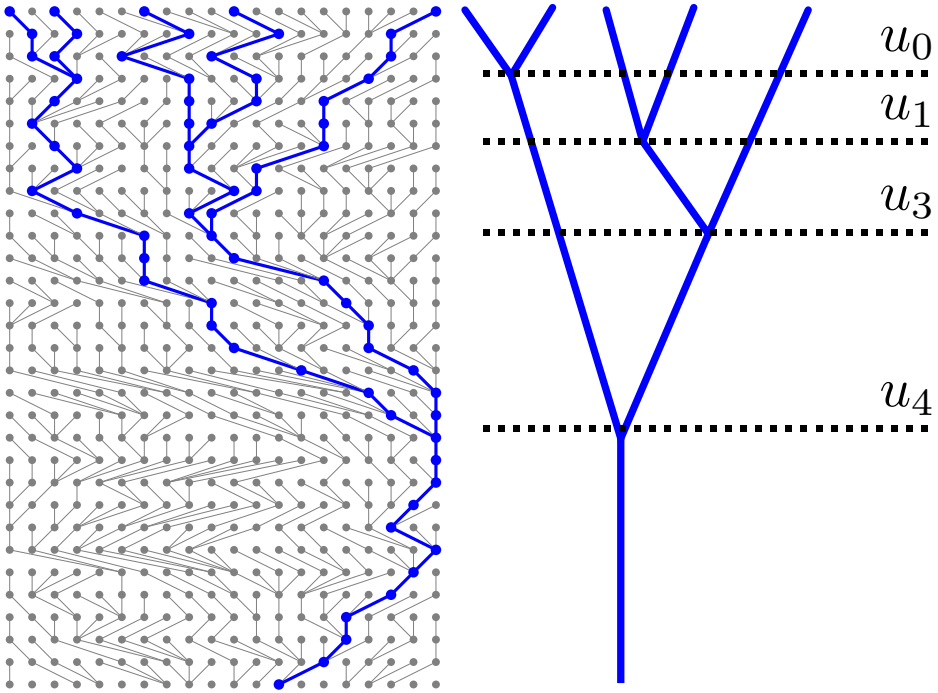
with the scaled coalescence rate

$$\lambda = \binom{k}{2}\frac{1}{2N} = \frac{k(k-1)}{2(2N)} = \frac{k(k-1)}{4N}$$

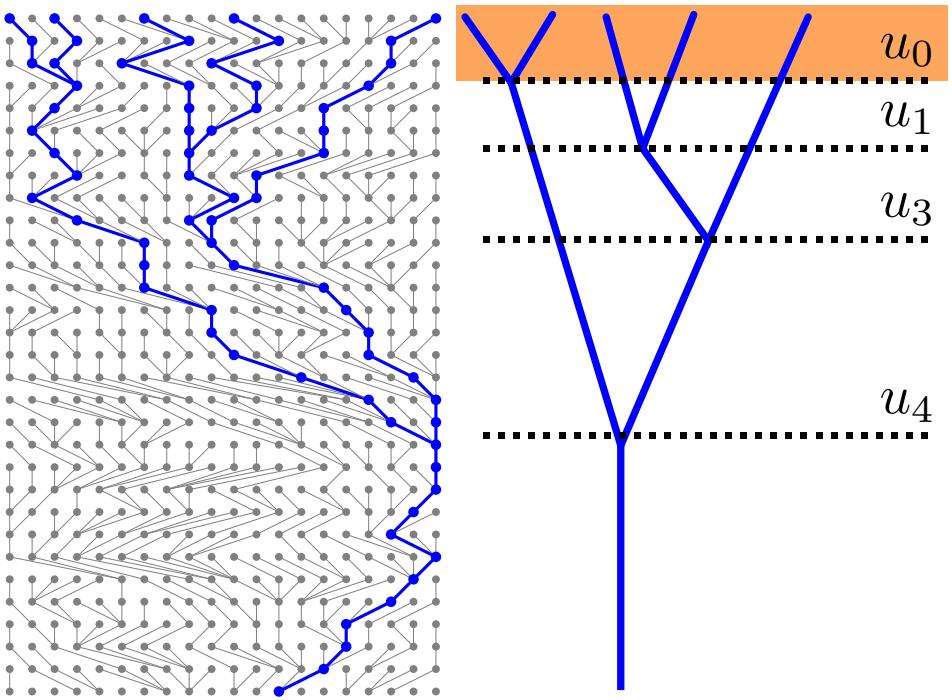
Samples larger than two



We are now able to calculate the probability of a whole relationship tree (Genealogy G). We assume that each coalescence is independent from any other:

$$\mathrm{P}(G|N)$$

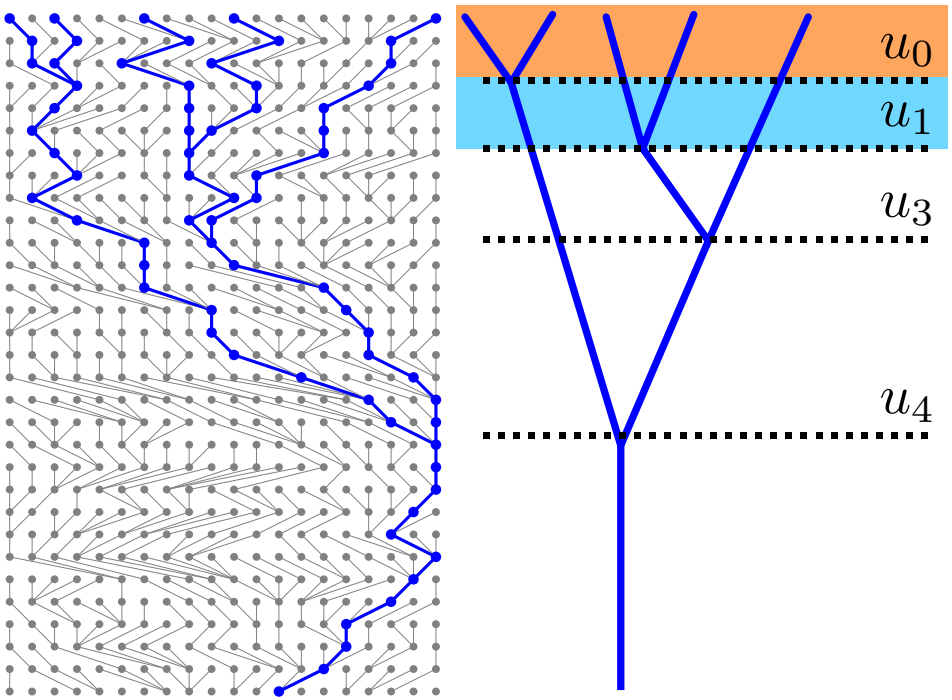
Samples larger than two



We are now able to calculate the probability of a whole relationship tree (Genealogy G). We assume that each coalescence is independent from any other:

$$\mathrm{P}(G|N) = \mathrm{P}(u_0|N, i_1, i_2) \times$$

Samples larger than two



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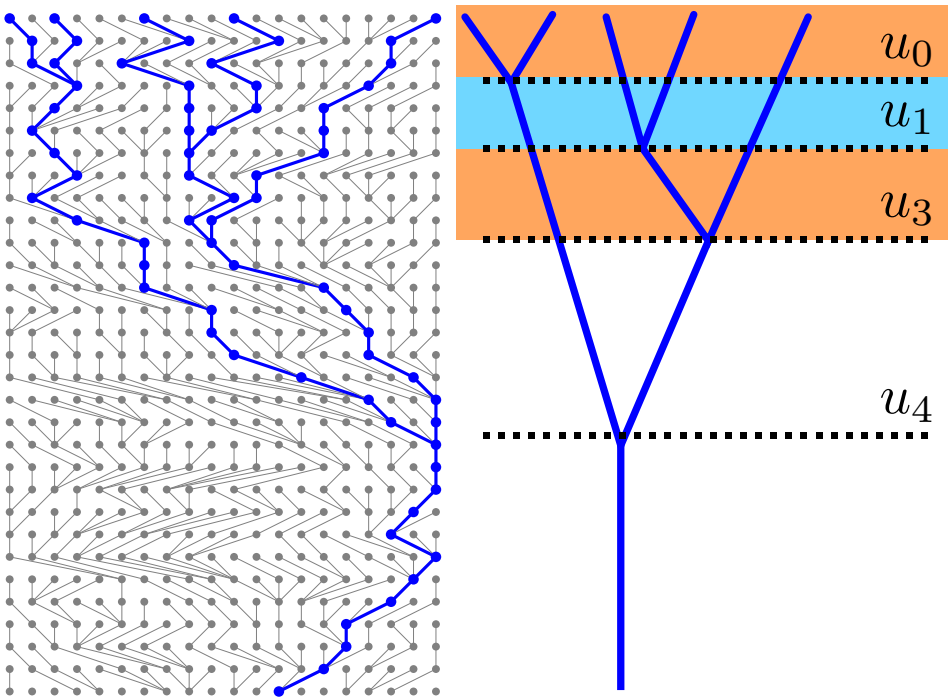
$$\mathrm{P}(G|N) = \mathrm{P}(u_0|N, i_1, i_2) \times \mathrm{P}(u_1|N, i_3, i_4)$$

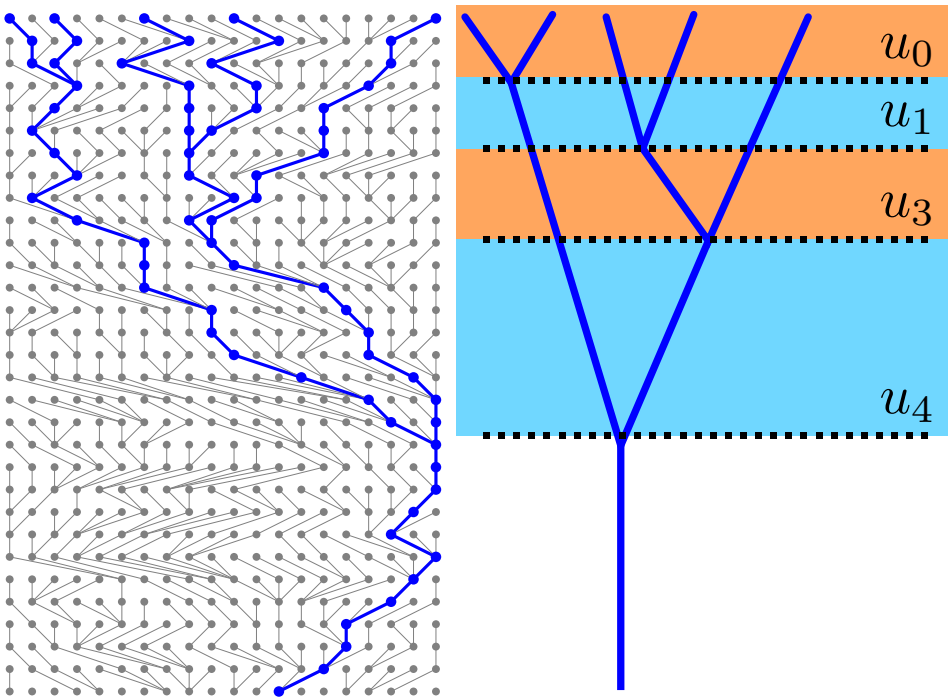
Samples larger than two



We are now able to calculate the probability of a whole relationship tree (Genealogy G). We assume that each coalescence is independent from any other:

$$\begin{aligned}\mathrm{P}(G|N) = \quad & \mathrm{P}(u_0|N, i_1, i_2) \\ & \times \mathrm{P}(u_1|N, i_3, i_4) \\ & \times \mathrm{P}(u_3|N, i_{3,4}, i_5)\end{aligned}$$

Samples larger than two



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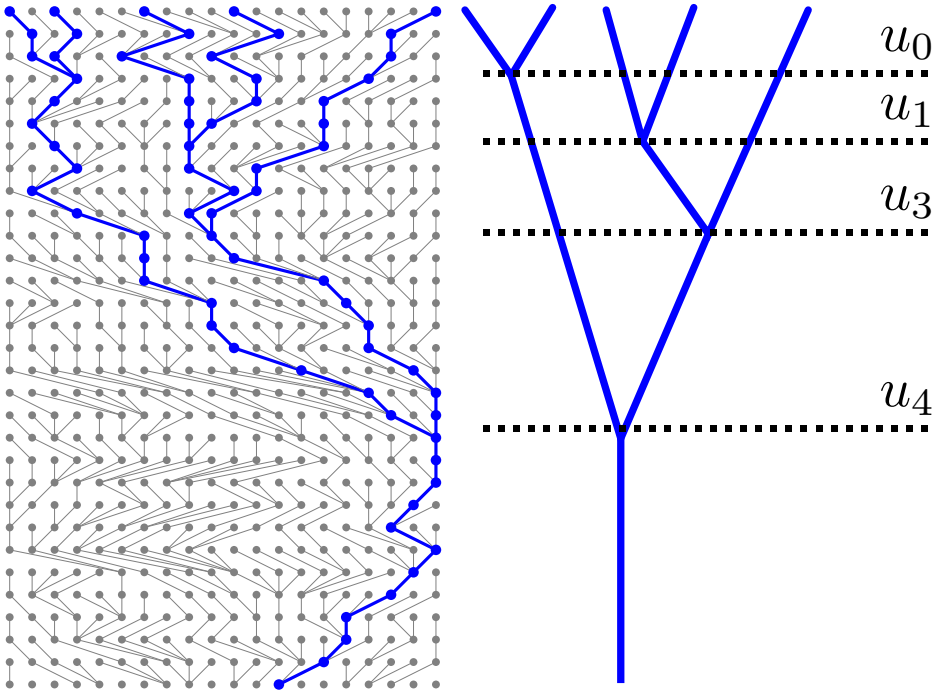
$$\begin{aligned}\mathrm{P}(G|N) = \quad & \mathrm{P}(u_0|N, i_1, i_2) \\ & \times \mathrm{P}(u_1|N, i_3, i_4) \\ & \times \mathrm{P}(u_3|N, i_{3,4}, i_5) \\ & \times \mathrm{P}(u_4|N, i_{1,2}, i_{3,4,5})\end{aligned}$$

Samples larger than two



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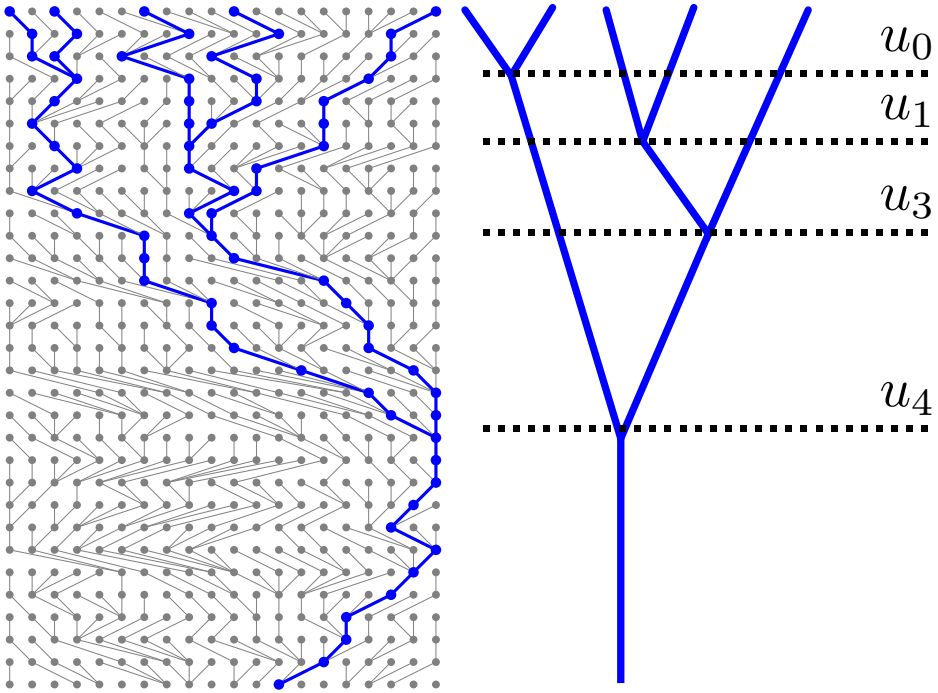
$$\mathrm{P}(G|N) = \prod_{j=0}^{T} e^{-u_j \frac{k_j(k_j-1)}{4N}} \frac{2}{4N}$$

Samples larger than two



Each interval u_j is independent of the others, the expected length of the interval is the inverse of the coalescent rate. Thus we can sum these expectations to get to expectation of the depth of the genealogy.

$$\mathbb{E}(\tau_{\text{MRCA}}) = \text{Sum of the expectation of each time interval} = \sum_{j=0}^{J} \frac{4N}{k_j(k_j-1)}$$

$$\lim_{k\to\infty} \mathbb{E}(\tau_{\text{MRCA}}) = 2N + \frac{2}{3}N + \frac{1}{3}N + \frac{1}{5}N + \frac{2}{15}N + \ldots = 4N \qquad \lim_{k\to\infty} \sigma(\tau_{\text{MRCA}}) = 4N$$

What is it good for?

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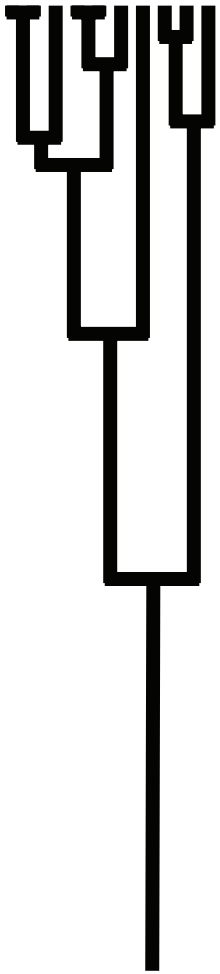
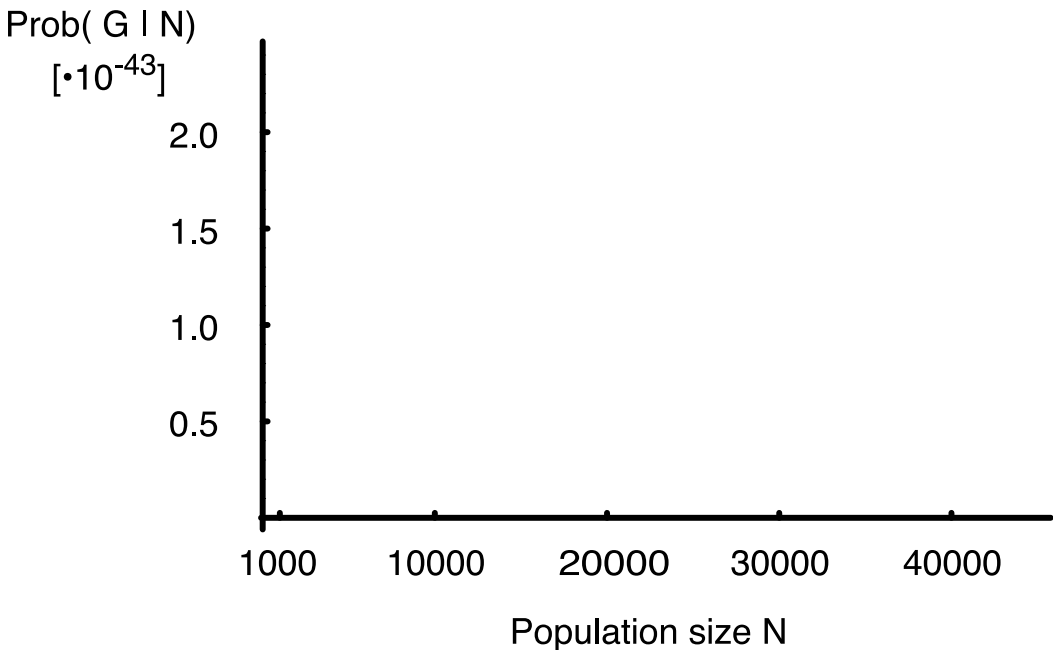


Population size estimation

using an oracle



If an oracle gives us the true relationship tree G then we can calculate the population size N.

$$\mathrm{p(G|N,n)} = \prod_{\mathrm{k=2}}^{\mathrm{n}} \exp\left(-\mathrm{u_k}\frac{\mathrm{k(k-1)}}{\mathrm{4N}}\right)\frac{2}{\mathrm{4N}}$$

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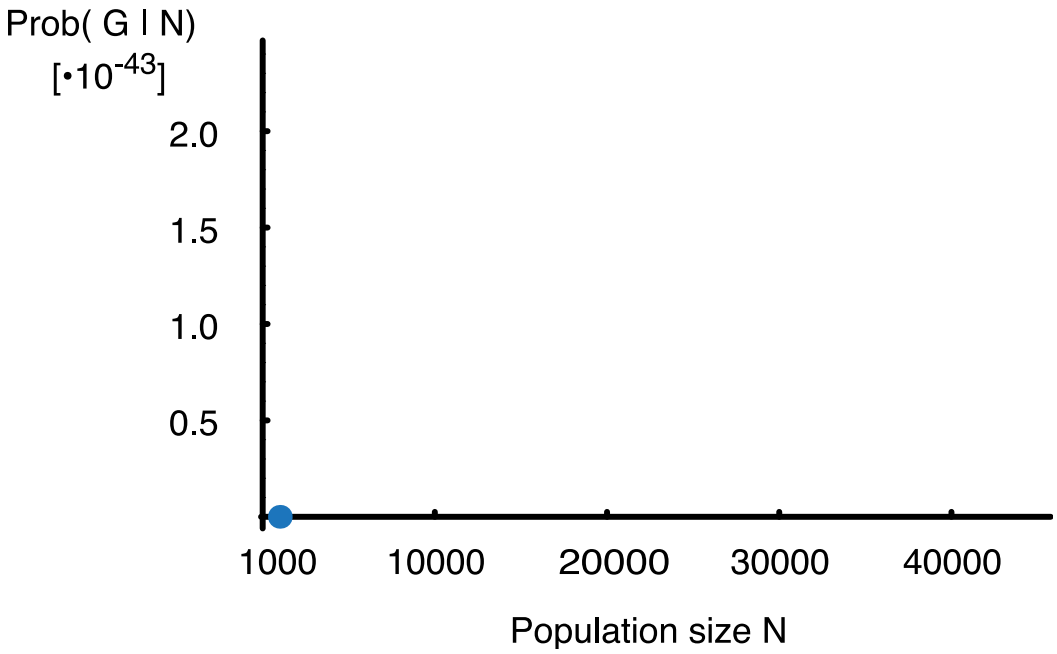

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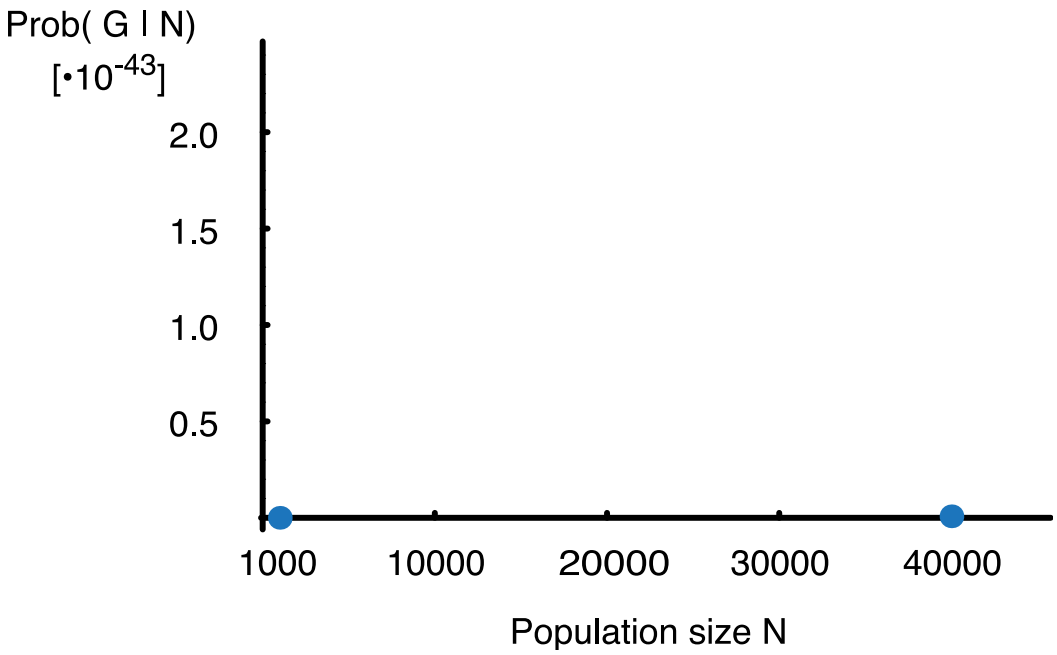

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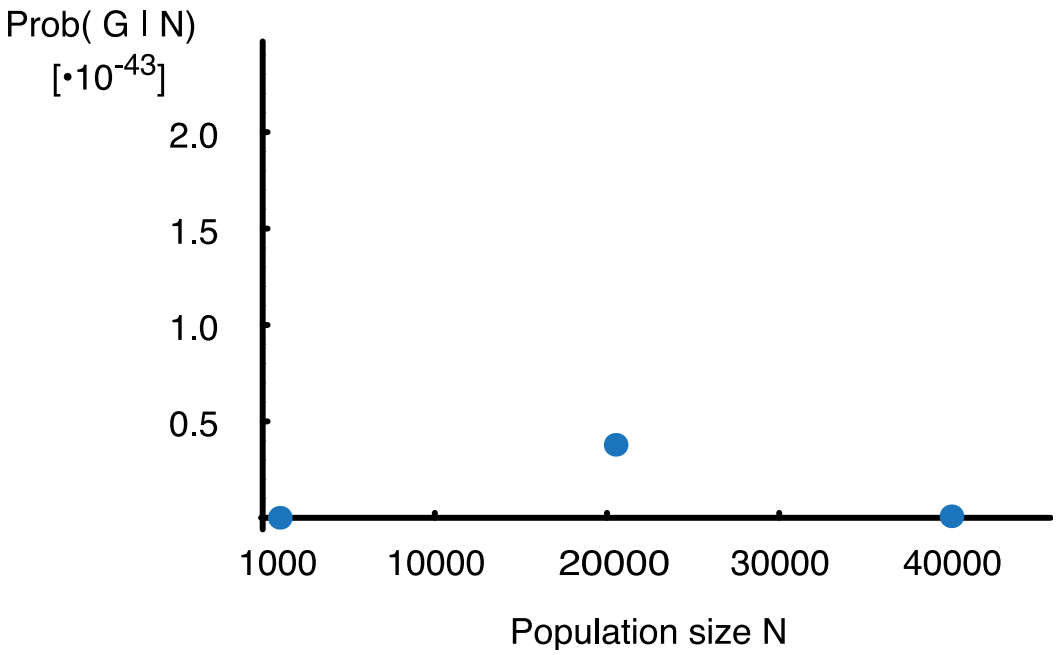

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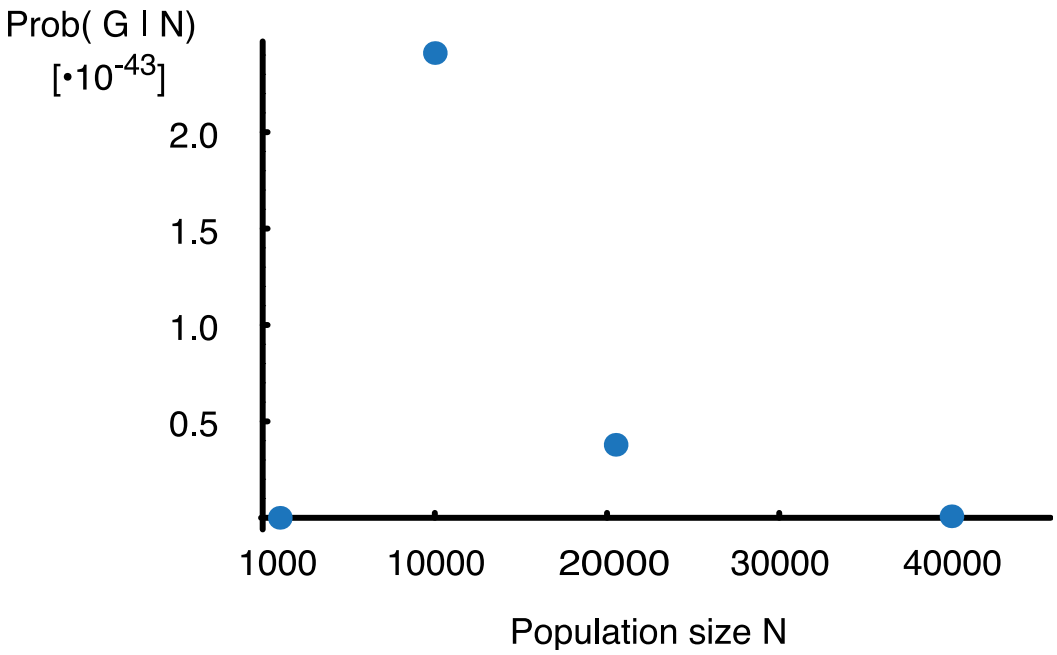

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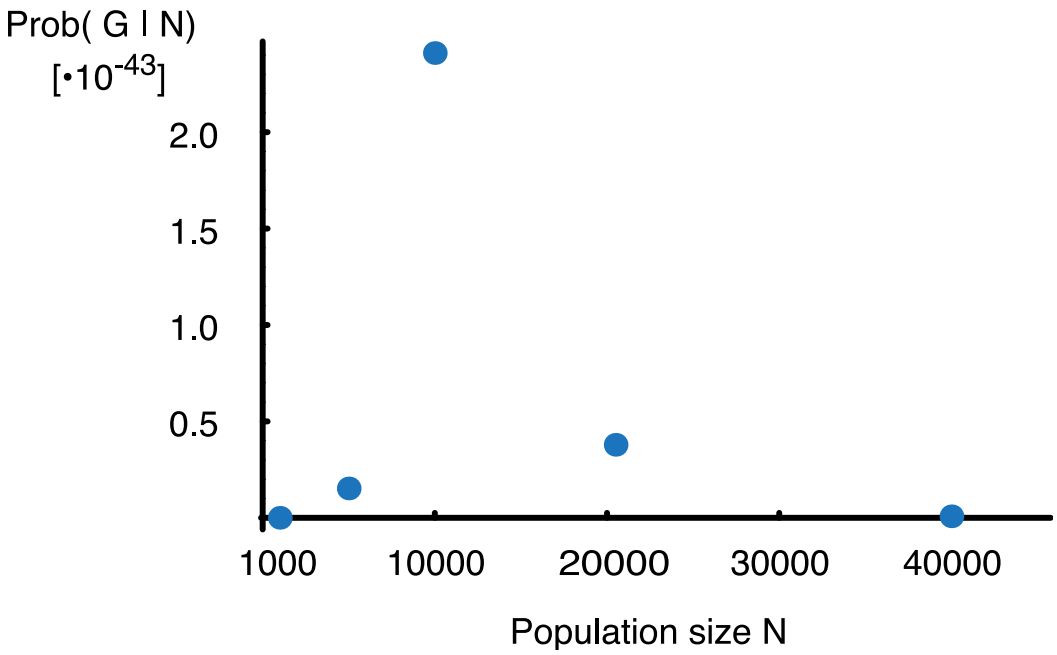

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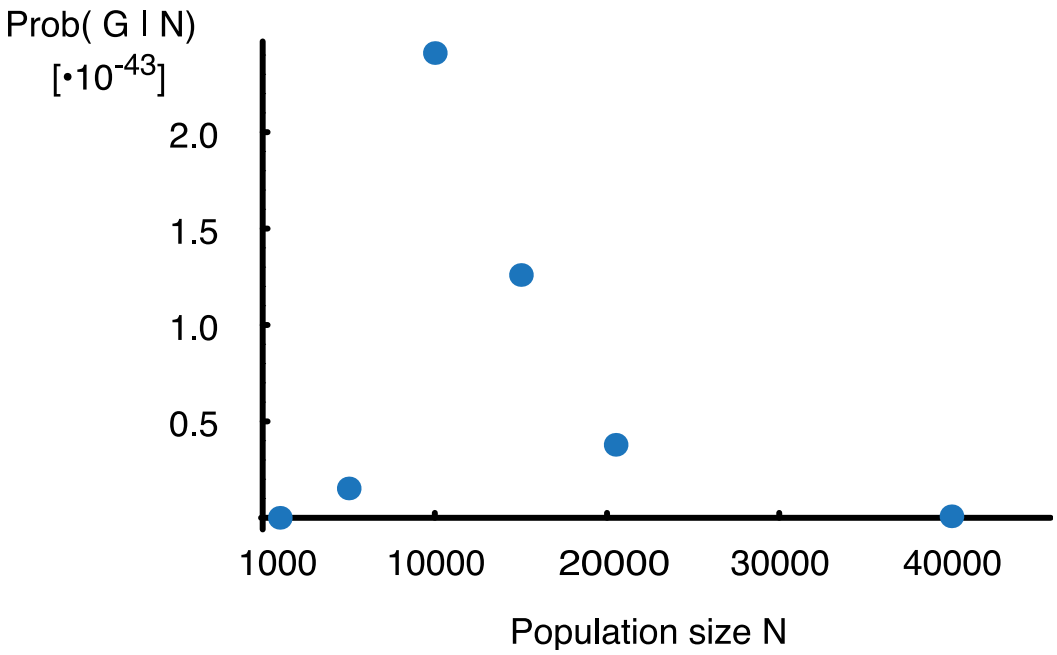

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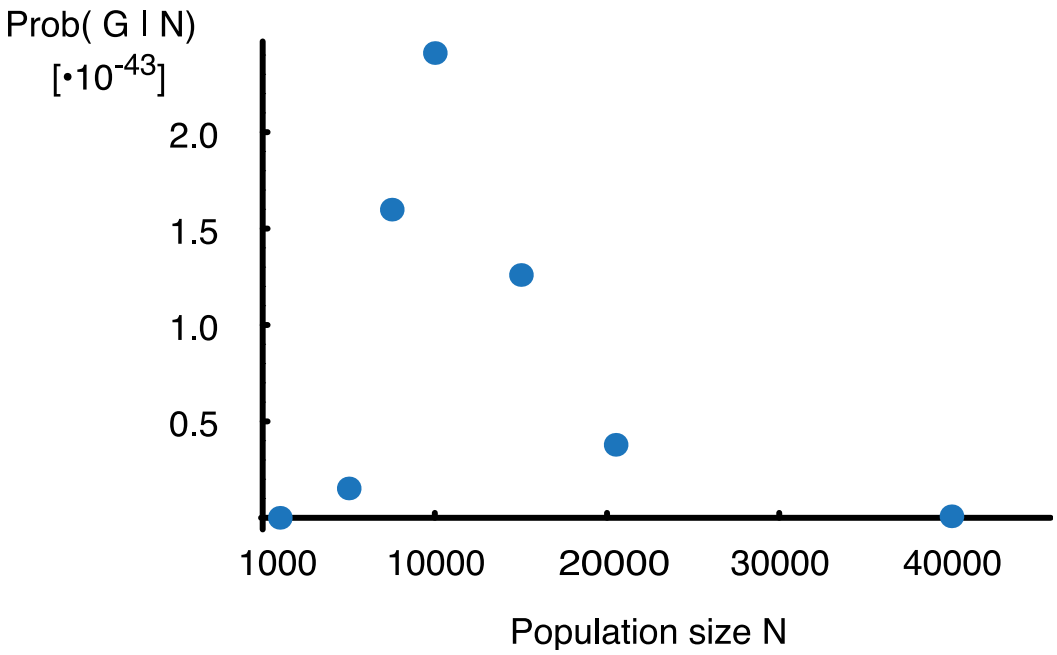

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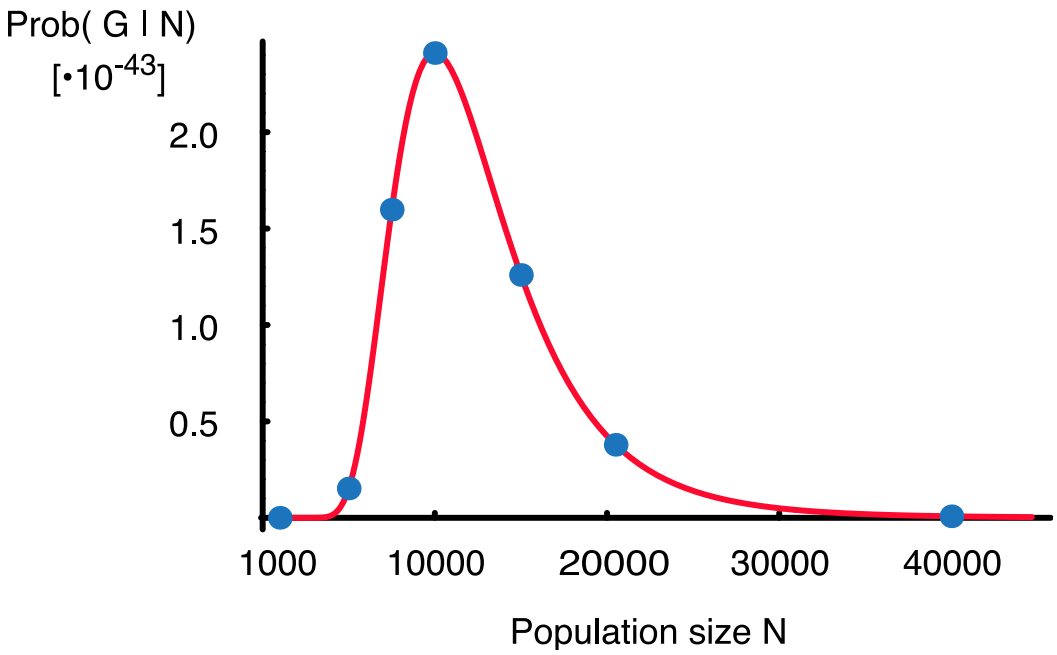

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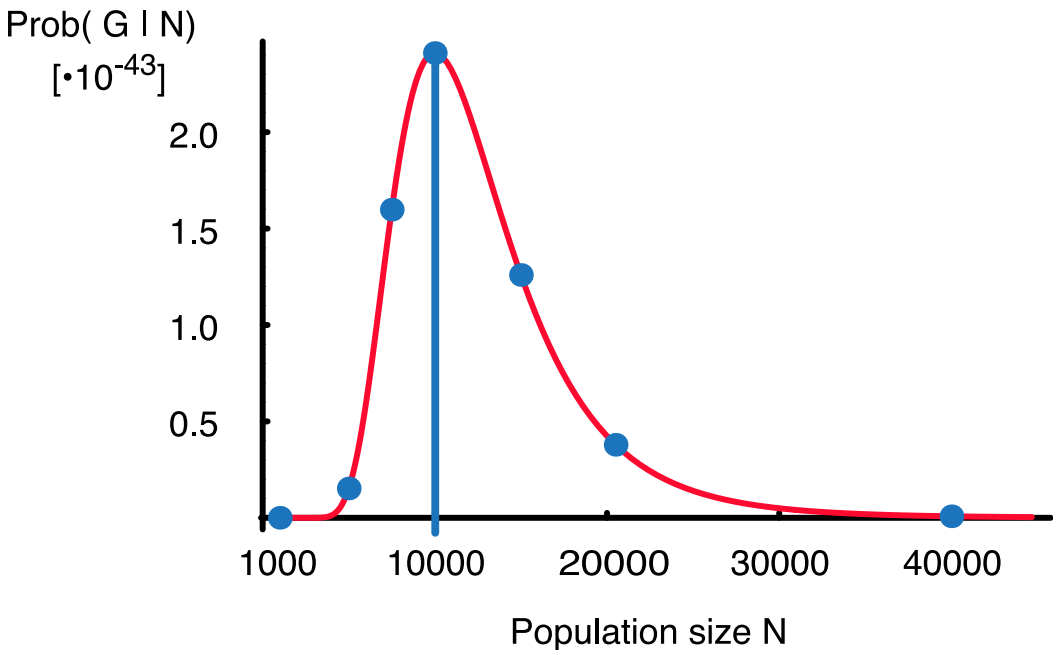

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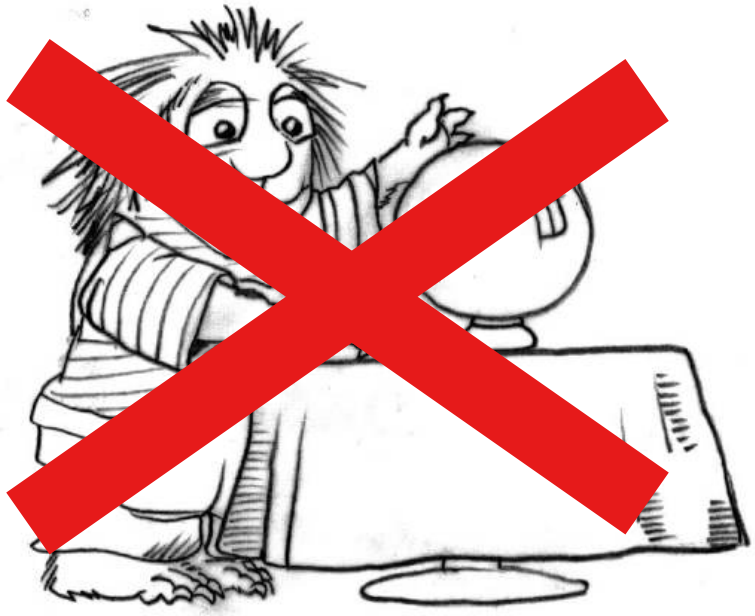
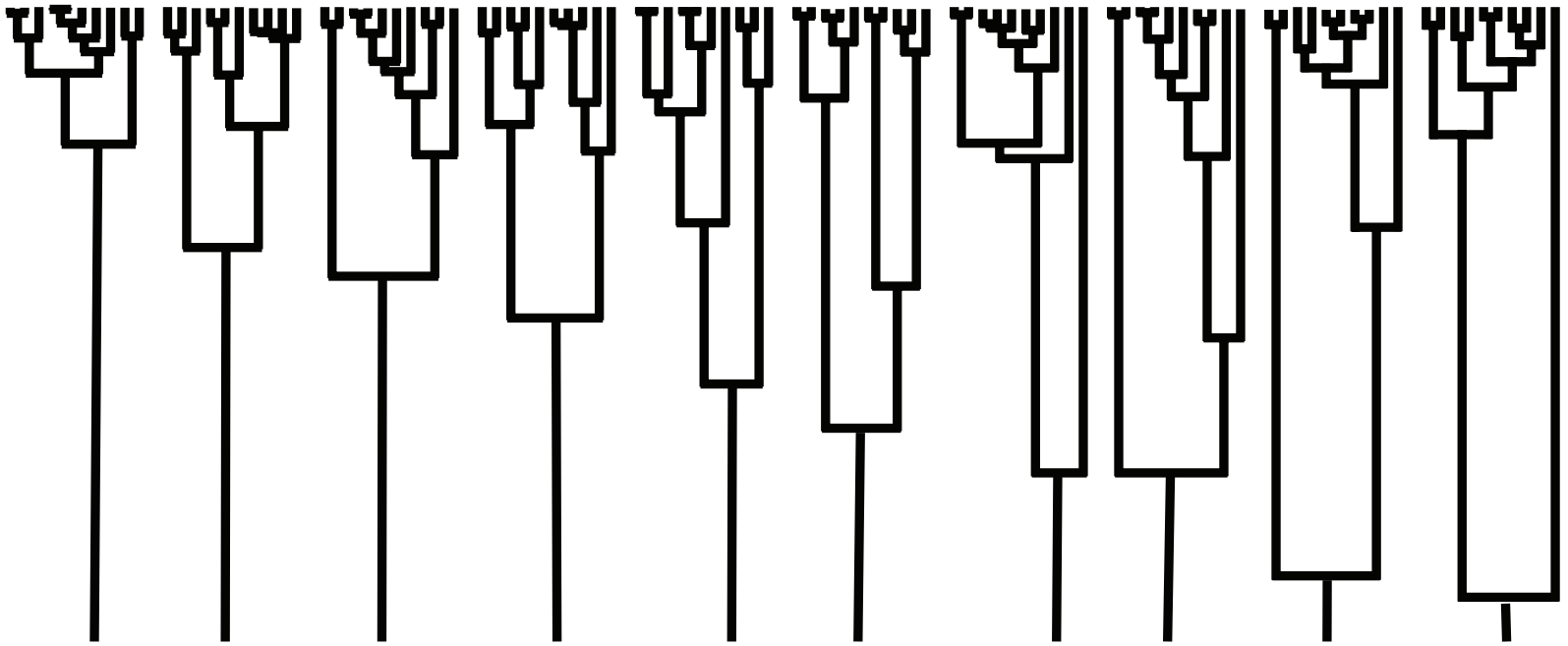
Population size estimation

There are at least two problems with the oracle-approach:

- There is no oracle to gives us clear information!
- We do not record genealogies, our data are sequences, microsatellite loci!
- What about the variability of the coalescence process?

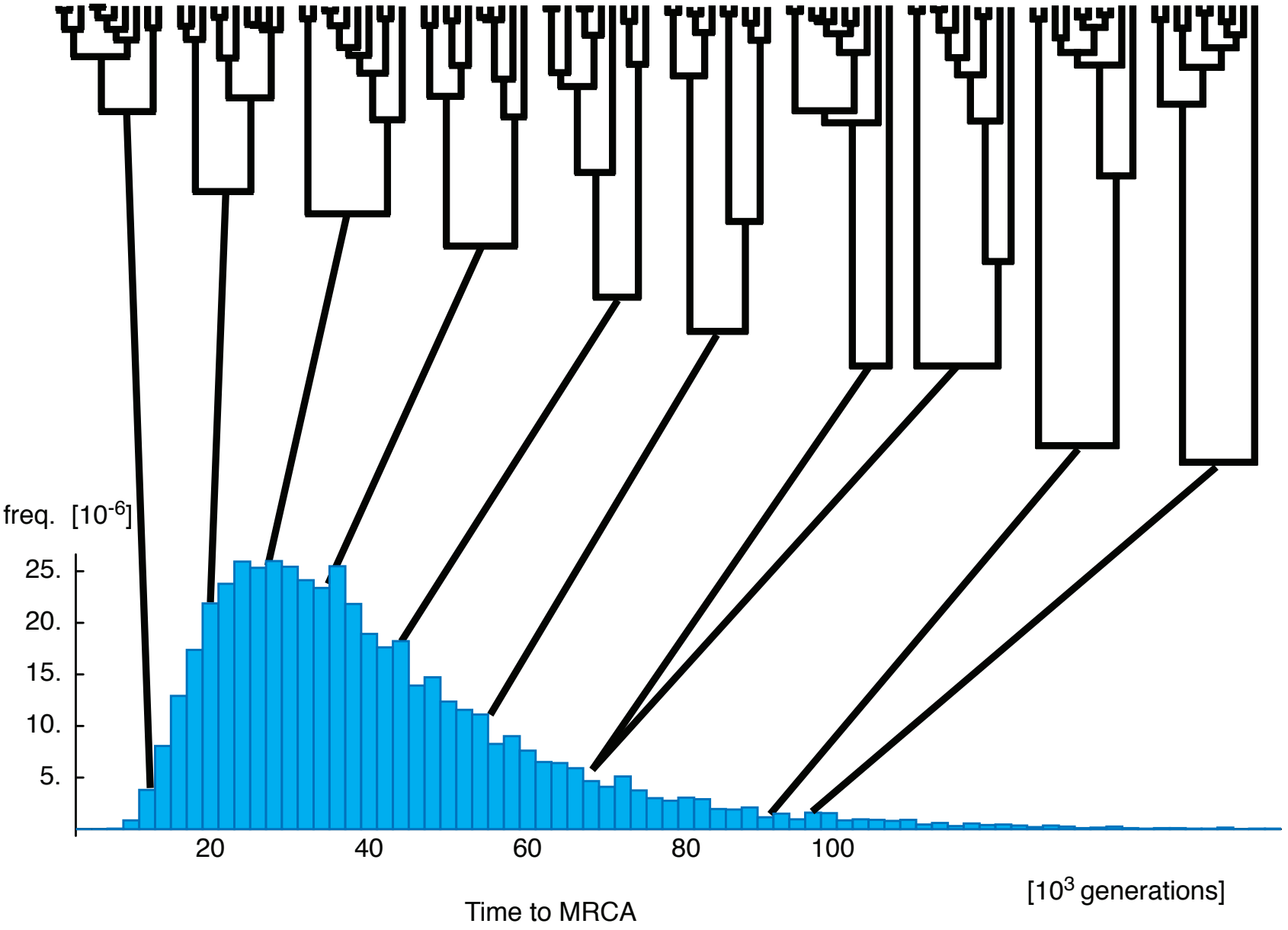


Variability of the coalescent process



All genealogies were simulated with the same population size $N_e = 10,000$

Variability of the coalescent process



MRCA = most recent common ancestor (last node in the genealogy)

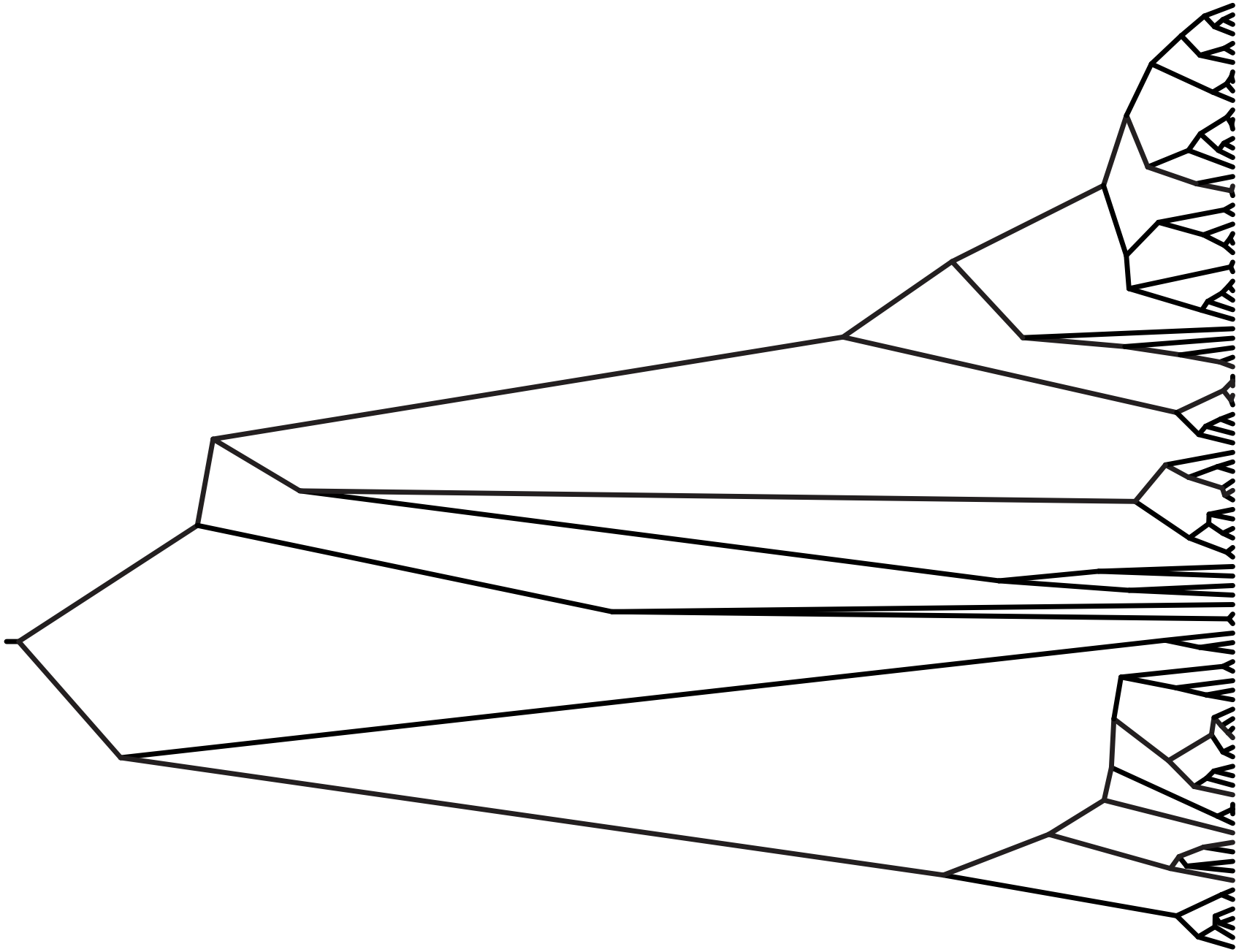
Kingman's n-coalescent is an approximation

- All individuals have the same fitness (no selection).
- All individuals have the same chance to be in the sample (random sampling).
- The coalescent allows only merging two lineages per generation. This restricts us to to have a much smaller sample size than the population size.

$$n << N$$

- Yun-Xin Fu (2005) described the exact coalescent for the Wright-Fisher model and derived a maximal sample size $n < \sqrt{4N}$ for a diploid population. Although this may look like a severe restriction for the use of the coalescence in small populations, it turned out that the coalescence is rather robust and that even sample sizes close to the effective population size are not biasing immensely.

Kingman's n-coalescent is an approximation



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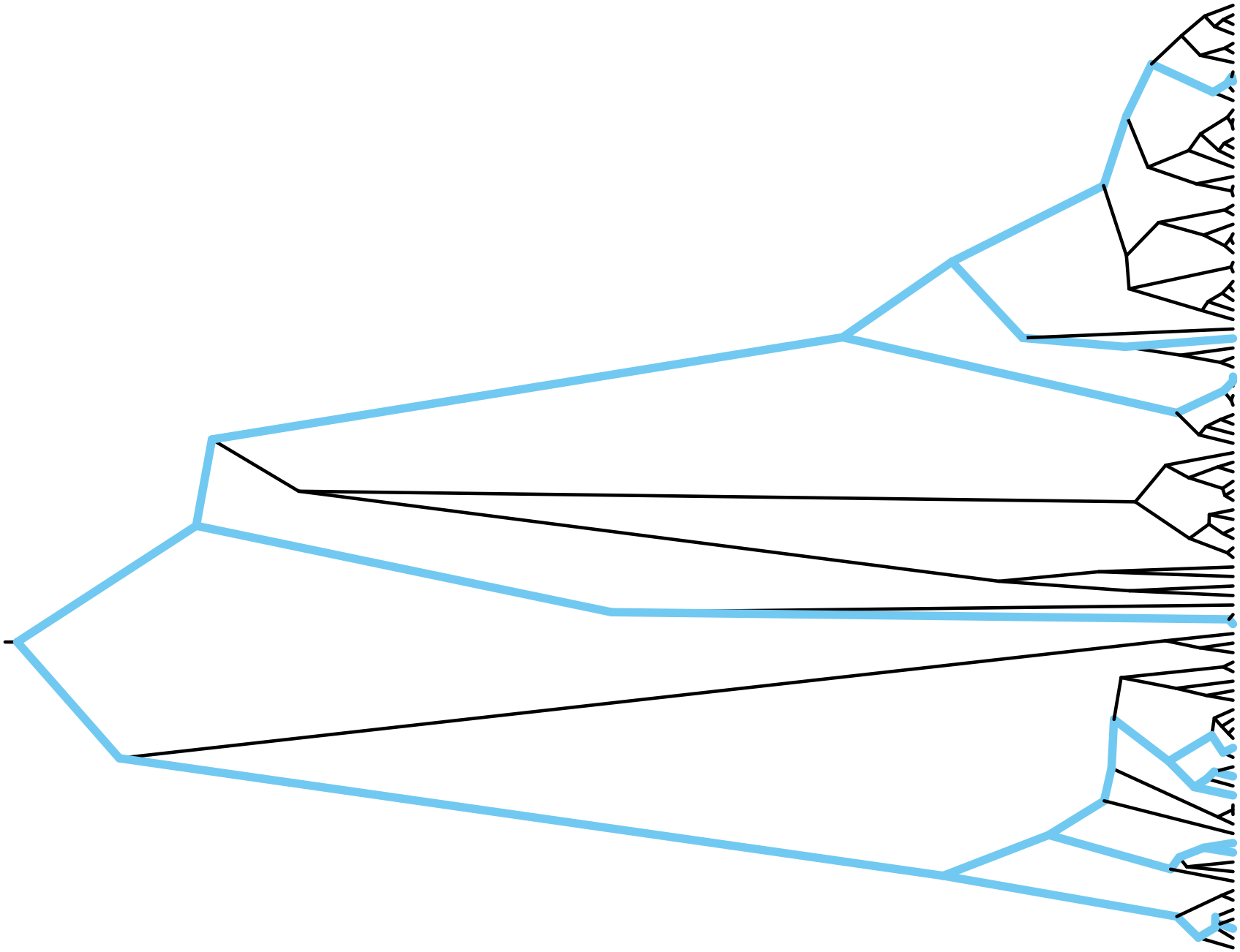

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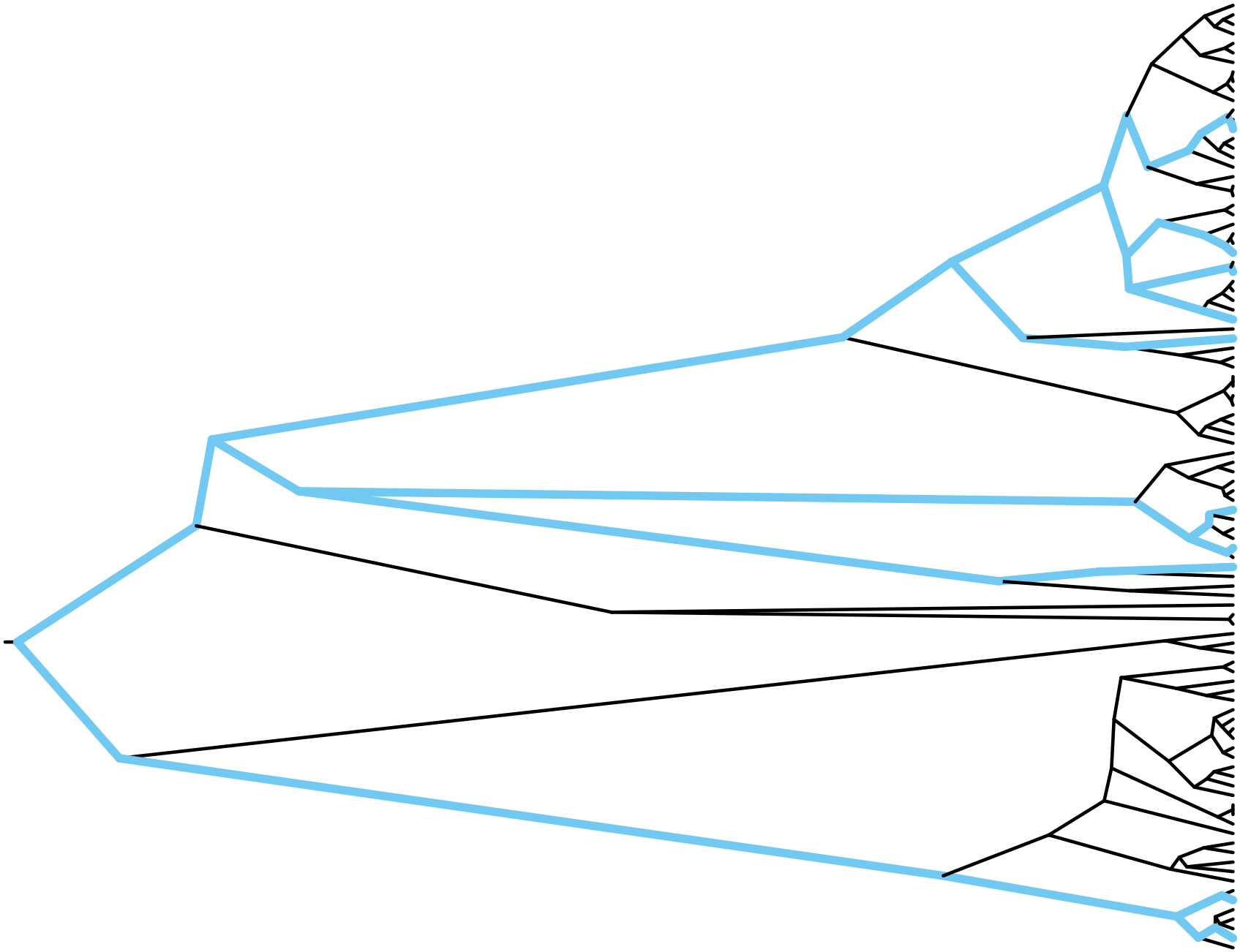

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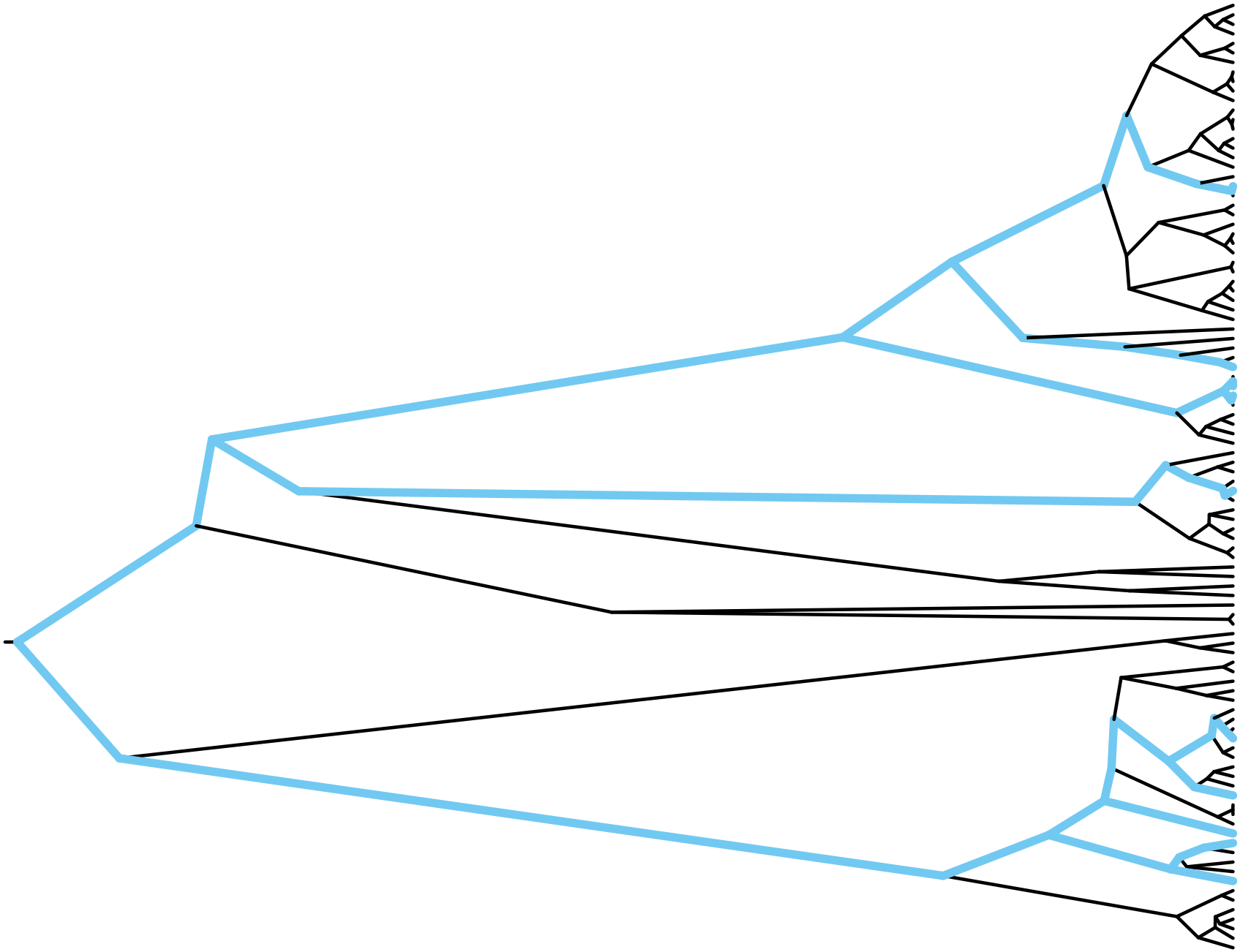
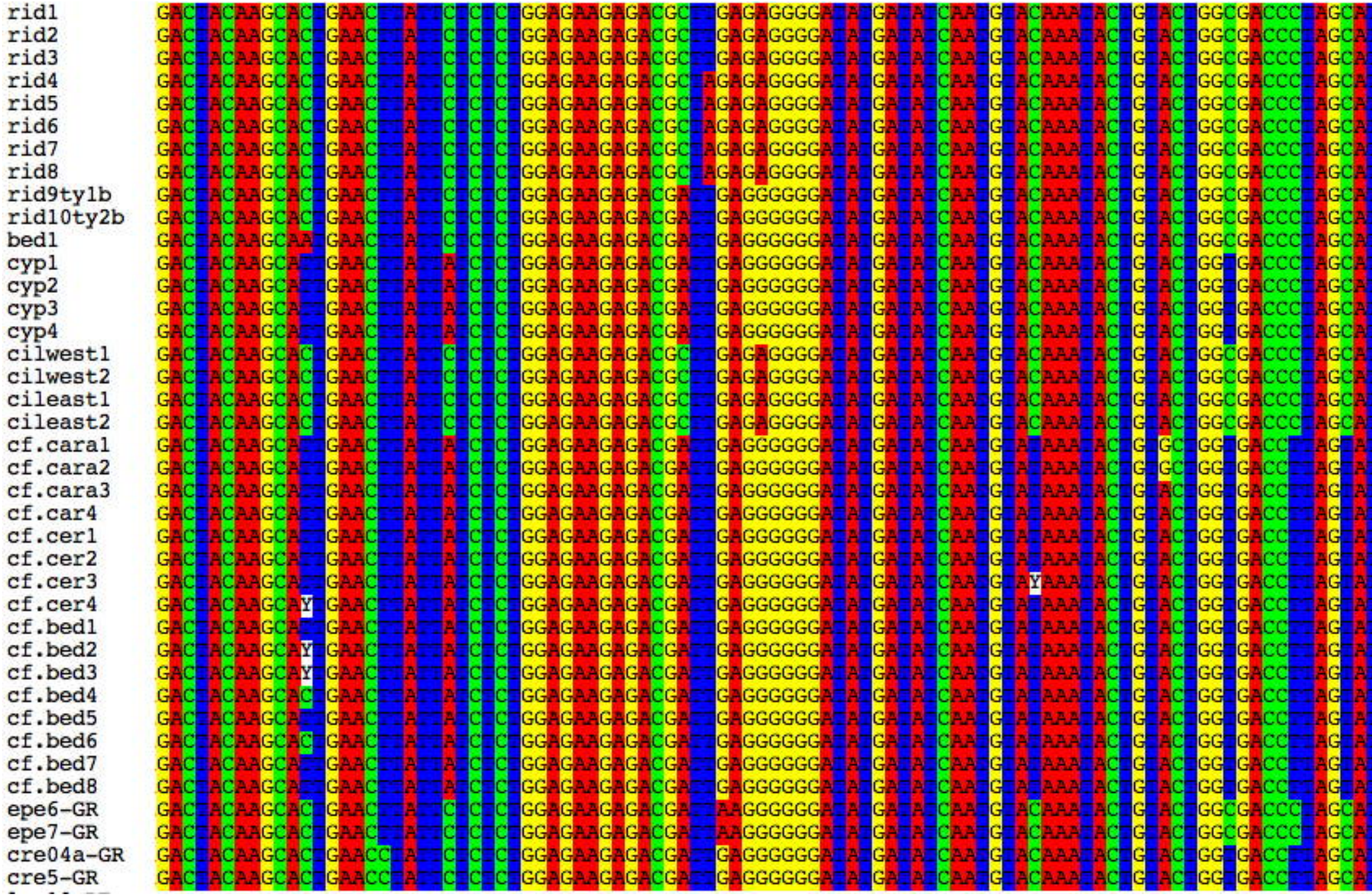


Observations

- Large samples coalesce on average in $4N$ generations.
- The time to the most recent common ancestor (TMRCA) has a large variance
- Even a sample with few individuals can most often recover the same TMRCA as a large sample.
- The sample size should be much smaller than the population size, although severe problems appear only with sample sizes of the same magnitude as the population size, or with non-random samples because Kingman's coalescence process assumes that maximally two sample lineages coalesce in any generation.
- With a known genealogy we can estimate the population size. Unfortunately, the true genealogy of a sample is rarely known.

Genealogy and data

our data looks like this:



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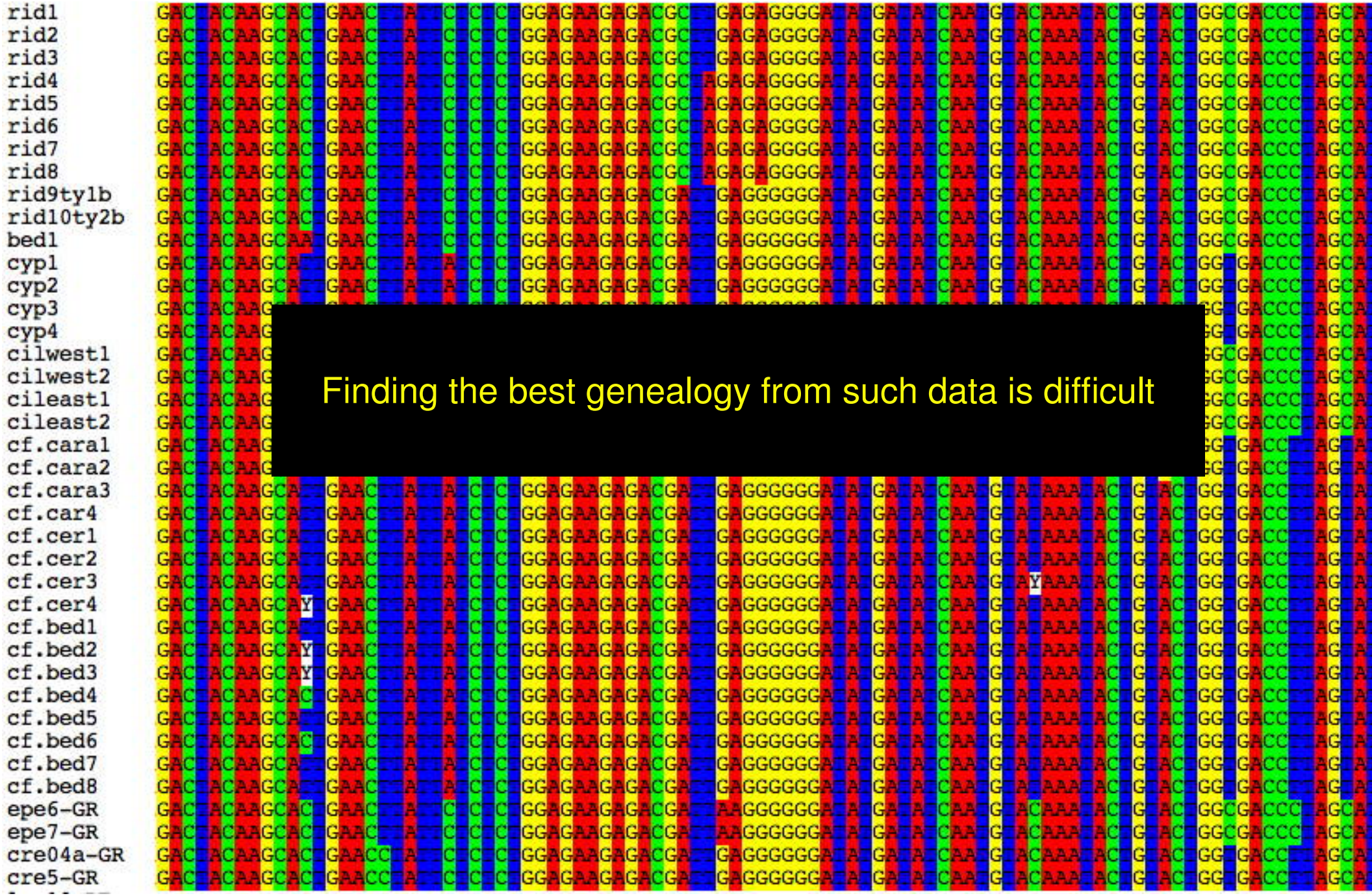

Genetic data and the coalescent

- Finite populations loose alleles due to genetic drift
- Mutation introduces new alleles into a population at rate μ
- With $2N$ chromosomes we can expect to see every generation $2N\mu$ new mutations. The population size N is positively correlated with the mutation rate μ.
- With genetic data sampled from several individuals we can use the mutational variability to estimate the population size.

Population size

The observed genetic variability

$$\mathcal{S} = f(N, \mu, n).$$

Different N and appropriate μ can give the same number of mutations. For example, for 100 loci sampled from 20 individuals with 1000bp each, we get :

N	μ	$4N\mu$	$\hat{S}$	σ_S^2
1250	10^{-5}	0.05	153.95	16.25
12500	10^{-6}	0.05	152.89	16.05

Using genetic variability alone therefore does not allow to disentangle N and μ.

With multiple dated samples and known generation time we can estimate N and μ independently.

Mutation-scaled population size

By convention we express most results as the compound $N\mu$ and an inheritance scalar x, for simplicity we call this the mutation-scaled population size

$$\Theta = xN\mu,$$

where μ is the mutation rate per generation and per site. With a mutation rate per locus we use θ.

- for diploids: $\Theta = 4N\mu$.
- for haploids: $\Theta = 2N\mu$.
- For mtDNA in diploids with strictly maternal inheritance this leads to $\Theta = 2N_f\mu$, and if the sex ratio is $1:1$ then $\Theta = N\mu$

Most real populations do not behave exactly like Wright-Fisher populations, therefore we subscript N and call it the effective population size N_e, and consider Θ the mutation-scaled EFFECTIVE population size.

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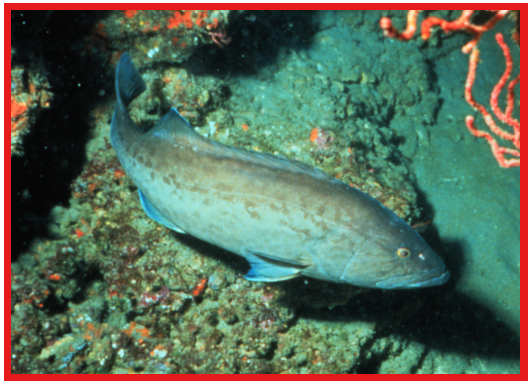

Gag Grouper starts out as a female and later in live becomes male.

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Historical humpback whale population size

Humpback whales in the North Atlantic: Census population size around 12,000.

Historical humpback whale population size

using the data by Joe Roman and Stephen R. Palumbi (Science 2003 301: 508-510)

$\Theta = 2N_{♀}\mu$	0.01529	Population size of the North Atlantic population, estimated using migrate
$N_{♀} = \frac{\Theta}{2\mu}$	31,854	with $\mu = 2.0 \times 10^{-8} \mathrm{bp}^{-1} \mathrm{year}^{-1}$ and a generation time of 12 years
$N_e = N_{♀} + N_{♂}$	63,708	Sex ratio is 1:1
$N_B = 2N_e$	127,417	ratio N_B/N_e assumed, using other data
$N_T = N_B \frac{N_{\text{juveniles}} + N_{\text{adults}}}{N_{\text{adults}}}$	203,867	from catch and survey data (used a ratio of 1.6)

More modern estimates for mtDNA: $150,000$ [improved estimation of mutation rate]; for nucDNA: $112,000 (45,000 - 235,000)$ [Conservation Genetics (2013) 14:103114]

Using the infinite sites model we use the number of variable sites S per locus to calculate the mutation-scaled population size:

$$\theta_W = \frac{S}{\sum_{k=1}^{n-1} \frac{1}{k}}$$

from a sample of n individuals. For a single population the Watterson's estimator works marvelously well, but it is vulnerable to population structure.

Watterson's θ_W uses a mutation rate per locus! To compare with other work use mutation rate per site.

[in the second part of the coalescence talk I will discuss estimates based on site frequency spectra]

Construction of a versatile estimator

For Bayesian inference we want to calculate the probability of the model parameters given the data $\mathrm{p}(\text{model}|\mathrm{D})$.

Coalescent — to describe the population genetic processes.

Mutation model — to describe the change of genetic material over time.

Wikimedia: Neon sign at Autonomy in Cambridge UK

Construction of a versatile estimator

We calculate the Posterior distribution $\mathrm{p}(\Theta|\mathrm{D})$ using Bayes' rule

$$\mathrm{p}(\Theta|\mathrm{D}) = \frac{\mathrm{p}(\Theta)\mathrm{p}(\mathrm{D}|\Theta)}{\mathrm{p}(\mathrm{D})}$$

where $\mathrm{p}(\mathrm{D}|\Theta)$ is the likelihood of the parameters.

Wikimedia: Neon sign at Autonomy in Cambridge UK

(almost) Felsenstein equation

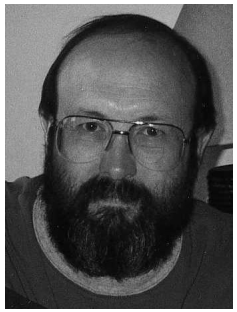
aka Likelihood calculation

$$p(D|\mathbf{\Theta}, G) = \mathrm{p}(\mathrm{G}|\mathbf{\Theta})\mathrm{p}(\mathrm{D}|\mathrm{G})$$

$p(G|\mathbf{\Theta})$

The probability density of a genealogy given parameters.

$\mathrm{p}(\mathrm{D}|\mathrm{G})$



The probability density of the data for a given genealogy. Phylogeneticists know this as the tree-likelihood.

Felsenstein equation

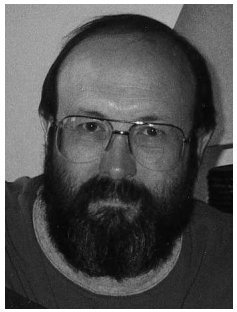
aka Likelihood calculation

$$p(D|\mathbf{\Theta}) = \int_G \mathrm{p(G|\mathbf{\Theta})p(D|G)dG}$$

$p(G|\mathbf{\Theta})$

The probability density of a genealogy given parameters.

$\mathrm{p(D|G)}$



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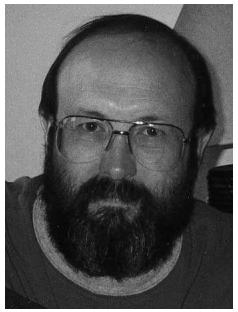

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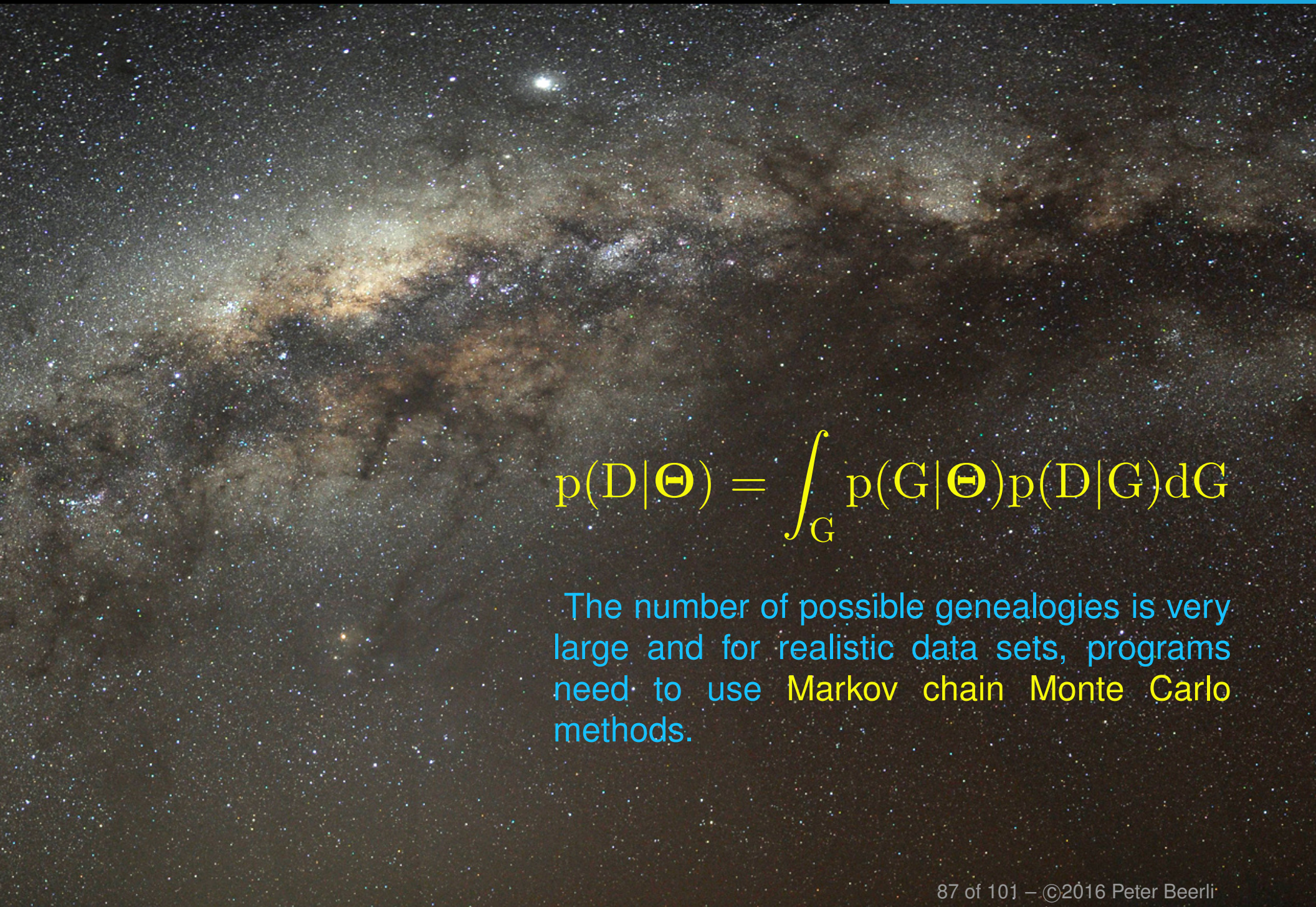
Problem with integration formula

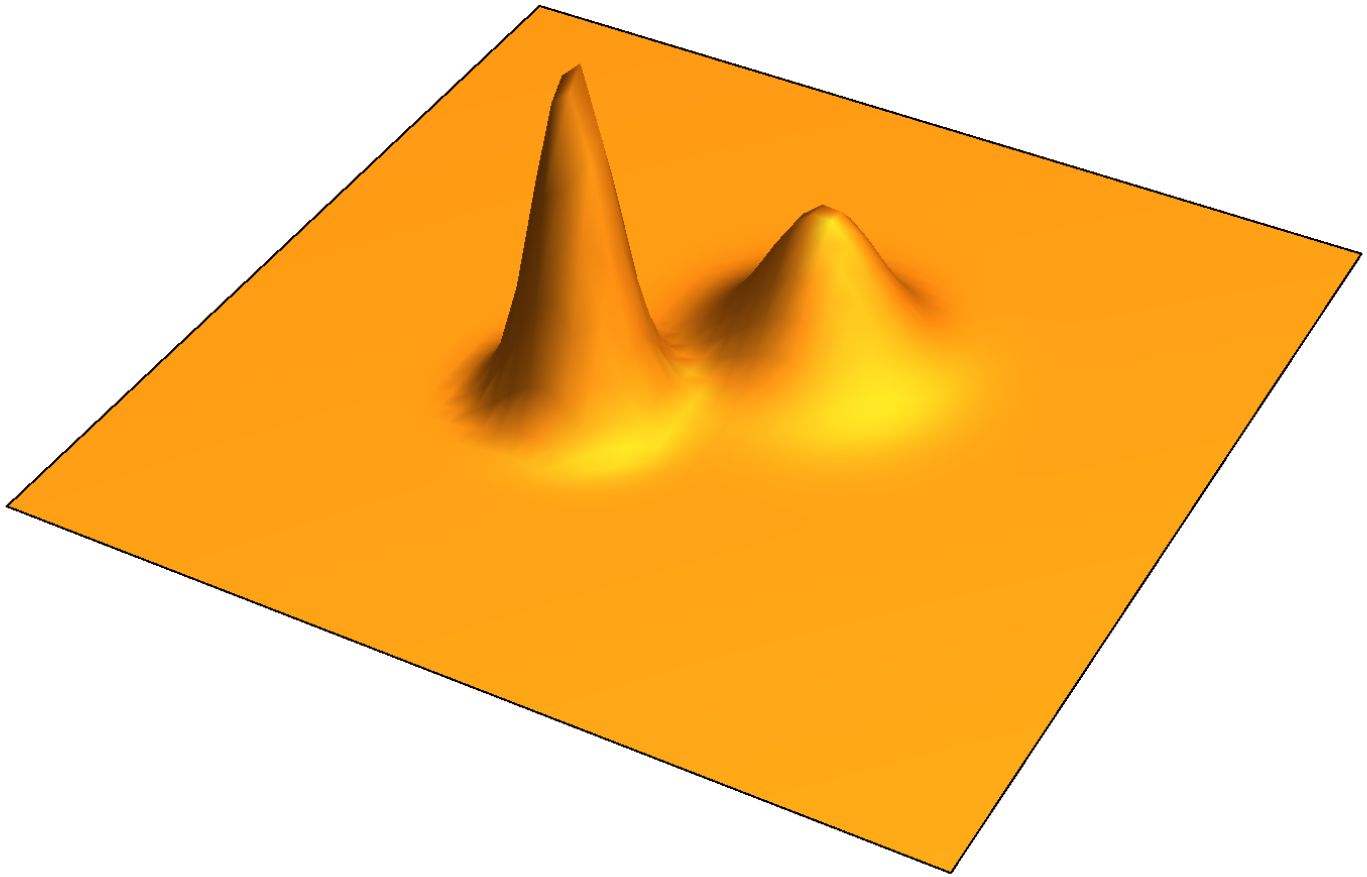


$$p(D|\Theta) = \int_G p(G|\Theta)p(D|G)dG$$

The number of possible genealogies is very large and for realistic data sets, programs need to use Markov chain Monte Carlo methods.

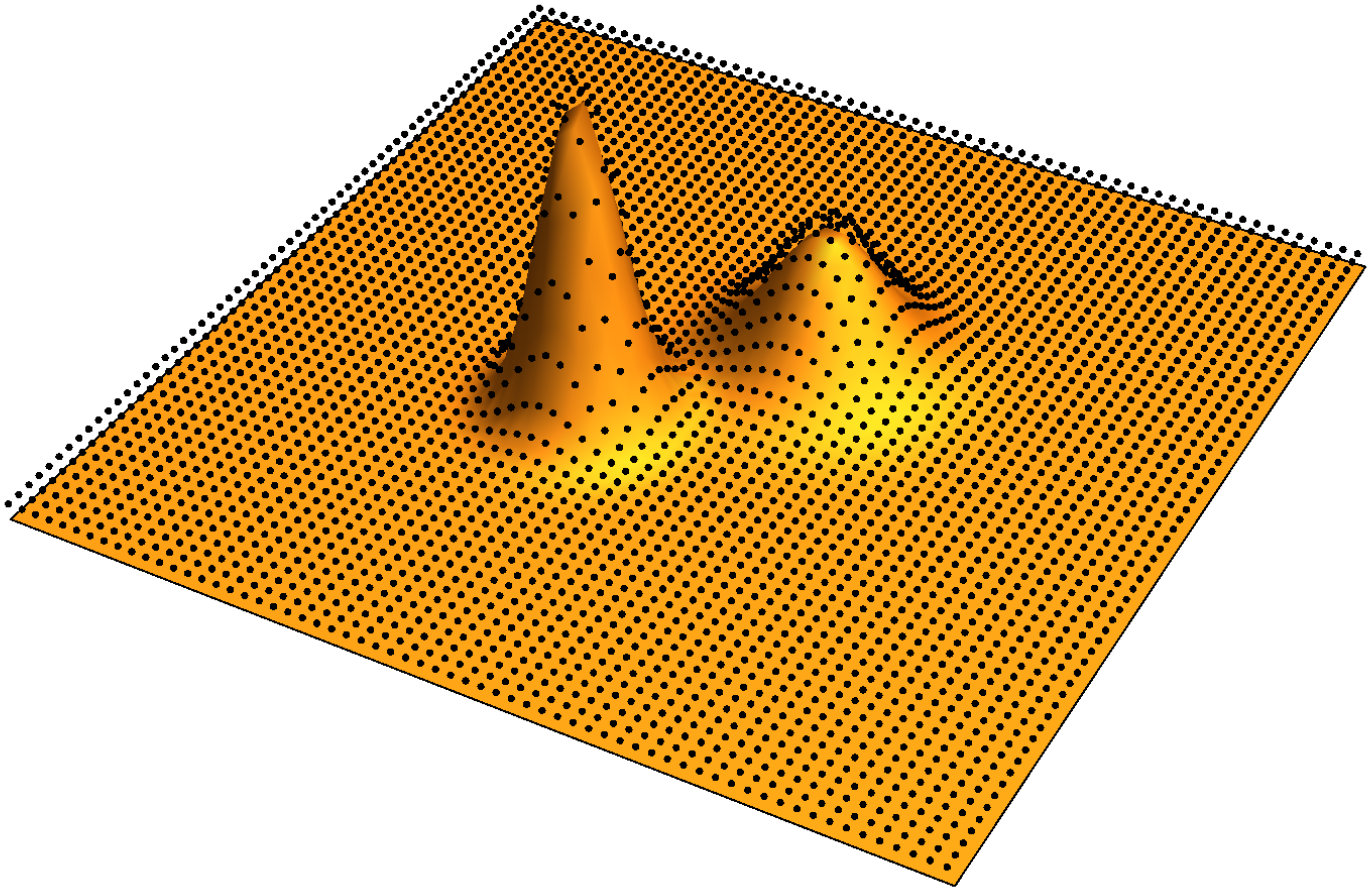
Naive integration approach

Plane of all trees

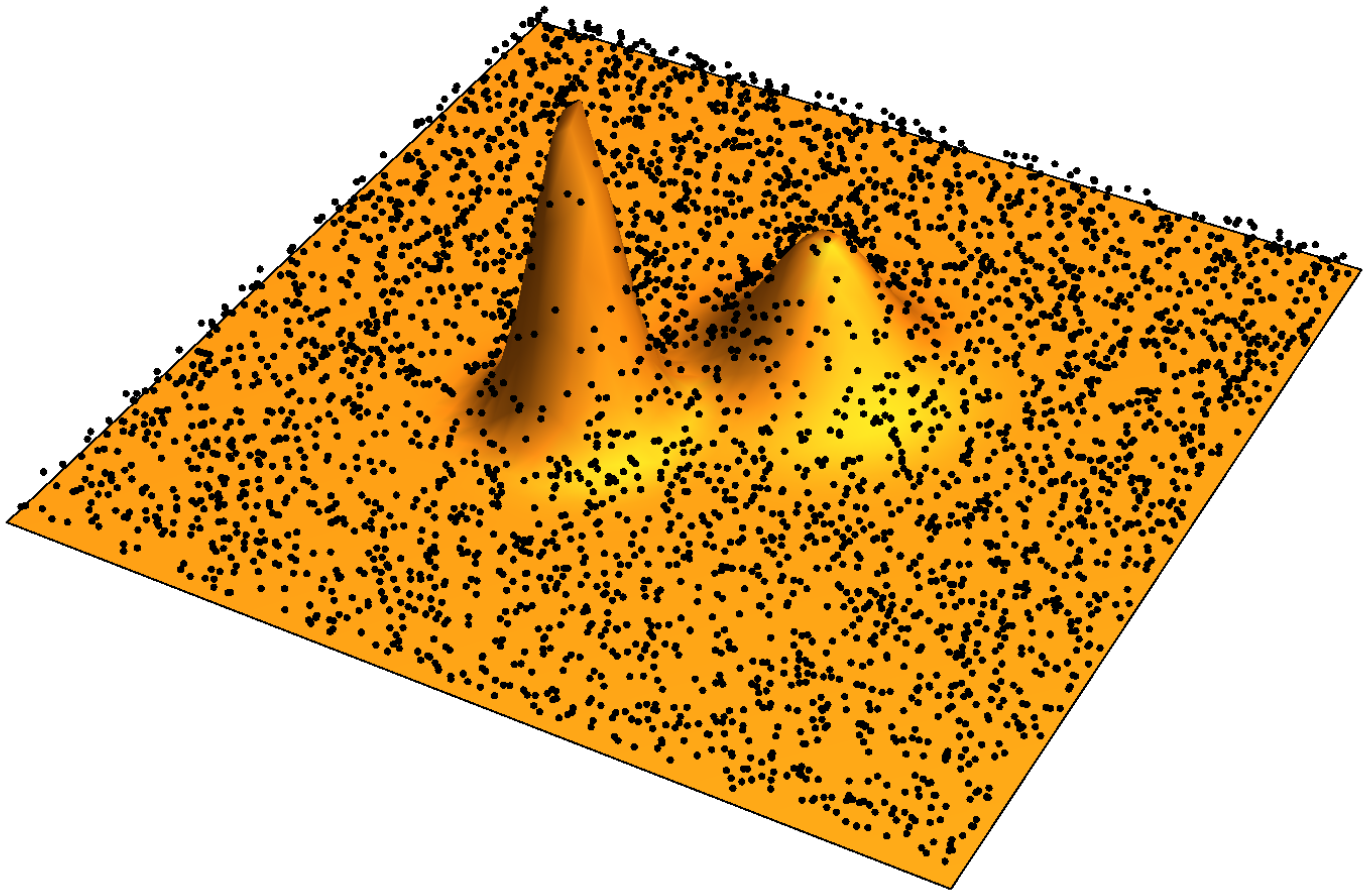


Naive integration approach

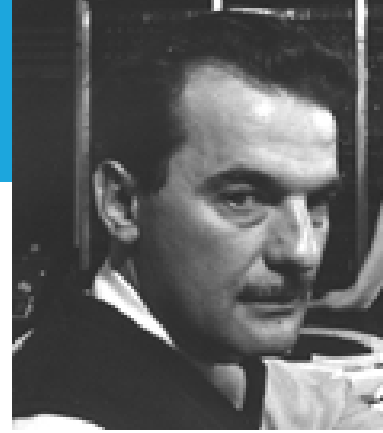
Riemann's sum

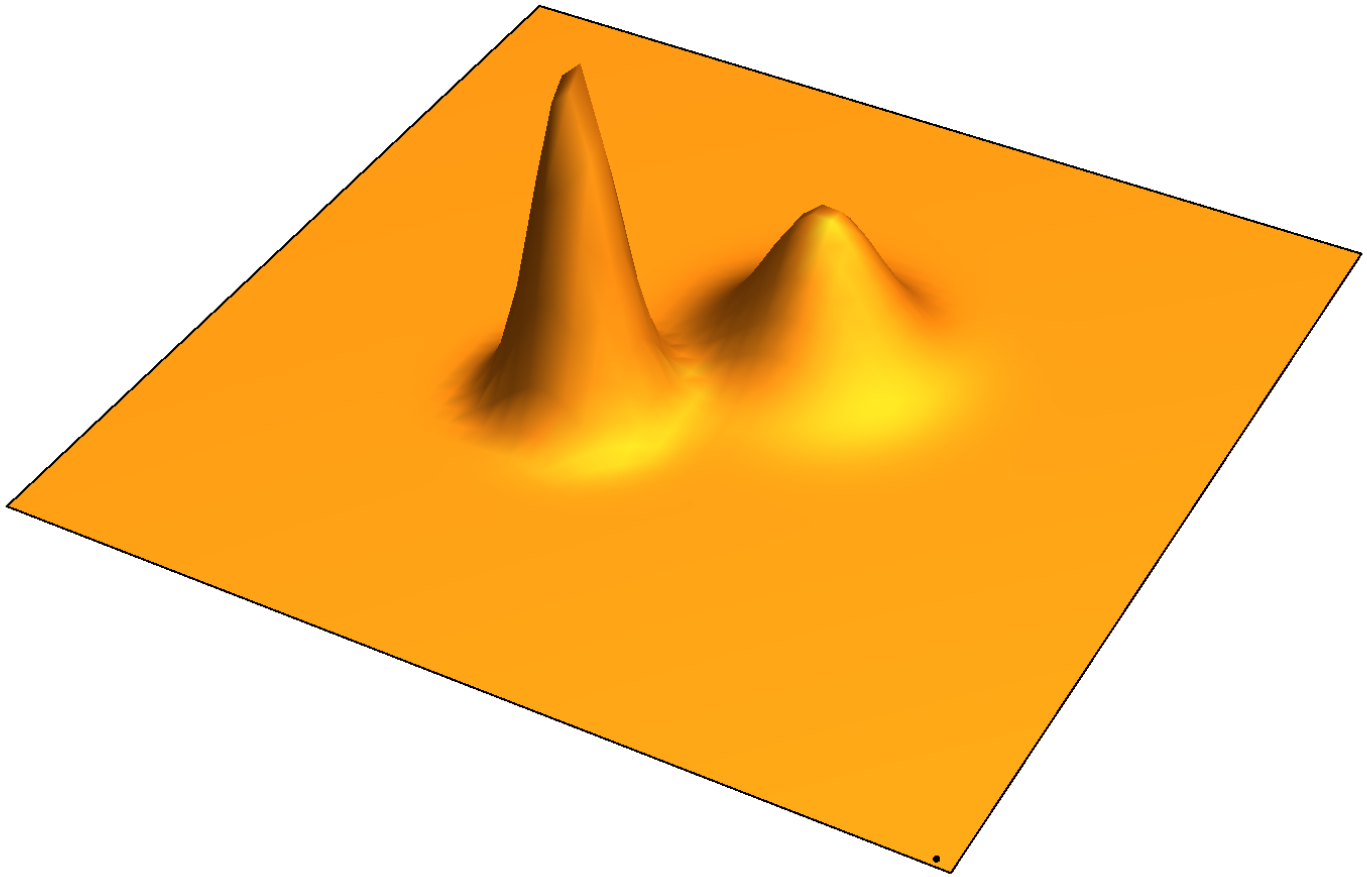


Another naive integration approach



Metropolis-Hastings algorithm





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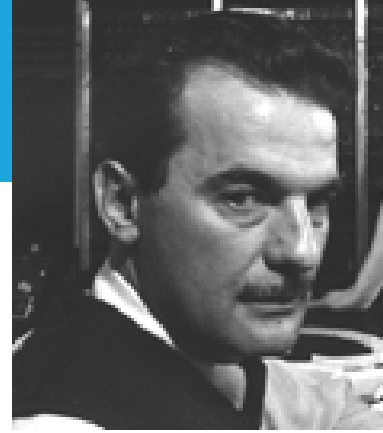

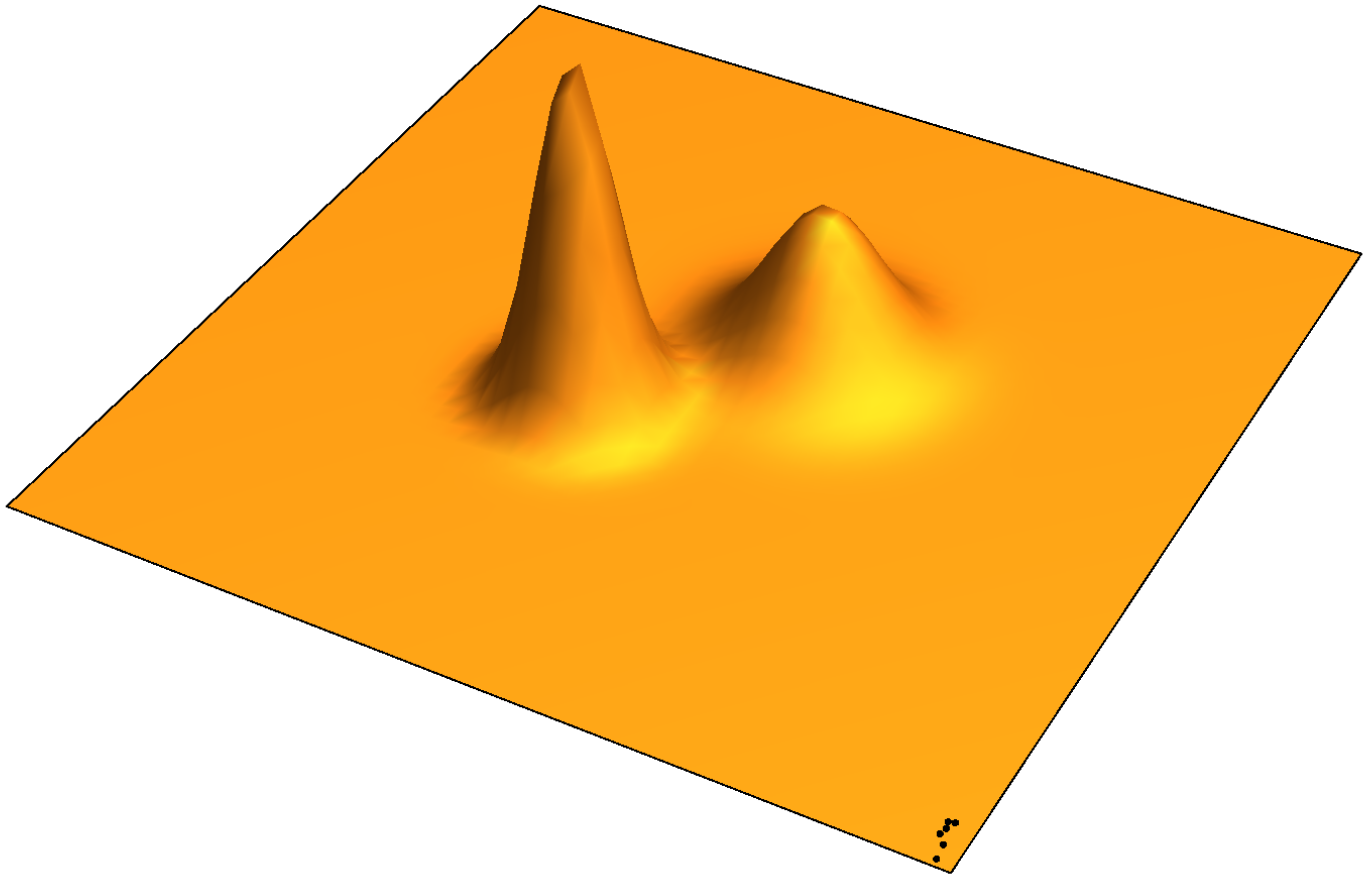

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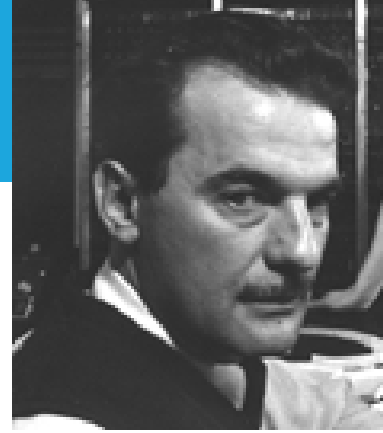

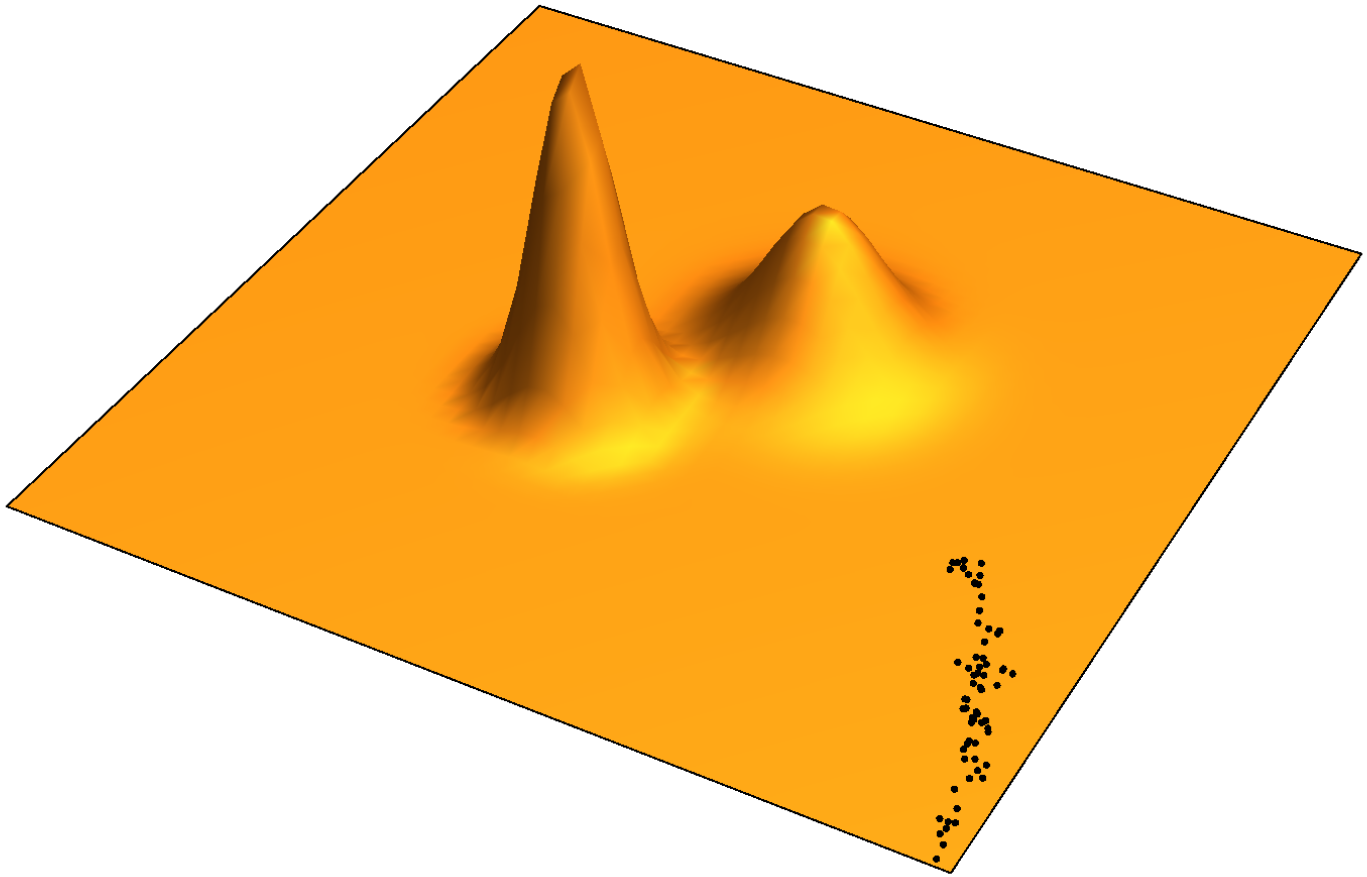

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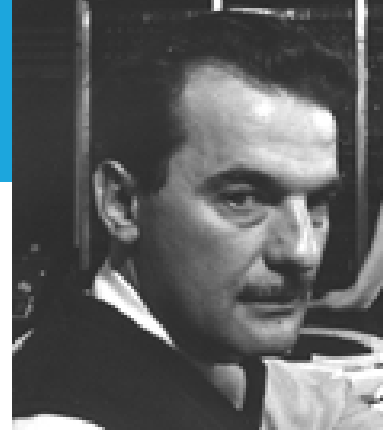

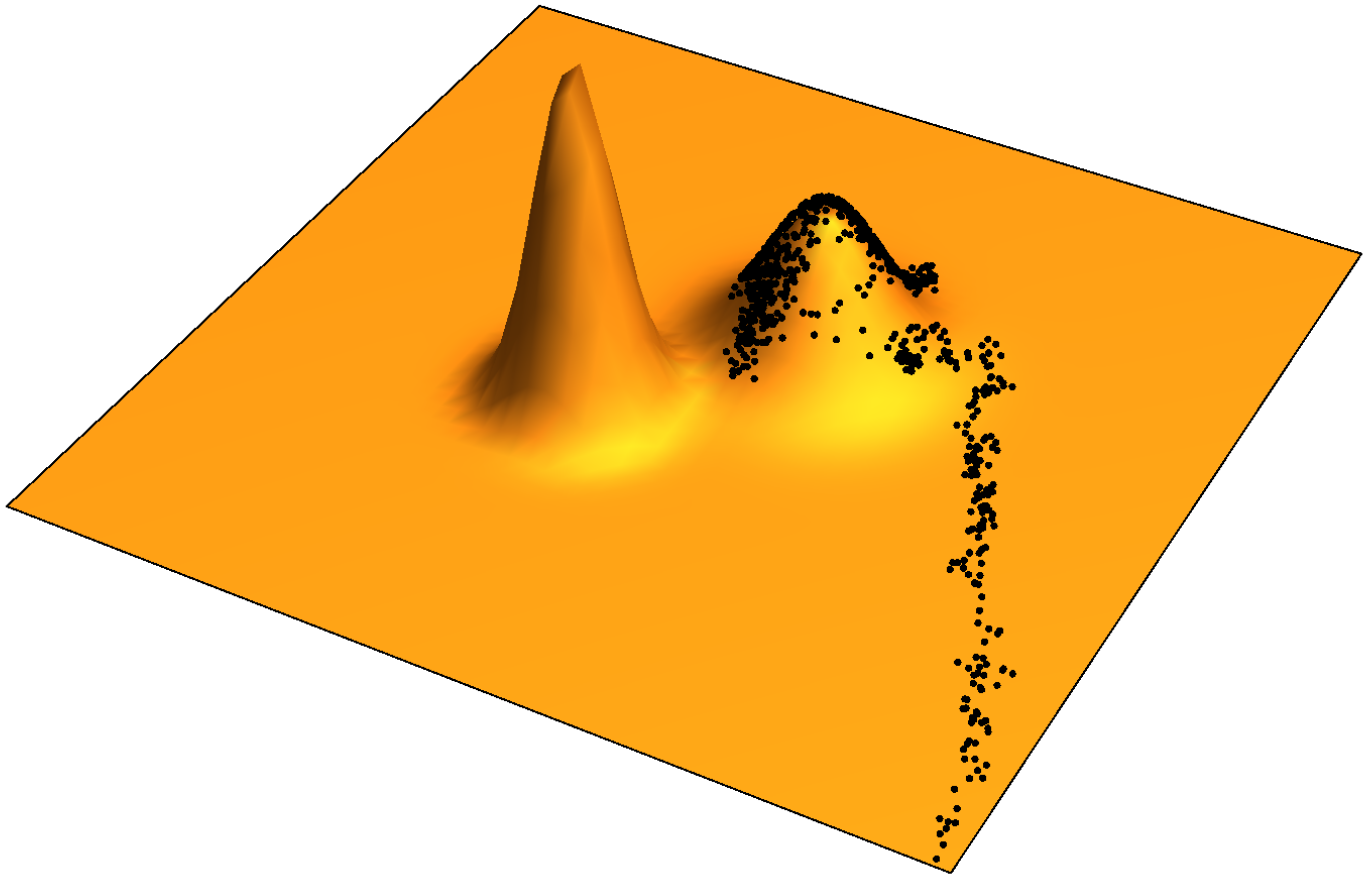

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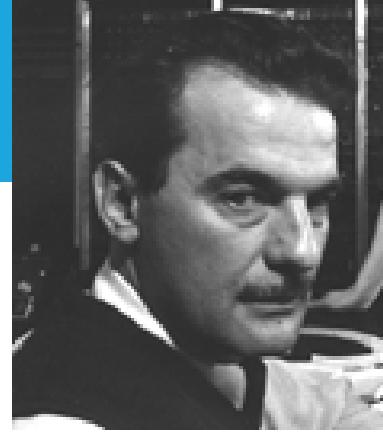

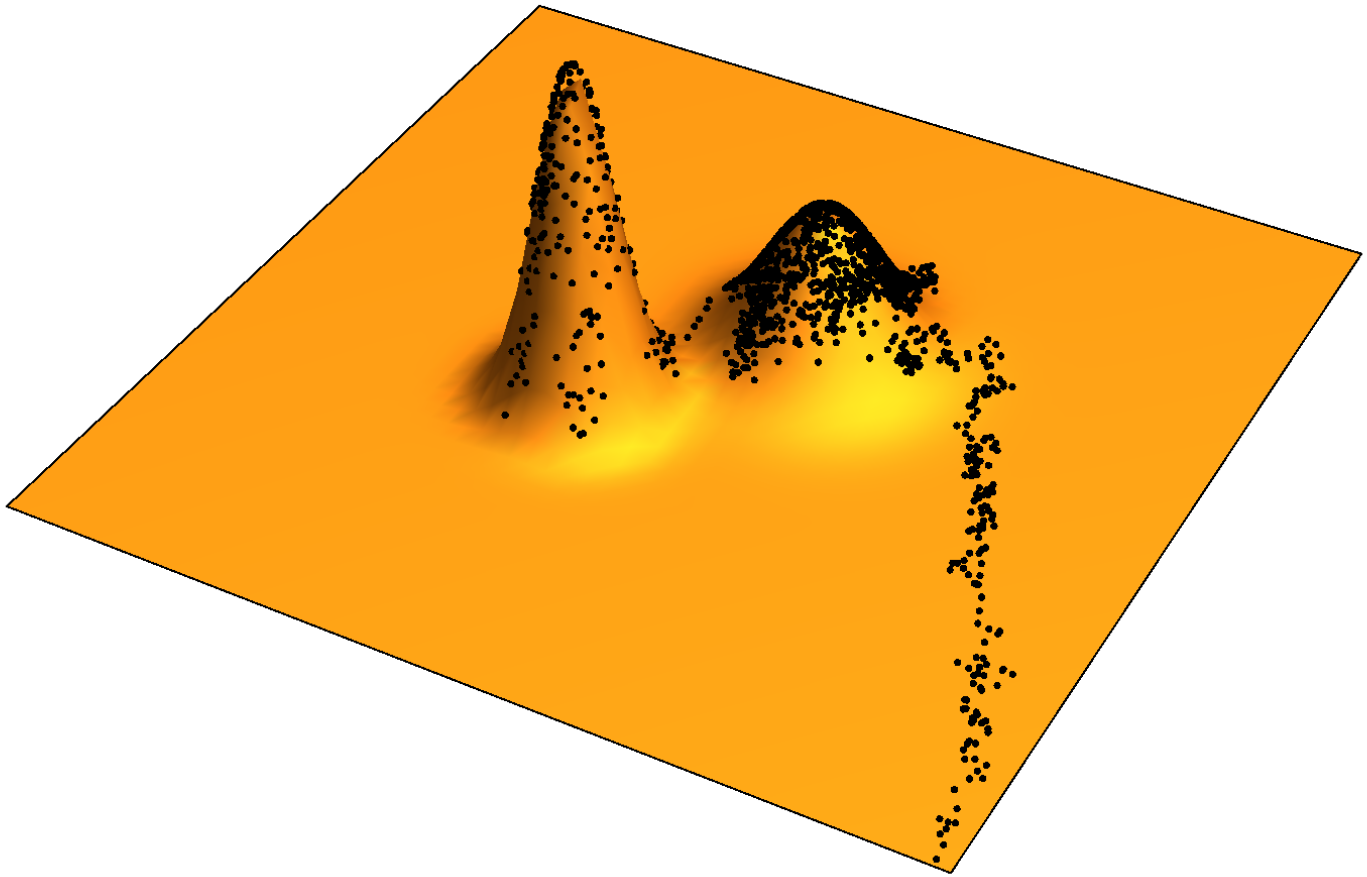

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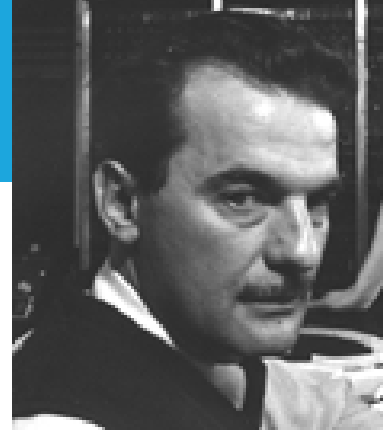

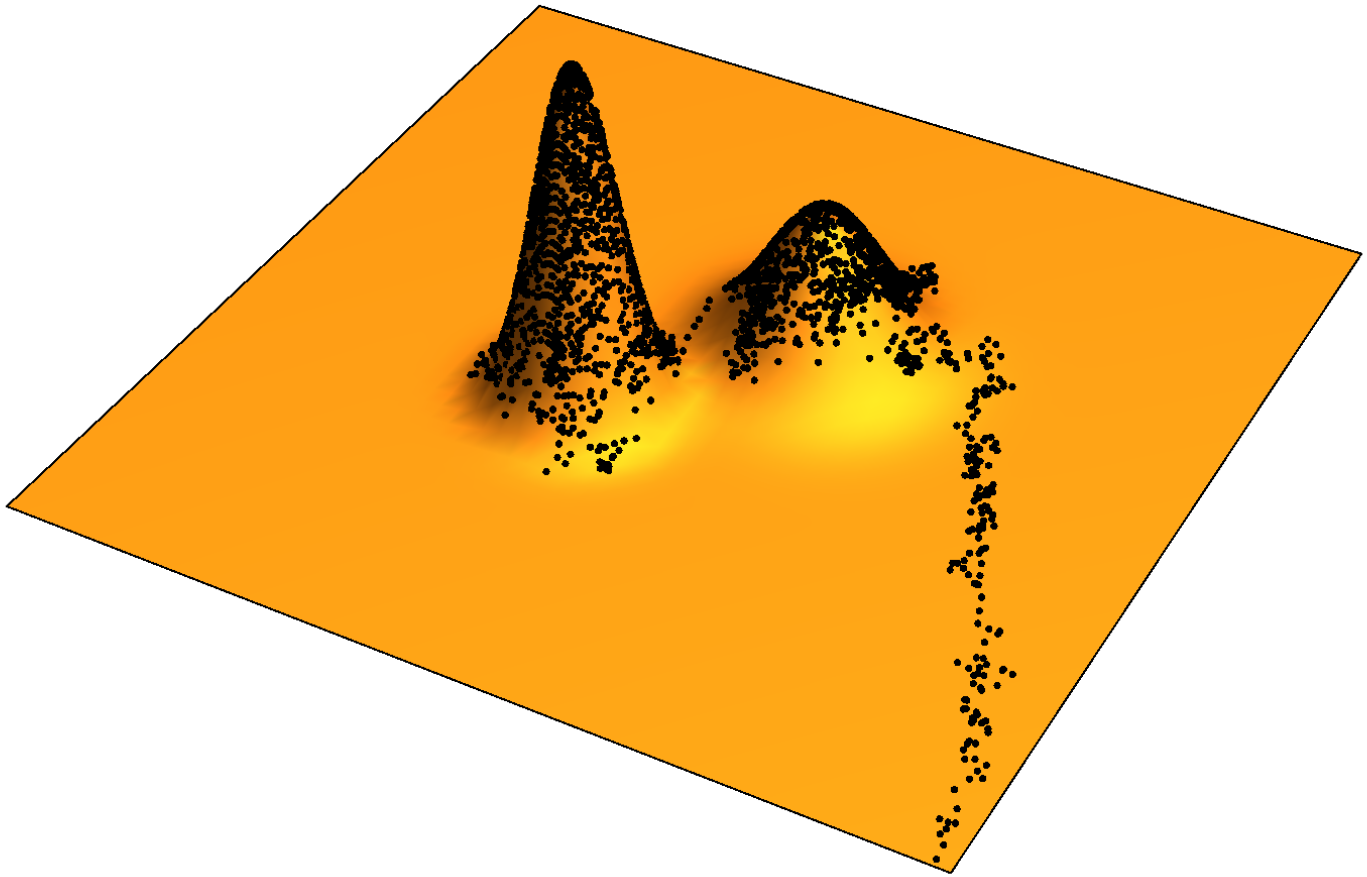

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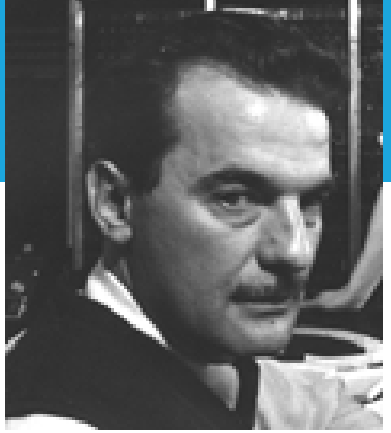
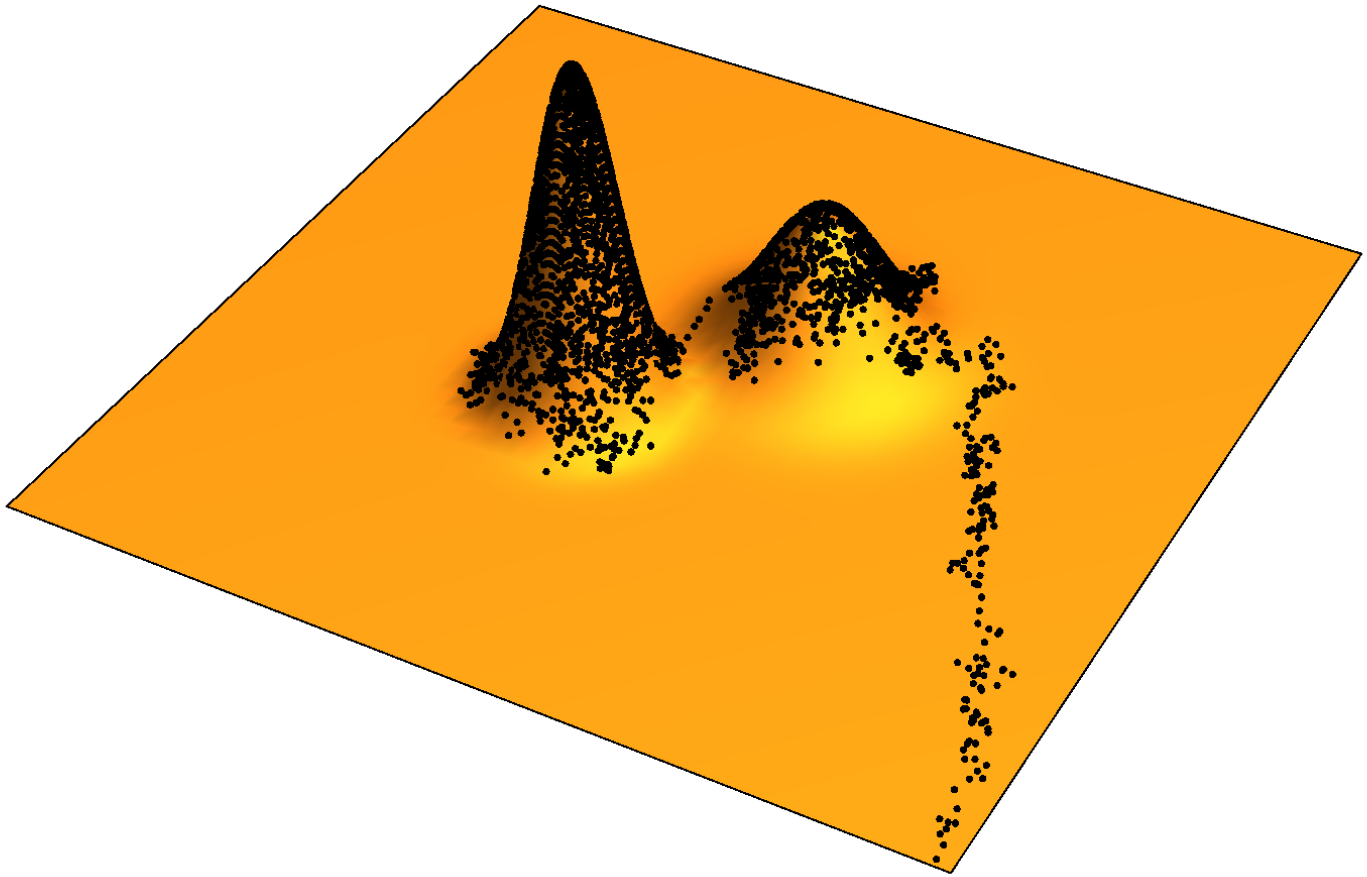


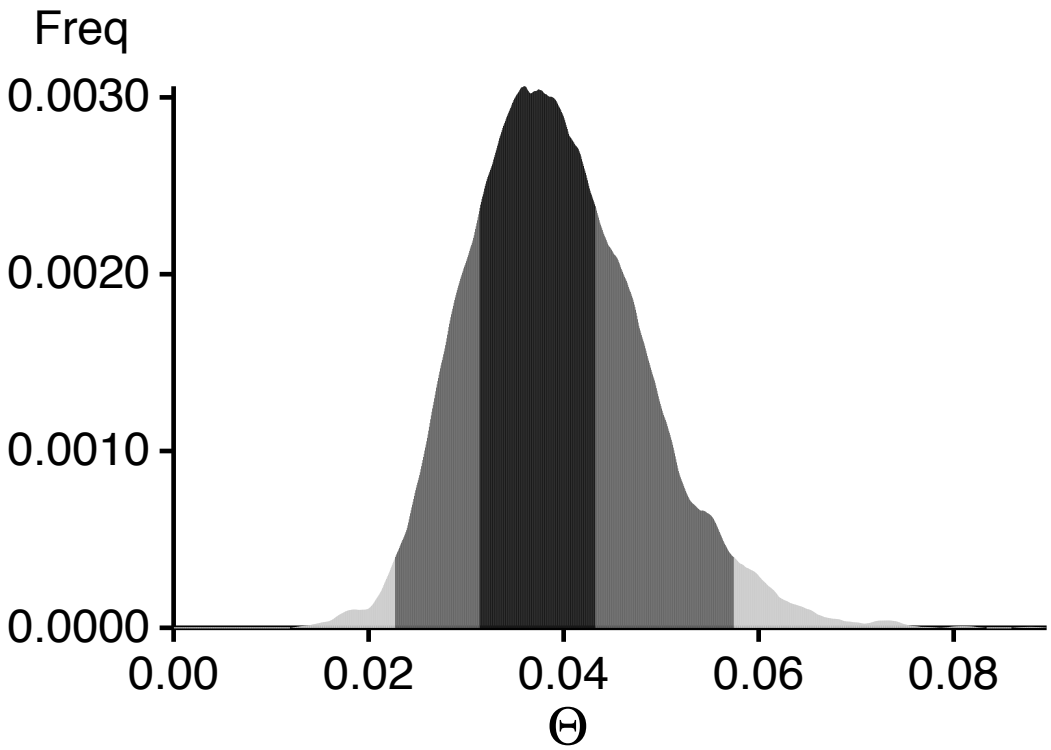


Inference of population size

Nuu-Chah-Nulth

Around 1930 – Friendly Cove, Vancouver Island

Inference of population size



Bayesian inference: $\Theta = 0.036$

Ward *et al* calculated $\Theta_{Ewens} = 0.043$

With a mutation rate of 0.32/site/million year and a generation time of 27 years we get $N_{\text{women}} = 2082$. Assuming same numbers of men and women and on average 2 children we get $N = 8328$.

Proc. Natl. Acad. Sci. USA
Vol. 88, pp. 8720–8724, October 1991
Evolution

Extensive mitochondrial diversity within a single Amerindian tribe

(population genetics/molecular anthropology/Pacific Northwest/human evolution)

R. H. WARD*, BARBARA L. FRAZIER*, KERRY DEW-JAGER*, AND SVANTE PÄÄBO†

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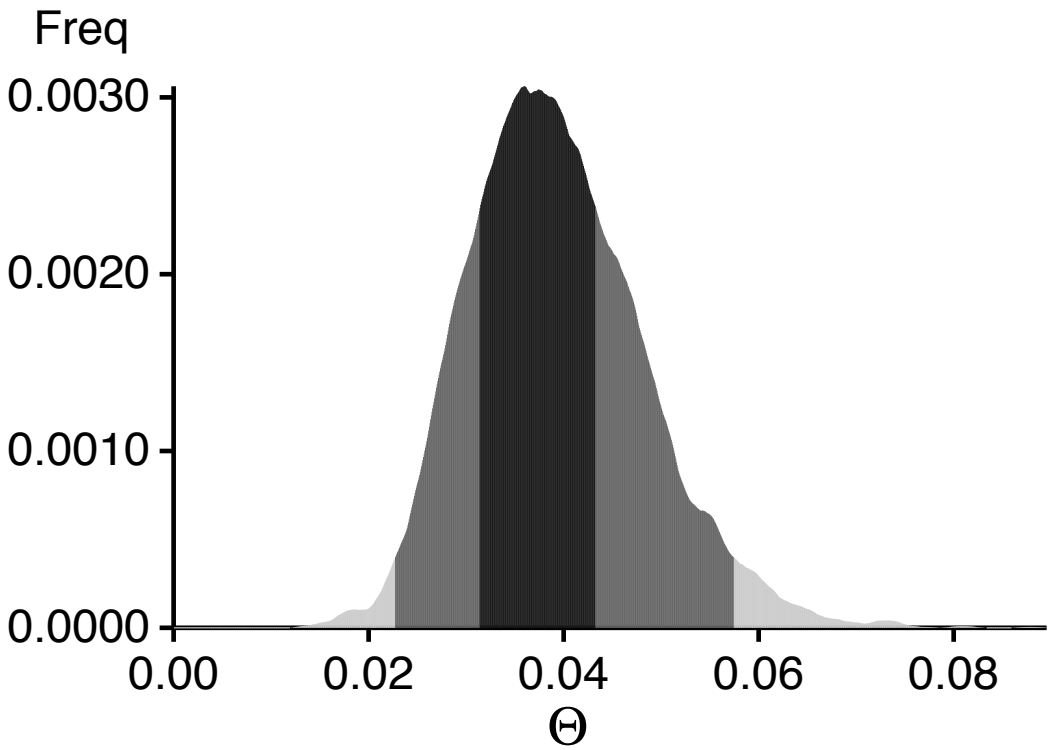
[The Nuu-Cha-Nulth are organized in 14 nations totaling 8147 (Nuuchahnulth tribal council Indian registry from February 2006)]

Inference of population size

Nuu-Chah-Nulth



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This sounds very good, but the 95%credibility interval is $5000 - 13500$

References

Coalescent:

Nuu-Cha-Nulth population size: J. Felsenstein. 1971. Inbreeding and variance effective numbers in populations with overlapping generations. Genetics 68:581-597;

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Sample size:

Felsenstein, J.2005. Accuracy of coalescent likelihood estimates: Do we need more sites, more sequences, or more loci? MBE 23: 691-700.

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